X-RAY LUMINOSITY FUNCTIONS OF NORMAL GALAXIES IN THE GREAT OBSERVATORIES ORIGINS DEEP SURVEY

Andrew Ptak, Bahram Mobasher, Ann Hornschemeier, Franz Bauer, And Colin Norman Received 2005 October 19; accepted 2007 May 29

ABSTRACT

We present soft (0.5-2 keV) X-ray luminosity functions (XLFs) in the Great Observatories Origins Deep Survey (GOODS) fields derived for galaxies at $z \sim 0.25$ and 0.75. SED fitting was used to estimate photometric redshifts and separate galaxy types, resulting in a sample of 40 early-type galaxies and 46 late-type galaxies. We estimate k-corrections for both the X-ray/optical and X-ray/NIR flux ratios, which facilitates the separation of AGNs from the normal/starburst galaxies. We fit the XLFs with a power-law model using both traditional and Markov-Chain Monte Carlo (MCMC) procedures. A key advantage of the MCMC approach is that it explicitly takes into account upper limits and allows errors on "derived" quantities, such as luminosity densities, to be computed directly (i.e., without potentially questionable assumptions concerning the propagation of errors). The slopes of the early-type galaxy XLFs tend to be slightly flatter than the late-type galaxy XLFs, although the effect is significant at only the 90% and 97% levels for $z \sim 0.25$ and 0.75. The XLFs differ between z < 0.5 and z > 0.5 at >99% significance levels for earlytype, late-type, and all (early- and late-type) galaxies. We also fit Schechter and lognormal models to the XLFs, fitting the low- and high-redshift XLFs for a given sample simultaneously assuming only pure luminosity evolution. In the case of lognormal fits, the results of MCMC fitting of the local FIR luminosity function were used as priors for the faint- and bright-end slopes (similar to "fixing" these parameters at the FIR values, except here the FIR uncertainty is included). The best-fit values of the change in $\log L^*$ with redshift were $\Delta \log L^* = 0.23 \pm 0.16$ dex (for early-type galaxies) and 0.34 ± 0.12 dex (for late-type galaxies), corresponding to $(1+z)^{1.6}$ and $(1+z)^{2.3}$. These results were insensitive to whether the Schechter or lognormal function was adopted.

Subject headings: stars: formation — surveys — X-rays: galaxies

Online material: color figures

1. INTRODUCTION

Recently, Norman et al. (2004, hereafter N04) presented the first X-ray luminosity functions (XLFs) of normal/starburst galaxies at cosmologically interesting redshifts (at $z \sim 0.3$ and 0.7). These XLFs were derived from the Chandra Deep Field North (2 Ms) and South (1 Ms) surveys (hereafter the CDF fields). It was found that the normal galaxy XLF was consistent in normalization and shape to the far-infrared LF, assuming a pure luminosity evolution of $\sim (1+z)^3$. While the errors were large due to the limited numbers of galaxies in each luminosity bin, the estimated star formation rate (SFR) derived from the XLFs was found to be consistent with other SFR measures, such as the H α luminosity. As discussed in N04, an interesting aspect of the XLF is that the X-ray emission from star-forming galaxies often has a large component due to X-ray binaries, particularly at hard X-ray energies above 2 keV. Therefore, the XLF is in part a probe of the binary stellar mass functions and hence indirectly a measure of the SFR. It is quite possible that binaries play a critical role in the evolution of the majority of stellar systems, and X-ray emission provides one of very few *direct* probes of such phenomena.

The Great Observatories Origins Deep Survey (GOODS; Giavalisco et al. 2003) is a multiwavelength survey of a subarea of the CDF fields. This survey entails deep imaging by three of NASA's Great Observatories: the Hubble Space Telescope, the Chandra X-ray Observatory, and the Spitzer Space Telescope, as well as extensive photometric and spectroscopic observation by ground-based facilities. These data allow the extension of the N04 results in several new directions. First, improved redshift determinations are now available for many of the sources. Second, the multiband data have been used to model the spectral energy distributions (SEDs) of the sources and estimate their spectral types. Therefore, XLFs can now be generated as a function of galaxy spectral type, which is the primary goal of this paper. Finally, we have improved the selection of AGNs versus normal/starburst galaxies over N04 by applying k-corrections to both X-ray/ optical and X-ray/NIR flux ratios. Such k-corrections are critical to correctly separating AGNs from normal galaxies (Bauer et al. 2004).

A motivation for deriving galaxy type–selected XLFs is the fact that there are multiple contributors to the X-ray emission of galaxies, namely low-mass X-ray binaries (LMXRB), high-mass X-ray binaries (HMXRB), hot ISM, AGNs, and to a lesser extent, individual supernovae (SNe) and massive stars, with the relative contribution of these expected to be dependent on galaxy type (see Ptak 2001 for a review). Specifically, the X-ray emission of early-type galaxies is known to be dominated by LMXRB emission and in the case of massive, gas-rich ellipticals, hot ISM (at the virial temperature of the galaxy). Late-type spiral galaxies should have significant contributions from both LMXRB and HMXRB populations, with the former being associated with the older ($t \gtrsim 10^8$ – 10^9 yr) populations. Late-type spiral galaxies often exhibit hot ISM

The Johns Hopkins University, Homewood Campus, Baltimore, MD 21218.
 Space Telescope Science Institute, 3700 San Martin Drive, Baltimore, MD

 ^{21218.} NASA/GSFC Laboratory for X-ray Astrophysics, Code 662, Greenbelt, MD 20771.

⁴ Chandra Fellow, Columbia University, 550 West 112th Street, New York, NY 10027.

⁵ European Southern Observatory, Karl-Schwarzschild-Strasse 2, Garching, D-85748, Germany.

due to heating associated with recent episodes of star formation. Finally, starburst galaxies should have the largest contribution from hot ISM (including potentially a superwind outflow) and HMXRB (see Persic & Rephaeli 2003).

We would therefore naively expect a somewhat different evolution in the X-ray luminosity density of these various galaxy types, with the LMXRB contribution following the global SFR history of the universe with a delay of the order of the evolutionary timescale of low-mass stars, i.e., $\sim 10^9$ yr (Ghosh & White 2001), and the HMXRB and hot ISM contribution tracking the SFR history instantaneously (relative to a Hubble time).

In this paper, we are primarily concerned with sources with $L_{\rm X} < 10^{42}~{\rm erg~s^{-1}}$, and hence any AGN present would be a lowluminosity AGN (LLAGN). LLAGN are found in all galaxy types, with LINERs preferentially being found in early-type galaxies (see Ho et al. 2003 and references therein). As in N04, we assess the impact of AGNs by classifying sources based on a Bayesian statistical analysis. In Appendix C we also separately address the properties of the X-ray flux ratios (i.e., ratio of X-ray to optical and near-IR flux), which provide important criteria for source classification. Ascertaining the presence of any luminosity or redshift dependence in the X-ray/optical or X-ray/NIR flux ratios would also indicate possible evolution of the relative contributions of different sources of X-ray flux in galaxies, i.e., LMXRB, HMXRB, hot gas, and AGNs (for examples of study of the evolution of X-ray/optical flux ratios, please see Ptak et al. 2001; Hornschemeier et al. 2002).

This paper is organized as follows. In § 2 we describe our sample selection and data analysis. The results of our analysis are given in § 3, and we discuss the results in § 4. In the Appendices, we give the details of our galaxy/AGN source classification procedure, basic statistics concerning the sample, and a discussion of the X-ray/optical and X-ray/NIR flux ratios. We assume the WMAP cosmology of $H_0 = 70 \text{ km s}^{-1} \text{ Mpc}^{-1}$, $\Omega_m = 0.3$, and $\Omega_{\Lambda} = 0.7$.

2. METHODOLOGY

2.1. X-Ray Sample Selection and Redshifts

The X-ray data used in our analysis are taken from Alexander et al. (2003), where the positions and X-ray fluxes for sources in both the CDF-N and CDF-S are tabulated. F. E. Bauer et al. (2007, in preparation) have carried out detailed matching of the GOODS ACS data with the Alexander et al. (2003) X-ray catalog, assigning matching probabilities based on optical faintness of potential counterparts, resulting in 263 (185) sources in the CDF-N (CDF-S) with a single ACS counterpart. We found 44 (28) CDF-N (CDF-S) sources with no optical counterpart, while 3 (4) with multiple counterparts were not included in our analysis. We also included the off-nuclear X-ray sources (14 in the CDF-N and 6 in the CDF-S; e.g., Hornschemeier et al. 2004). We then used the ACS coordinates in the Bauer catalog to match with the publicly available GOODS spectroscopic redshift catalogs⁶ and Mobasher et al. (2004) photometric redshift catalog to obtain both redshifts and spectral types for all sources. Stars were excluded.

The main parameters for each galaxy are the soft band (0.5-2.0 keV) flux (from Alexander et al. 2003), the hardness ratio [defined as (H-S)/(H+S), where H is the 2-8 keV vignetting-corrected count rate and S is the 0.5-2.0 keV vignetting-corrected count rate; from Alexander et al. 2003], the redshifts, and the optical spectral types. We broadly divided the spectral types into the groups early-type, late-type, and starburst/irregular galaxies.

However, we only extract XLFs for early-type and late-type galaxies, since the numbers of irregular galaxies were very small.

2.1.1. Photometric Redshifts

The photometric redshifts for GOODS fields are estimated using template fitting technique (Mobasher et al. 1996; Benitez 2001; Bolzonella et al. 2000; Dahlen et al. 2007). The rest-frame spectral energy distributions (SEDs) for galaxies of different types are convolved with filter response functions of the filters used in photometric observations of galaxies. The convolved SEDs, shifted in redshift space, were then fitted to observed SEDs of individual galaxies by minimizing the χ^2 function

$$\chi^2 = \sum_{i=1}^n \left[\left(F_{\rm obs}^i - \alpha F_{\rm template}^i \right) / \sigma^i \right]^2,$$

where the summation, i, is over the passbands (i.e., number of photometric points), and n is the total number of passbands. F_{obs}^{i} and F_{template}^{i} are, respectively, the observed and template fluxes at any given passband. Here σ^i is the uncertainty in the observed flux, and α is the normalization. The redshift and SED (i.e., spectral type) corresponding to the minimum χ^2 value for a given galaxy were then assigned to that galaxy. We used priors based on luminosity functions (LFs). The main effect of a LF prior is to discriminate between cases in which the redshift probability distribution function, which identifies the most likely redshift, shows two or multiple peaks (i.e., more than a single optimum redshift) due to confusion between the Lyman break and the 4000 Å break features. The absolute magnitudes of the object at the redshift peaks can then discriminate between these possibilities. Absorption due to intergalactic H I is included using the parameterization in Madau (1995).

We use template spectral energy distributions (SEDs) for normal galaxies consisting of E, Sbc, Scd, and Im from Coleman et al. (1980) and two starburst templates from Kinney et al. (1996; SB2 and SB3). To increase the spectral resolution, we construct intermediate-type templates by using the weighted mean of the adjacent templates and interpolating between them. This is done by defining five intermediate-type templates between the main spectral types used. Therefore, we use a total of 31 SED templates in this study. Photometric redshifts were then measured using the observed photometry in GOODS-N (*UBVRizJK*) and GOODS-S (*UBVRiJHK*).

2.1.2. Spectroscopic Redshifts

For a subset of the galaxies, we use the available spectroscopic redshifts. The majority of the spectroscopic redshifts were derived from Keck DEIMOS data for the CDF-N (see Wirth et al. 2004 and references therein) and VLT FORS2 or VIMOS data for the CDF-S (Szokoly et al. 2004; Vanzella et al. 2006).

The limiting magnitude for spectroscopic redshift determination was typically $R_{\rm AB}\sim 24$. If a spectroscopic redshift was not available, then a photometric redshift was used, with the photometric redshift limiting magnitude typically being $R_{\rm AB}\sim 25$ (Mobasher et al. 2004). In cases where a quality assessment was available for the spectroscopic redshift and was considered to be poor, and a photometric redshift with error $\delta z < 0.2$ was available, 7 the photometric redshift was used. In cases where there was no spectroscopic redshift quality given and a photometric redshift with error $\delta z < 0.1$ was available, the photometric redshift was

⁶ See http://www.stsci.edu/science/goods.

⁷ δz is defined as the 68% uncertainty on the photometric redshift derived from the posterior probability.

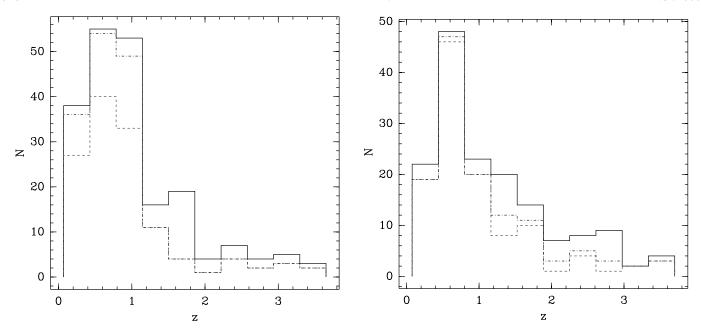


Fig. 1.—Redshift distributions for the CDF-N (*left*) and CDF-S (*right*) fields. The spectroscopic redshift distributions are shown with dot-dashed lines. In cases where the spectroscopic redshift quality was unknown or flagged as low and the photometric redshift error was small (see text for details), the photometric redshift was adopted. This distribution of sources for which a spectroscopic redshift was adopted is shown with dashed lines. [See the electronic edition of the Journal for a color version of this figure.]

used. This resulted in 204 (129) and 157 (116) total (spectroscopic) redshift determinations for the CDF-N and CDF-S sources, respectively (i.e., 73% and 82% of the GOODS X-ray sources with a unique optical counterpart have a redshift estimate). The redshift distributions are shown in Figure 1.

In cases where both a photometric and spectroscopic redshift were available, the mean absolute redshift difference was 0.09 (CDF-N) and 0.13 (CDF-S) after removing a small number of outliers with $\Delta z \equiv \left|z_{\rm spec} - z_{\rm phot}\right| > 0.5$. In Figure 2, we plot the difference between the photometric and spectroscopic redshift for sources with both measurements. The errors plotted in the figure are based solely on the photometric redshift errors. This figure shows that our typical redshift uncertainty is $\Delta z \sim 0.1-0.2$ for

 $z \lesssim 1.5$ and increases significantly for z > 1.5. We conservatively limit our analysis to $z \le 1.2$. This leaves a total of 148 sources in the CDF-N and 95 sources in the CDF-S region, of which 138 and 81 have soft band X-ray detections. Note that 131 (CDF-N) and 75 (CDF-S) of the X-ray sources have $L_X < 1 \times 10^{43}$ ergs s⁻¹, the highest X-ray luminosity considered in the X-ray luminosity functions discussed here.

2.2. K-Corrections

The soft band X-ray fluxes, listed in Alexander et al. (2003), are based on the observed count rates and a count rate to flux conversion computed from an effective photon index (based on band ratios). This photon index (Γ) was used to "k-correct"

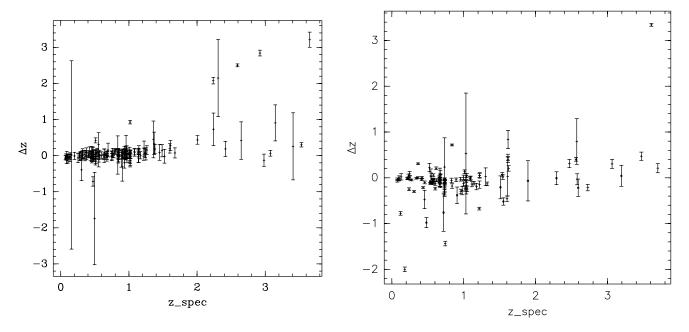


Fig. 2.— $\Delta z = z_{\text{spec}} - z_{\text{phot}}$ plotted as a function of z_{spec} for CDF-N (*left*) and CDF-S (*right*). In the CDF-S plot, the sources with poor spectroscopic redshift determinations (often based on a single line identification) are shown with crosses. The errors plotted for the points are based solely on the photometric redshift error. [*See the electronic edition of the Journal for a color version of this figure*.]

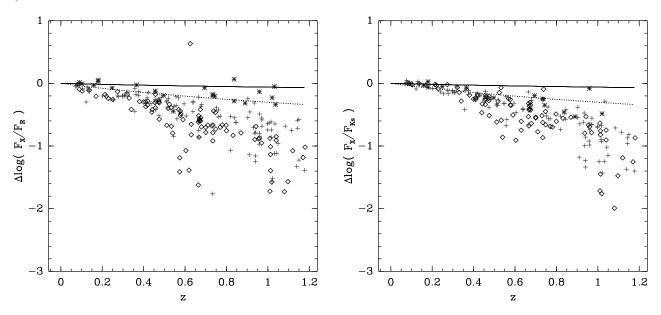


Fig. 3.—K-corrections applied to the X-ray flux ratios for the R (left) and K_s (right) bands, for both of the GOODS fields. The points show $\Delta \log [F_X/F_{\rm opt}] = \log [(F_X/F_{\rm opt})_{z=0}] - \log [(F_X/F_{\rm opt})_{z=0}]$, where $\log [(F_X/F_{\rm opt})_{z=0}]$ is the flux ratio after applying k-corrections, and $\log [(F_X/F_{\rm opt})_z]$ is the "observed" flux ratio. Early-type sources are marked with diamonds, late-type sources are marked with plus symbols, and irregular/starburst sources are marked with asterisks. The solid and dotted lines show the X-ray k-correction for $\Gamma = 1.8$ and $\Gamma = 1.0$, respectively. [See the electronic edition of the Journal for a color version of this figure.]

the observed fluxes, i.e., $F_{X,kcor} = (1+z)^{\Gamma-2}F_{X,obs}$ and $L_X = 4\pi d_L^2 F_{X,kcor}$, where d_L is the luminosity distance. In Appendix A, we present the sources, the adopted redshifts, and X-ray luminosities.

For a given band Q, we define F_Q to be νF_{ν} at the central frequency of the band ν_O . Since the GOODS magnitudes are in the AB system, $\log (F_X/\tilde{F}_Q)$ is given by $\log F_X + 0.4(Q + 48.6) \log \nu_O$, e.g., in the case of the R band, $\log (F_X/F_R) = \log F_X +$ $0.4R_{AB} + 4.8$. The optical magnitudes were k-corrected by interpolating between the magnitudes whose central wavelengths bracket $(1+z)\lambda$, where λ is the central wavelength of the filter of interest (e.g., 4400 Å for the B band). The bandpass correction term (1+z) was also included in the k-correction (Blanton et al. 2003). In Figure 3, we show the k-corrections applied to R- and K_s -band flux ratios. Note that in the case of K_s band magnitudes, the interpolation discussed above necessarily amounts to an extrapolation since K_s is the longest wavelength band in the data set and every source considered here has z > 0. The corrections are plotted separately for each spectral type, where it can be seen that the largest k-corrections occur for B-band fluxes from early-type galaxies, which is not surprising, considering their red color. Also, as mentioned in Bauer et al. (2004), the k-corrections to the flux ratios are dominated by the optical correction rather than the X-ray correction.

2.3. Luminosity Function Estimation

We construct binned luminosity functions using the estimator described in Page & Carrera (2000). Briefly, this assumes negligible variation in both the luminosity function and its evolution across each $\Delta L \Delta V(z)$ bin, where V(z) is the comoving volume:

$$N \simeq \phi(L,z) \int_{L_{\min}}^{L_{\max}} \int_{z_{\min}}^{z_{\max}} C(L,z) V(z) \, dz \, dL, \tag{1}$$

N is the number of sources in the XLF bin bounded by L_{\min} , L_{\max} , z_{\min} and z_{\max} , and C(L,z) is a completeness correction. Note that z_{\max} is a function of luminosity, since the limit is chosen at the highest observable redshift given the limiting flux [or,

equivalently, C(L,z)=0 where $F(L,z) < F_{\rm lim}$]. The error for each XLF bin is accordingly derived from the Poisson error on the number of galaxies in each bin (note that this assumes negligible statistical and systematic error in the volume integral terms), and for bins with no galaxies, we used N=1.841 as the $1\,\sigma$ upper limit (Gehrels 1986). Small luminosity function bins are preferred; however, larger bins minimize Poisson noise and reduce the effect of luminosity uncertainties within a bin (due to redshift uncertainties). For example, for the uncertainty in luminosity resulting from redshift error to be on the order of the bin size, bin sizes of $\Delta \log L=0.5$ are required for sources with $z \gtrsim 0.4$ and redshift errors of 0.1 or less. Given the small number of sources, we divide our sample into just two redshift bins, z < 0.5 and 0.5 < z < 1.2. We used the soft band GOODS sky coverage shown in Treister et al. (2004).

2.4. Completeness Correction

The completeness correction C(L, z) is given by the product of the X-ray detection completeness, the probability of a given X-ray source having an optical/NIR counterpart, and the probability of a counterpart having a redshift. The latter two terms reduce to the probability of an X-ray source having a redshift. For the X-ray detection completeness, we used the results of simulations originally performed for the full CDF areas discussed in Bauer et al. (2004), where the effective solid angle for each source was computed (i.e., the maximum solid angle over which the source could have been detected). The ratio of the effective solid angle to the (total) CDF solid angle at the flux of the source is then an estimate of the completeness for sources at a similar flux and position. We then computed the completeness as a function of flux by multiplying the mean per-source completeness (i.e., effectively averaging over off-axis angle) and the fraction of GOODS X-ray sources with a redshift in the given flux range (the bin sizes were 0.5 dex for the GOODS-N sources and 0.75 dex for the GOODS-S sources), shown in Figure 4. In both the GOODS-N and GOODS-S, the completeness ranged from $\sim 60\%$ at $F_{0.5-2.0 \, \rm keV} = 2 \times 10^{-17} \, \rm ergs \, cm^{-2} \, s^{-1}$ in the North and $F_{0.5-2.0 \, \rm keV} = 7 \times 10^{-17}$ in the South to 100% for $F_{0.5-2.0 \, \rm keV} > 10^{-14} \, \rm ergs \, cm^{-2} \, s^{-1}$, although with errors on the order of 20%

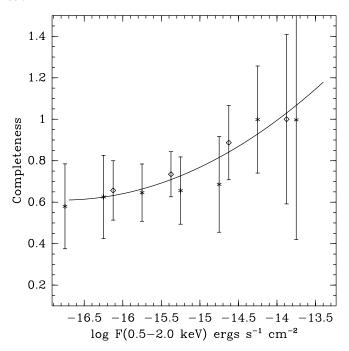


Fig. 4.—Completeness corrections as a function of flux calculated for the GOODS-N (*crosses*) and GOODS-S (*diamonds*) regions. The line shows a quadratic fit to the combined completeness corrections (i.e., to be applied to either data set).

in each flux bin. We fit the completeness with a quadratic function, both to avoid the impact of statistical fluctuations and to have an analytic expression for use in calculating equation (1). The best-fitting relation was $C(F_{\rm X}) = 14.56 + 1.66 \log F_{\rm X} + 0.0495 (\log F_{\rm X})^2$, which is shown in Figure 4.

2.5. Galaxy Classification

N04 classified sources as AGN1 (broad-line AGN), AGN2 (narrow-line AGN), and "galaxy" (i.e., normal/starburst galaxy) using a prior based on the observed distributions of X-ray luminosity (log L_X), the X-ray to optical flux ratio (log $F_X + R$), and X-ray hardness ratio derived from a subset of the sources in which the optical classification was secure. These distributions were used as priors along with the likelihoods of these measured values in computing the posterior probabilities. A source was then classified as a galaxy if the galaxy probability exceeded the AGN1 and AGN2 probabilities.

Here we employ a similar algorithm, discussed in Appendix A. However, there are a number of differences with the approach in N04, including use of the X-ray/K-band flux ratio, inclusion of k-corrections in the computation of X-ray/optical and X-ray/NIR flux ratios, and a more conservative galaxy selection threshold. Note that our priors are based on the spectroscopic data in Szokoly et al. (2004), which covers the entire CDF-S area. In Szokoly et al., only R- and K-band values are listed, and accordingly, k-corrections could not be interpolated, as was done with the GOODS data (see § 2.2). However, we estimated the spectral type of the sources based on their R-K color and applied a mean k-correction derived for each type.

In the case of the X-ray hardness ratio, the means and standard deviations of the prior samples were -0.19 ± 0.46 for galaxies, 0.16 ± 0.37 for AGN2, and -0.51 ± 0.05 for AGN1, very similar to those used in N04. For comparison, Peterson et al. (2006) artificially redshifted a sample of local galaxies to z = 0.3 and found that hardness ratios of AGN2 were split evenly around HR = 0. They noted that soft emission from starbursts could be

resulting in HR < 0 in AGN2/starburst composites (the case for four out of the seven AGN2 with HR < 0). Our final normal/starburst galaxy sample consisted of 64 CDF-N sources and 23 CDF-S sources. We also consider an "optimistic" normal/starburst galaxy sample which includes all sources that were *not* classified as an AGN. This sample therefore includes sources with ambiguous classification, similar to the original selection in N04. The optimistic normal/starburst sample contains 98 CDF-N and 54 CDF-S sources.

2.6. XLF Model Fitting

In order to quantify differences among the various XLFs discussed here, we fit the XLFs with several models (note that all luminosity functions discussed here are binned). The simplest model is a linear function in $\log \phi$ - $\log L_{\rm X}$ space, i.e., $\log \phi(L_{\rm X}) = a \log L_{\rm X} + b$. We performed linear fits with two conventional approaches, namely, using the survival analysis linear regression (hereafter Method 1) discussed in Isobe et al. (1986) and least-squares fitting (using the qdp program) after excluding the upper limits (Method 2). The linear regression approach has the disadvantage of not explicitly including any error information (i.e., the algorithm only takes as input the value of each detection or limit), and the confidence level used in computing an upper limit is arbitrary.

2.6.1. Markov-Chain Monte Carlo Fitting

Bayesian parameter estimation has been gaining popularity in various fields of astronomy, most notably in cosmology (see, e.g., Spergel 2007). A key advantage of Bayesian parameter estimation is that common statistical issues, such as the treatment of upper limits in fitting and propagation of errors, are inherently handled properly. In addition, the probability distributions for parameters of interest are computed and can be shown, rather than simply a summary, such as the 68% confidence interval. This latter point is particularly relevant when there are relative minima in the fitting statistic, in which case the meaning of the traditional error bar is not well defined.

The basis of Bayesian parameter estimation is the computation of the posterior probability distribution,

$$p(\theta|D) = \frac{p(\theta)p(D|\theta)}{p(D)},$$
 (2)

where θ is the vector of model parameters (e.g., $\theta = \{a,b\}$ for the linear model), D represents the data, $p(\theta)$ is the prior probability distribution for the parameters, $p(D|\theta)$ is the likelihood function, and p(D) is a normalization constant (i.e., since it does not depend on the parameter values). Here p(D) is given by $\int p(\theta)p(D|\theta)d\theta$. In practice, computing $p(\theta|D)$ is computationally difficult, since it requires an n-dimensional integral when fitting an n parameter model. The Markov-chain Monte Carlo (MCMC) procedure circumvents this by performing a directed random walk through the parameter space (van Dyk et al. 2001; Ford 2006). Here we assume Gaussian priors for each parameter (with mean and standard deviation μ_{θ_i} and σ_{θ_i} for parameter θ_i) and choose very large σ_{θ_i} values for the case of a flat (uninformative) prior. The likelihood function for the number of counts in an XLF bin is the Poisson distribution, giving

$$p(\theta|D) \propto \prod_{i=1}^{n} G(\theta_{i}|\mu_{\theta_{i}}, \sigma_{\theta_{i}})$$

$$\times \prod_{j=1}^{m} Pois[N_{j}|\phi(\theta_{1},...,\theta_{n}, L_{X,j}, z_{j})V(L_{X,j}, z_{j})], \quad (3)$$

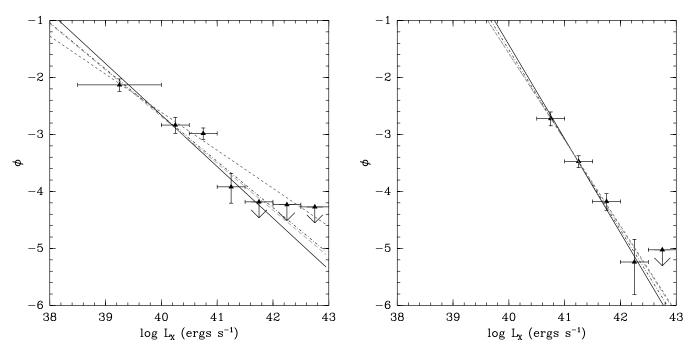


Fig. 5.—Full-sample (early + late type galaxies) XLFs from the GOODS fields for the redshift intervals z < 0.5 (left) and 0.5 < z < 1.2 (right). The solid, dashed, dot-dashed, and long-dashed lines show, respectively, the best-fitting models from Method 1 (survival analysis linear regression), Method 2 (least-squares, excluding upper limits), MCMC fitting with Method 1 results used to initialize the fit, and MCMC fitting with Method 2 results used to initialize the fit. [See the electronic edition of the Journal for a color version of this figure.]

where there are *n* model parameters and *m* XLF bins. $G(\theta_i|\mu_{\theta_i},$ σ_{θ_i}) is the Gaussian prior for parameter θ_i . N_i is the number of sources, and $V(L_{X,j},z_i)$ is the comoving volume corresponding to the *j*th bin (see eq. [1]). Here $\phi(\theta_1,...,\theta_n,L_{X,j},z_j)$ is the model XLF evaluated at $L_{X,j}$ and z_j . The "Pois" terms give the likelihood of detecting N_i galaxies in XLF bin j given an expectation of ϕV . In addition to the simple linear model, we also fitted Schechter and lognormal functions to the individual XLFs. We used the Metropolis-Hastings MCMC algorithm, in which a "proposal" distribution is used to guide the variation of the parameters (see Ford 2005; Russell et al. 2007). In this procedure, random offsets for each parameter are drawn from the proposal distribution, and accordingly, the step sizes (i.e., the Gaussian σ) are preferably on the order of the final error for that parameter. A step is accepted if the probability of the model given the new parameter values is higher, and also at random intervals when the probability is lower (i.e., occasionally the fit is allowed to proceed "downhill" to avoid relative minima). For a given run, three chains were produced with a length of at least 2×10^6 iterations, and the parameter step sizes were adjusted during the first chain to achieve acceptance rates in the range 0.35-0.4. Only the last chain in a given run was used for analysis. At least 10 runs were performed for a given fit, and we computed the convergence Rstatistic from Gelman et al. (2004). R values ≤1.2 indicate convergence, and in every case the R values were <1.01. In practice, the linear model parameters a and b were highly correlated, which leads to inefficient MCMC fitting. As discussed in Gelman et al. (2004), we addressed this by instead fitting for a_1 , a_2 , and b, where $a = a_1 + a_2 b$. The initial values of a_1 and a_2 were determined from a linear regression fit to the output of a short MCMC run, and the value of a_2 was held approximately fixed by using a tight prior (a_1 and a_2 are "nuisance" parameters and are not discussed further). The initial parameter values for a given run were chosen randomly, and the range of the allowed values was ± 2 times the initial step size from the prior mean. The dispersion between the best-fitting parameter values from the runs in a given fit was always <10% of the final error in the parameters, showing that the chains converged independently of the starting values. For each parameter of interest, the chain values were binned into normalized histograms, which represent the marginalized posterior probability distribution of the parameter. The peak value (i.e., the mode) was taken as the bestfit value, and the 68% confidence interval was derived from the probability density by steeping from the peak in the direction of the smallest decrease until the integrated area equaled 68%, as discussed in Kraft et al. (1991). In order to assess how constraining the prior is, we computed the ratio of the prior σ values to the standard deviation of the MCMC parameter values (i.e., the 68% error in the case of a Gaussian posterior probability distribution). In tables showing the fit results, we marked parameter values and errors as having a tightly constraining prior when this ratio is <1.1 and as a moderately constraining prior when it is between 1.1 and 2.0.

2.6.2. Fit Quality

Unfortunately, the MCMC procedure does not directly provide a model probability estimate, since this requires that the normalization of equation (2) be computed. Since efficient model selection is complicated (and the subject of active research; see, e.g., Trotta 2007), we defer this, and here we compute χ^2 for the models after excluding upper limits. This also has the advantage that it can be consistently applied to all of the model fitting methods discussed here in the case of the linear model.

3. RESULTS

3.1. Linear Model Fits

In Figures 5, 6, 7, 8, and 9, we show the XLFs derived from the GOODS fields based on the early-type, late-type, and total

⁸ We used the Gaussian distribution as the proposal distribution, as is common practice.

PTAK ET AL.

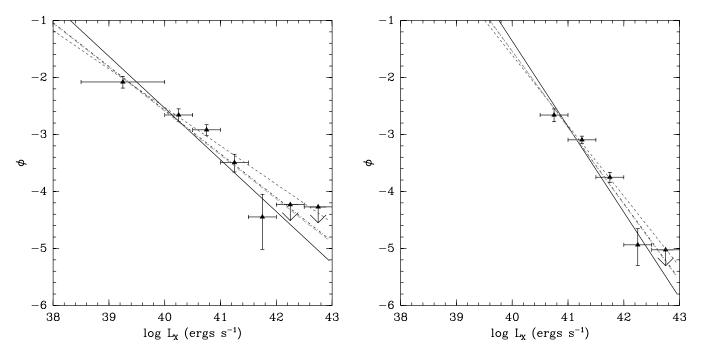


Fig. 6.—"Optimistic" sample (early + late type galaxies) XLFs from the GOODS fields for the redshift intervals z < 0.5 (left) and 0.5 < z < 1.2 (right). Lines are as in Fig. 5. [See the electronic edition of the Journal for a color version of this figure.]

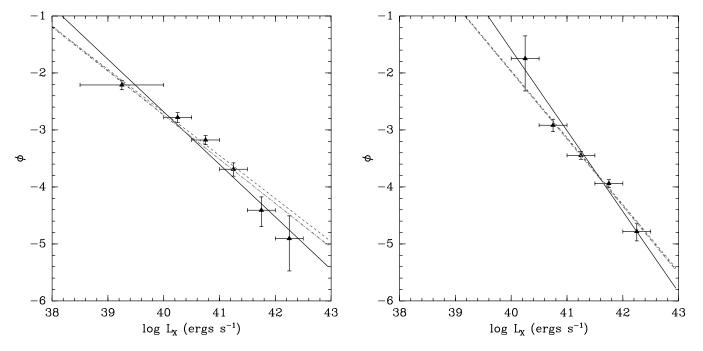
early- + late-type galaxy samples along with the linear model fits. We also show the fits to the "optimistic" galaxy sample and the N04 XLFs for comparison. The best-fit parameters and errors are given in Table 1 along with the associated χ^2 values and the probability with which the model can be rejected (again with the caveat that this probability does not include upper limit data).

As is evident from the fit results, the MCMC fitting was not very sensitive to whether the Method 1 or Method 2 fit values were used as the prior means. Interestingly, in every case, the MCMC fits resulted in curves that were intermediate to the Method 1 and Method 2 fits and tended toward the Method 1 results when upper limits were constraining (e.g., the early-type

and late-type galaxy XLFs) and toward Method 2 when the no upper limits were used or the upper limits were not constraining (e.g., the N04 XLFs). The marginalized posterior probability distributions for a and b were close to Gaussian, resulting in nearly symmetric errors. In general, the linear models cannot be rejected at high confidence (i.e., >3 σ). Hereafter, the fit results presented refer to MCMC fitting.

3.2. Joint Linear Model Fits

In Figures 10, 11, and 12, we show the results of fitting corresponding pairs of XLFs simultaneously (i.e., the low- and high-redshift XLFs for a given sample and the early- and late-type



 $Fig. \ 7. \\ --N04 \ XLFs \ for \ z < 0.5 \ (\textit{left}) \ and \ 0.5 < z < 1.2 \ (\textit{right}), \ with \ linear \ fits \ shown \ as \ in \ Fig. 5. \ [\textit{See the electronic edition of the Journal for a color version of this figure.}]$

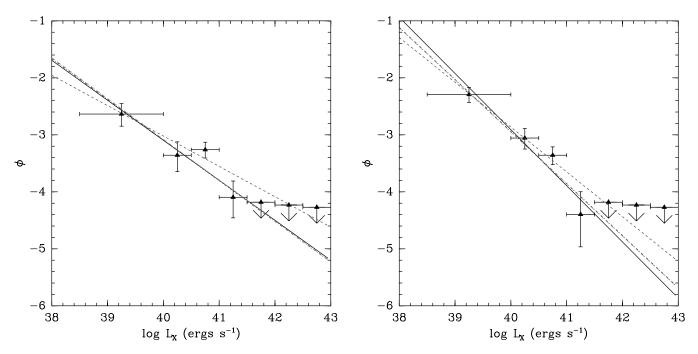


Fig. 8.—z < 0.5 early-type (left) and late-type (right) galaxy XLFs from the GOODS fields. Lines are as in Fig. 5. [See the electronic edition of the Journal for a color version of this figure.]

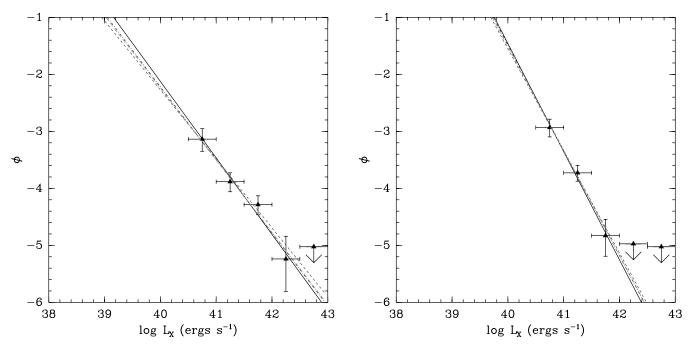


Fig. 9.—0.5 < z < 1.2 early-type (left) and late-type (right) galaxy XLFs from the GOODS fields. Lines are as in Fig. 5. [See the electronic edition of the Journal for a color version of this figure.]

TABLE 1 XLF Linear Fits

Sample	Method	a	b	$\chi^2/{ m dof}$	$p_{\chi^2} \ (\%)$
Early-type galaxies, low-z	1	-0.70 ± 0.11	25.0 ± 4.5	3.58/2	83.2
	2	$-0.53^{+0.13}_{-0.13}$	$18.3^{+5.4}_{-5.4}$	6.34/2	95.6
	3*	$-0.71^{+0.12}_{-0.09}$	$25.4_{-3.3}^{+5.2}$	3.66/2	83.9
	3^{\dagger}	$-0.72^{+0.11}_{-0.10}$	$25.6_{-3.3}^{+5.2}$	3.70/2	84.2
Late-type galaxies, low-z	1	-0.99 ± 0.14	36.5 ± 5.5	3.78/2	84.8
	2	$-0.79^{+0.11}_{-0.12}$	$28.6^{+4.7}_{-4.6}$	5.55/2	93.6
	3*	$-0.91^{+0.09}_{-0.11}$	$33.6_{-3.9}^{+4.1}$	2.95/2	77.0
	3^{\dagger}	$-0.91^{+0.112}_{-0.08}$	$33.6_{-4.0}^{+4.1}$	2.97/2	77.2
All galaxies, low-z	1	-0.90 ± 0.11	33.5 ± 4.6	8.95/2	98.6
	2	$-0.67^{+0.09}_{-0.08}$	$24.1^{+3.4}_{-3.5}$	12.47/2	99.5
	3*	$-0.81^{+0.06}_{-0.08}$	$29.8_{-2.3}^{+3.3}$	6.64/2	96.2
	3^{\dagger}	$-0.82^{+0.06}_{-0.08}$	$30.1_{-2.5}^{-2.3}$	6.64/2	96.2
All galaxies (optimistic), low-z	1	-0.91 ± 0.11	33.8 ± 4.6	15.98/3	99.9
	2	$-0.67^{+0.07}_{-0.07}$	$24.4^{+2.9}_{-2.9}$	19.03/3	100.0
	3*	$-0.77^{+0.06}_{-0.06}$	$28.1_{-2.4}^{-2.4}$	9.85/3	98.0
	3†	$-0.77^{+0.06}_{-0.06}$	$28.2_{-2.5}^{-2.4}$	9.70/3	97.9
Norman et al. (2004), low-z	1	-0.92 ± 0.07	34.2 ± 2.8	20.24/4	100.0
,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,	2	$-0.76^{+0.05}_{-0.05}$	$27.5^{+2.2}_{-2.1}$	13.85/4	99.2
	3*	$-0.77_{-0.04}^{+0.05}$	28 2+2.0	10.56/4	96.8
	3†	$-0.78^{+0.04}_{-0.06}$	$28.4^{+1.7}_{-2.1}$	10.43/4	96.6
Early-type galaxies, hi-z	1	-1.34 ± 0.10	51.5 ± 4.2	1.17/2	44.4
, J. J. S	2	$-1.21^{+0.21}_{-0.21}$	$46.2^{+8.7}_{-8.7}$	1.25/2	46.3
	3*	$-1.28^{\substack{-0.21 \ +0.20}{+0.20}}_{\substack{-0.21}}$	$49.0_{-8.0}^{-8.7}$	1.07/2	41.4
	3 [†]	$-1.27^{+0.18}_{-0.23}$	$48.4_{-7.5}^{-8.0}$	1.07/2	41.3
Late-type galaxies, hi-z	1	-1.90 ± 0.10	74.5 ± 4.2	0.58/1	53.9
71 - 3	2	$-1.79^{+0.29}_{-0.29}$	$70.1^{+11.7}_{-11.7}$	0.52/1	51.5
	3*	$-1.85^{+0.29}_{-0.23}$	$72.6_{-8.8}^{-11.7}$	0.54/1	52.3
	3†	$-1.83^{+0.19}_{-0.32}$	$71.7^{+13.0}_{-8.1}$	0.49/1	50.5
All galaxies, hi-z	1	-1.65 ± 0.08	64.6 ± 3.3	0.99/2	38.9
8, =	2	$-1.51^{+0.16}_{-0.16}$	$58.9^{+6.7}_{-6.7}$	0.62/2	26.8
	3*	$-1.58^{+0.17}$	$61.7^{+5.7}_{-7.3}$	0.64/2	27.2
	3†	$\begin{array}{c} -1.58^{+0.17}_{-0.14} \\ -1.52^{+0.13}_{-0.18} \end{array}$	$59.3^{+8.1}_{-4.9}$	0.60/2	25.9
All galaxies (optimistic), hi-z	1	-1.50 ± 0.17	58.7 ± 6.9	11.75/2	99.4
in galantes (optimistes), in 2 illinois	2	$-1.24^{+0.12}_{-0.12}$	$47.8^{+5.1}_{-5.1}$	13.83/2	99.6
	3*	$-1.33^{+0.11}_{-0.09}$	$51.8^{+4.3}_{-3.1}$	9.16/2	98.7
	3 [†]	$-1.34^{+0.10}_{-0.10}$	$52.2^{+3.9}_{-4.5}$	9.12/2	98.7
Norman et al. (2004), hi-z	1	-1.42 ± 0.10	55.1 ± 4.2	15.25/3	99.8
	2	$-1.17^{+0.09}_{-0.09}$	45.0+3.9	4.10/3	74.9
	3*	$-1.18^{+0.10}_{-0.09}$	$45.0_{-3.9}$ $45.1^{+3.6}$	3.76/3	71.1
	3 [†]	$-1.18^{+0.09}_{-0.10}$	$45.2^{+3.6}_{-4.2}$	3.76/3	71.2

Notes.—Fit parameters are given for $\log \phi = a \log L_X + b$. Method 1: Linear regression. Method 2: Fit to data excluding upper limits. Method 3*: MCMC fit using Method 1 results to initialize the fit. Method 3†: MCMC fit using Method 2 results to initialize the fit (see text). p_{χ^2} gives the χ^2 probability at which the model fit can be rejected (note that χ^2 is computed excluding upper limits).

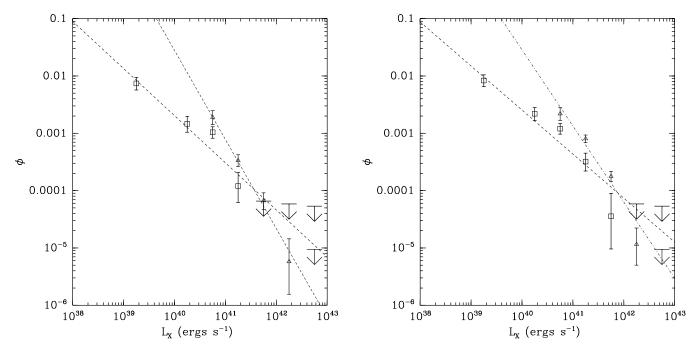


Fig. 10.—Joint linear fits to the low-z and high-z XLFs for the full (early + late type) galaxy sample (*left*) and the optimistic galaxy sample (*right*). The low-z XLF points are marked with squares, and the high-z XLF points are marked with triangles. The dashed lines and dot-dashed lines show the fits to the low-z and high-z XLFs, respectively. [See the electronic edition of the Journal for a color version of this figure.]

galaxy XLFs at low- or high-redshift), explicitly fitting for the offsets in *a* and *b* between the XLFs. Figure 13 shows the early-and late-type XLFs being similarly fitted jointly. Here also the posterior probabilities were mostly Gaussian in shape, and the best-fit parameter values and errors are given in Table 2. These fits resulted in parameters for the given XLFs that were equivalent to those obtained by fitting the XLFs separately with the linear model. However, here we are deriving the probability distribution for the offsets in slope and intercept. As discussed above, this approach avoids the need to propagate errors in determining the sig-

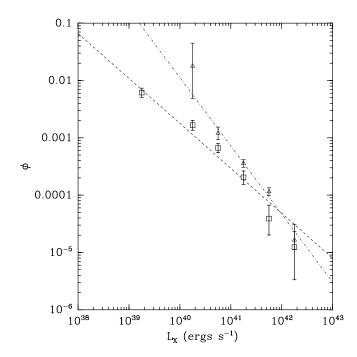


Fig. 11.—As in Fig. 10, but for the N04 XLFs. [See the electronic edition of the Journal for a color version of this figure.]

nificance of the change in the linear model parameters between XLFs. But in these fits, the final errors on Δa and Δb are very similar to adding the errors obtained from the individual fits in quadrature, which is perhaps not surprising, given that the posterior probabilities for these parameters were nearly Gaussian.

In every case, the best-fit values of Δa were <0 and the best-fit values of Δb were >0. In order to estimate the significance of a change in the XLF, we determined the fraction of simulations in which $\Delta a < 0$ and $\Delta b > 0$. Since in some cases this probability is very small, we performed 50 runs with a very large chain length (5 × 10⁶) for these fits. These probabilities are also listed in Table 2.

3.3. Lognormal and Schechter Fits

We fit the XLFs using the lognormal (Saunders et al. 1990) and Schechter (Schechter et al. 1976) functions, since these functional forms fit the FIR (and hence star-forming galaxy) and optical luminosity functions of galaxies well. Specifically, the functional forms were

$$\phi(L) = \phi^* \left(\frac{L}{L^*}\right)^{1-\alpha} \exp\left[-\frac{1}{2\sigma^2} \log^2\left(1 + \frac{L}{L^*}\right)\right] \tag{4}$$

and

$$\phi(L) = \ln(10)\phi^* \left(\frac{L}{L^*}\right)^{1+\alpha} \exp\left(-\frac{L}{L^*}\right),\tag{5}$$

where in both cases the units of ϕ are galaxies Mpc⁻³ $\log L^{-1}$. For both of these sets of fits, we placed very weak priors on $\log L^*$ and $\log \phi^*$. We first fitted the FIR LF published in Saunders et al. (1990) in order to obtain prior information based on the local star-forming galaxy luminosity function and to validate our methodology. We assumed that the errors listed for each LF point were Gaussian and excluded upper limits. Our results are shown in Table 3, along with the original results of Saunders et al. (1990)

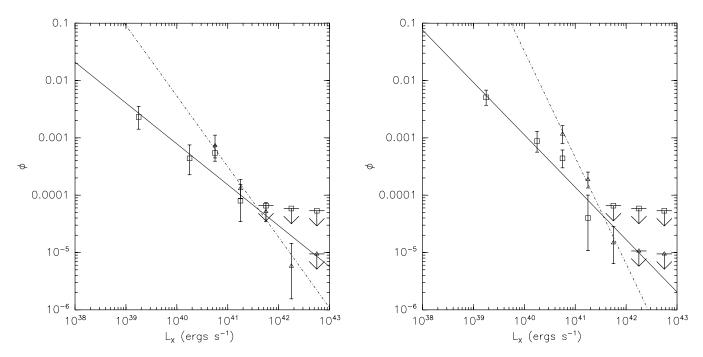


Fig. 12.—As in Fig. 10, but with the left panel showing the early-type galaxy XLFs and the right panel showing the late-type galaxy XLFs. [See the electronic edition of the Journal for a color version of this figure.]

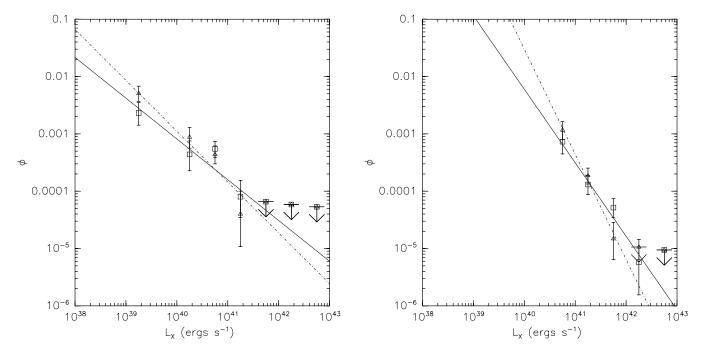


Fig. 13.—Joint linear fits to the low-z early- and late-type XLFs (*left*) and high-z early- and late-type XLFs (*right*). The early-type XLF points are marked with squares, and the late-type XLF points are marked with triangles. The dashed lines and dot-dashed lines show the fits to the early and late-type XLFs. [See the electronic edition of the Journal for a color version of this figure.]

TABLE 2
MCMC JOINT LINEAR FITS

Sample	Δa	Δb	$p_{\Delta a,\Delta b} \ (\%)$	$\chi^2/{ m dof}$	p_{χ^2} (%)
All galaxies	$-0.75^{+0.19}_{-0.15}$	$30.4_{-7.2}^{+6.9}$	>99.9	7.3/4	87.9
All galaxies (optimistic)	$-0.56^{+0.10}_{-0.13}$	$23.5_{-4.5}^{+5.1}$	>99.9	18.9/5	99.8
Norman et al. (2004)	$-0.38^{+0.09}_{-0.11}$	$17.0^{+3.9}_{-4.7}$	>99.9	14.0/7	94.9
Early-type galaxies	$-0.58^{+0.25}_{-0.21}$	$21.6^{+11.4}_{-7.7}$	99.4	4.7/4	68.4
Late-type galaxies	$-1.03^{+0.33}_{-0.23}$	$39.1_{-9.9}^{+12.8}$	>99.9	3.5/3	67.4
Early-/late-type galaxies, low-z	$-0.19_{-0.14}^{+0.15}$	$7.2^{+6.0}_{-5.6}$	89.8	6.6/4	84.4
Early-/late-type galaxies, hi-z	$-0.59_{-0.34}^{+0.32}$	$22.8^{+15.4}_{-11.7}$	96.9	1.6/3	33.2

Notes.—Best-fitting change in slope (Δa) and intercept (Δb) from joint fits to the low- and high-z samples. The "Early-/late-type" samples refer to comparing the early and late-type galaxy samples (i.e., see Fig. 13). $p_{\Delta a,\Delta b}$ refers to the probability that $\Delta a < 0$ and $\Delta b > 0$. p_{χ^2} gives the χ^2 probability at which the model fit can be rejected (note that χ^2 is computed excluding upper limits).

and the results of fitting the lognormal function to a more recent FIR sample in Takeuchi et al. (2003). Our results are consistent with the fitting based on traditional methods within the errors. Since the Takeuchi et al. LF fits were based on more recent data, we set the lognormal fit prior means to their values, with $\log L^*$ scaled to the X-ray band using $\log L_{\rm X} = \log L_{60~\mu \rm m} - 3.65$ (Ranalli et al. 1990) and by $(1 + z)^3$, where z = 0.25 for the z < 0.5 XLFs and z = 0.75 for the 0.5 < z < 1.2 XLFs. In the case of the early-type or late-type galaxy XLFs, we further reduced the prior mean for ϕ^* by a factor of 2, since these samples were $\sim 50\%$ of the total galaxy sample. We conservatively assumed the MCMC errors in Table 3, which were similar to, but larger than, the Takeuchi et al. errors. We set the prior standard deviations to 5 times these errors for α and σ and 50 times these errors for $\log L^*$ and $\log \phi^*$ (effectively weak). In practice, however, different values of α were preferred by the data, and we thus refit the data after setting the prior mean (and initial starting values) to the best-fitting α values from the original fits. The fits were not stable if we allowed the slope prior widths to be broader, and in the case of the earlytype galaxy XLFs, the prior for α was required to be 50% smaller in order to result in stable fits.

The lognormal XLF fits to the XLFs are shown in Figures 14, 15, 16, 17, and 18, with the best-fit parameters and errors given in Table 4. The posterior probability distributions for the fit parameters for the z < 0.5 early-type sample XLF are shown in Figure 19, and the probability distributions for the fit parameters from the other lognormal fits have similar shapes. Note that using a tighter prior for α for the early-type galaxy fits did not change the best-fit parameter values significantly, but resulted in $\sim\!25\%$ smaller errors for $\log L^*$ and $\log \phi^*$. The posterior probability distributions for σ were obviously completely dominated by the prior distribution. This is, nevertheless, an improvement over simply "fixing" fit parameters, which is the common procedure when parameters are not sufficiently constrained by the data, since the width of the priors are scaled from the "physical prior" of the 60 μ m LF fit results. A broad tail in the probability densities toward low

values of log ϕ^* and high values of log L^* is due to degeneracy between these parameters. This also can be seen in Figure 20, where the MCMC draws are plotted.

We computed the luminosity density, $\rho = \int \phi(L)L d \log L$, for each MCMC draw by numerically integrating the lognormal function over the range $10^{37} < L_{\rm X} < 10^{43} {\rm~ergs~s^{-1}}$. The posterior probabilities for ρ are shown in Figure 21. Note that this results in a statistically correct estimation of ρ and its error, since no (usually questionable) propagation of errors is required.

In the case of the Schechter function fits, we assumed $\alpha = -1.0$ with a prior width of 50% (i.e., prior $\mu_{\sigma} = -1.0, \sigma_{\alpha} = 0.5$), since the GOODS-S *J*-band 0.1 < z < 0.5 luminosity function α values were in the range $-1.4 < \alpha < -0.5$ (Dahlen et al. 2005). We also used the early-type galaxy 0.1 < z < 0.5 *J*-band LF fit parameters from Dahlen et al. (2005) for M_I^* (-22.97) and ϕ^* $(8.6 \times 10^{-4} \text{ Mpc}^{-3} \text{ mag}^{-1})$ to estimate the initial XLF fit parameters (and prior mean values) for $\log L^*$ and ϕ^* . We converted M_J to the X-ray band by computing the mean k-corrected value of $\log F_{\rm X} + 0.4J + 5.1$ (see also Appendix C) to be -3.3 for normal galaxies. We rescaled ϕ^* by 2.5 to be in the units of Mpc⁻³ dex⁻¹. As with the lognormal fits, the prior mean for ϕ^* was reduced by a factor of 2 for the early-type and late-type galaxy sample XLFs. The best-fitting Schechter models are also shown in Figures 14– 18, with the best-fitting parameter values and errors given in Table 5. The posterior probability densities for $\log L^*$, $\log \phi^*$, α , and ρ are shown in Figure 22 for the z < 0.5 early-type galaxy sample, and again other Schechter fits resulted in posterior probabilities with roughly similar shapes.

We also fitted the low- and high-z XLFs simultaneously, in this case only allowing the log L^* to vary between the low- and high-z models. This is, by definition, the case of pure luminosity evolution (PLE). The results of these fits are shown in Figures 23, 24, 25, 26, and 27, and the fit parameters are listed in Tables 6 and 7.

TABLE 3 Lognormal Fits to the 60 μm Luminosity Function

Parameter	Saunders et al. (1990)	Takeuchi et al. (2003)	MCMC (This Work)
α	$\begin{aligned} 1.09 &\pm 0.12 \\ 0.724 &\pm 0.031 \\ 0.026 &\pm 0.008 \\ 10^{8.47 \pm 0.23} \end{aligned}$	$\begin{aligned} 1.23 &\pm 0.04 \\ 0.724 &\pm 0.010 \\ 0.026 &\pm 0.003 \\ (4.34 &\pm 0.86) \times 10^8 \end{aligned}$	$\begin{array}{c} 1.04^{+0.07}_{-0.08} \\ 0.751^{+0.014}_{-0.015} \\ 0.026^{+0.003}_{-0.003} \\ 10^{8.39^{+0.15}_{-0.15}} \end{array}$

Note.—The data used in our MCMC fitting of the $60 \mu m$ LF was taken from Saunders et al. (1990), with several upper limit data points excluded.

⁹ Note that the prior is not biasing this result, since in both cases the prior peaks on the *opposite* side of the probability density to the broad tail.

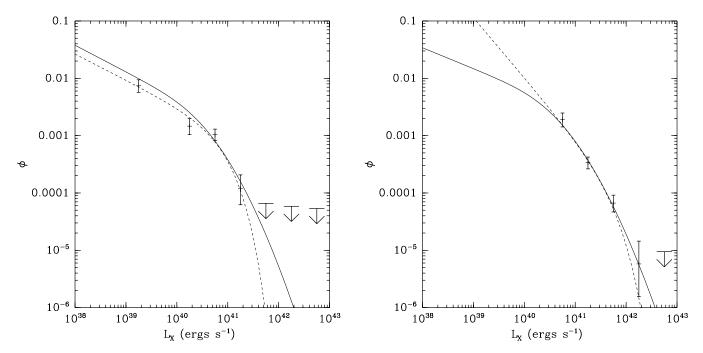


Fig. 14.—Lognormal (solid curves) and Schechter function (dashed curves) fits to the full (early + late type) galaxy sample. The left panel shows the z < 0.5 sample, and the right panel shows the 0.5 < z < 1.2 sample.

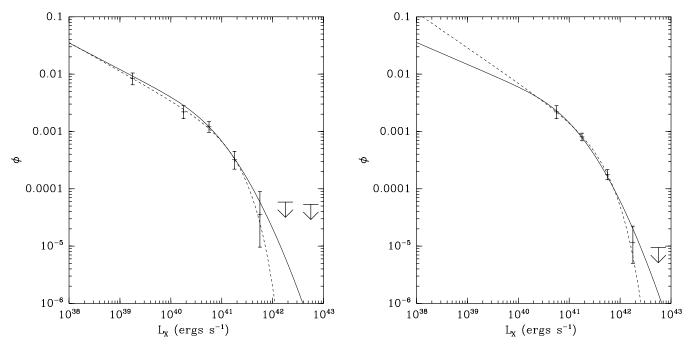


Fig. 15.—Lognormal (solid curves) and Schechter function (dashed curves) fits to the optimistic galaxy sample. The left panel shows the z < 0.5 sample, and the right panel shows the 0.5 < z < 1.2 sample.

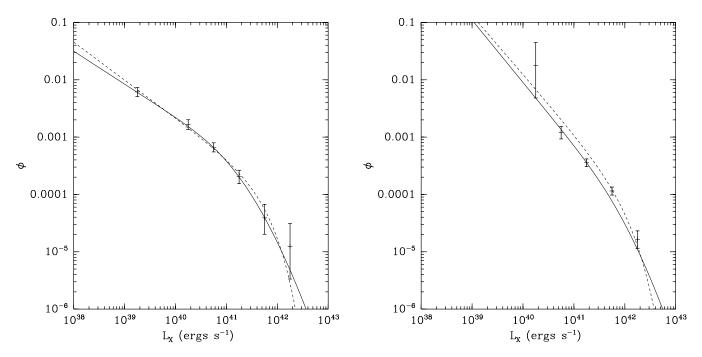


Fig. 16.—Lognormal (solid curves) and Schechter function (dashed curves) fits to the N04 galaxy sample. The left panel shows the z < 0.5 sample, and the right panel shows the 0.5 < z < 1.2 sample.

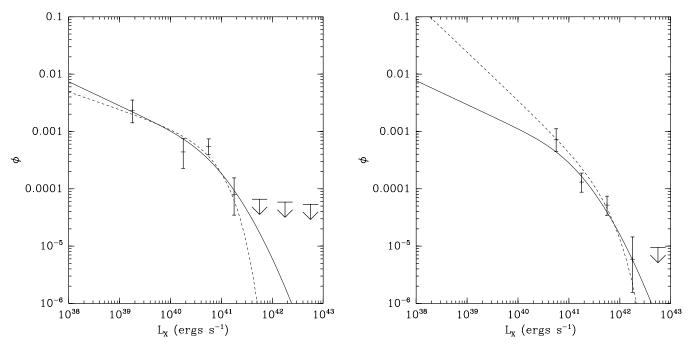


Fig. 17.—Lognormal ($solid\ curves$) and Schechter function ($dashed\ curves$) fits to the early-type galaxy sample. The left panel shows the z<0.5 sample, and the right panel shows the 0.5< z<1.2 sample.

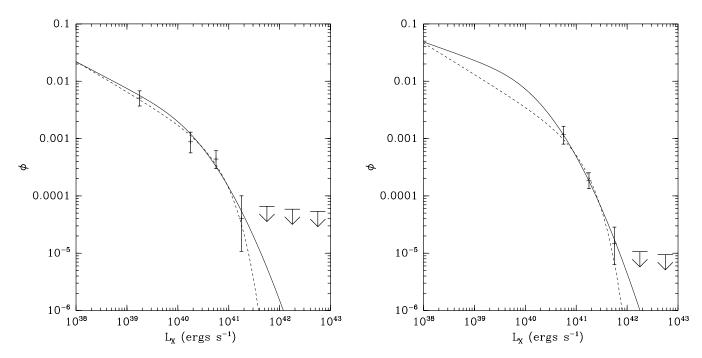


Fig. 18.—Lognormal (solid curves) and Schechter function (dashed curves) fits to the late-type galaxy sample. The left panel shows the z < 0.5 sample, and the right panel shows the 0.5 < z < 1.2 sample.

TABLE 4 MCMC LOGNORMAL FITS

Sample	$\log \phi^*$	$\log L^*$	α	σ	$\log ho$	$\chi^2/{ m dof}$	<i>p</i> _{χ²} (%)
All galaxies, low-z	$-2.28^{+0.18}_{-0.39}$	$39.84^{+0.57}_{-0.43}$	$1.46^{+0.20}_{-0.26}$	$0.71^{+0.08\mathrm{a}}_{-0.07}$	$37.94^{+0.07}_{-0.09}$	8.1	N/A
All galaxies (optimistic), low-z	$-2.46^{+0.31}_{-0.27}$	$40.17^{+0.44}_{-0.49}$	$1.46^{+0.176}_{-0.19}$	$0.73^{+0.08a}_{-0.07}$	$38.11^{+0.07}_{-0.08}$	1.8/1	79.9
Norman et al. (2004), low-z	$-2.88^{+0.34}_{-0.38}$	$40.40^{+0.51}_{-0.40}$	$1.57^{+0.15}_{-0.10}$	$0.74^{+0.08a}_{-0.07}$	$37.96^{+0.08}_{-0.07}$	0.5/2	21.7
Early-type galaxies, low-z	$-3.03^{+0.36}_{-0.30}$	$40.13^{+0.48}_{-0.47}$	$1.42^{+0.14b}_{-0.19}$	$0.74^{+0.06a}_{-0.09}$	$37.62^{+0.13}_{-0.16}$	3.0	N/A
Late-type galaxies, low-z	$-2.41^{+0.34}_{-0.47}$	$39.61^{+0.74}_{-0.67}$	$1.46^{+0.35}_{-0.30}$	$0.73^{+0.07a}_{-0.08}$	$37.63^{+0.12}_{-0.10}$	1.6	N/A
All galaxies, hi-z	$-2.14^{+0.48}_{-0.48}$	$39.85^{+0.68}_{-0.60}$	$1.36^{+0.45a}_{-0.36}$	$0.75^{+0.07a}_{-0.08}$	$38.27^{+0.24}_{-0.17}$	0.9	N/A
All galaxies (optimistic), hi-z	$-2.34^{+0.27}_{-0.31}$	$40.32^{+0.52}_{-0.48}$	$1.38^{+0.37a}_{-0.34}$	$0.71^{+0.08a}_{-0.07}$	$38.44^{+0.14}_{-0.10}$	1.2	N/A
Norman et al. (2004), hi-z	$-3.4^{+1.0}_{-2.2}$	$41.2^{+2.1}_{-1.3}$	$2.09^{+0.16b}_{-0.27}$	$0.73^{+0.08a}_{-0.07}$	$38.22^{+0.56}_{-0.09}$	6.0/1	94.1
Early-type galaxies, hi-z	$-3.13^{+0.54}_{-0.40}$	$40.43^{+0.42}_{-0.60}$	$1.42^{+0.20}_{-0.20}$	$0.73^{+0.08a}_{-0.07}$	$37.83^{+0.16}_{-0.13}$	1.6	N/A
Late-type galaxies, hi-z	$-1.76^{+0.81}_{-0.66}$	$39.43^{+0.57}_{-0.79}$	$1.32^{+0.40}_{-0.37}$	$0.72^{+0.08a}_{-0.06}$	$38.24_{-0.35}^{+0.34}$	0.1	N/A

Note.—Best-fitting parameters from fitting a lognormal function to the XLFs.

^a Parameter is tightly constrained by prior.

Parameter is tignity constrained by prior.

b Parameter is moderately constrained by prior. Luminosities are in ergs s⁻¹ in the 0.5–2.0 keV bandpass. ρ is in ergs s⁻¹ Mpc⁻³ in the 0.5–2.0 keV bandpass; p_{χ^2} gives the χ^2 probability at which the model fit can be rejected (note that χ^2 is computed excluding upper limits).

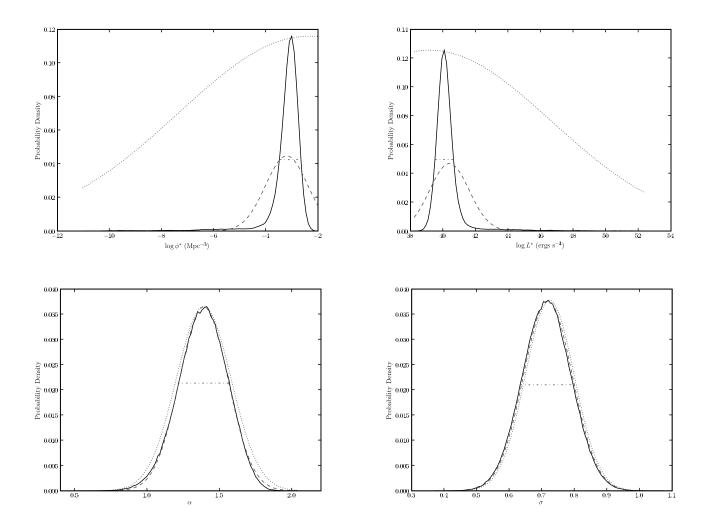


Fig. 19.—Marginalized posterior probabilities for the fit parameters $\log \phi^*$, $\log L^*$, α , and σ for the lognormal fit to the z < 0.5 early-type galaxy XLF. The solid lines show the posterior probability, the dotted lines show the prior, the dashed lines show a Gaussian distribution with the same mean and standard deviation as the posterior, and the dot-dashed line shows the 68% confidence interval. [See the electronic edition of the Journal for a color version of this figure.]

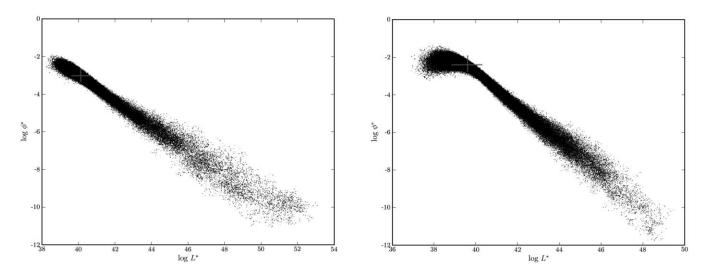


Fig. 20.-z < 0.5 early-type (left) and late-type (right) galaxy MCMC draws for $\log \phi^*$ and $\log L^*$ in the lognormal fit. The solid lines show the 68% confidence intervals determined from the marginalized (and hence one-dimensional) posterior probabilities for the parameters. [See the electronic edition of the Journal for a color version of this figure.]

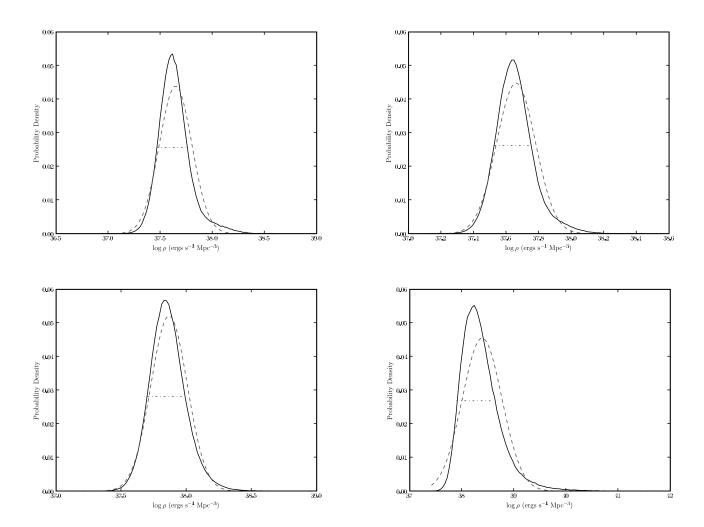


Fig. 21.—Posterior probability distributions for the luminosity density (ρ) derived from lognormal fits to the early-type (*left panels*) and late-type (*right panels*) galaxy samples. The z < 0.5 results are shown in the top row, and the 0.5 < z < 1.2 results are shown in the bottom row. The lines are as shown in Fig. 19 (there is no prior for ρ , since it is derived from other fit parameters). [See the electronic edition of the Journal for a color version of this figure.]

TABLE 5 MCMC SCHECHTER FITS

Sample	$\log \phi^*$	$\log L^*$	α	$\log ho$	$\chi^2/{ m dof}$	$p_{\chi^2} \ (\%)$
All galaxies, low-z	$-3.27^{+0.28}_{-0.26}$	$40.93^{+0.24}_{-0.15}$	$-1.45^{+0.16}_{-0.12}$	$37.90^{+0.08}_{-0.08}$	2.9/1	87.9
All galaxies (optimistic), low-z	$-3.45^{+0.24}_{-0.27}$	$41.28^{+0.23}_{-0.18}$	$-1.50^{+0.11}_{-0.11}$	$38.08^{+0.08}_{-0.07}$	0.4/2	16.9
Norman et al. (2004), low-z	$-4.18^{+0.27}_{-0.41}$	$41.74_{-0.22}^{+0.38}$	$-1.66^{+0.06}_{-0.09}$	$37.99^{+0.09}_{-0.08}$	2.1/3	45.0
Early-type galaxies, low-z	$-3.57^{+0.35}_{-0.36}$	$40.94^{+0.36}_{-0.22}$	$-1.31^{+0.24}_{-0.22}$	$37.55^{+0.13}_{-0.14}$	3.3/1	89.7
Late-type galaxies, low-z	$-3.51^{+0.41}_{-0.57}$	$40.83_{-0.26}^{+0.\overline{43}}$	$-1.52^{+0.18}_{-0.25}$	$37.63^{+0.11}_{-0.12}$	0.9/1	62.9
All galaxies, hi-z	$-4.14^{+0.59}_{-0.85}$	$41.71^{+0.54}_{-0.25}$	$-2.04^{+0.26a}_{-0.35}$	$38.27^{+0.87}_{-0.21}$	1.5/1	75.7
All galaxies (optimistic), hi-z	$-3.54^{+0.25}_{-0.31}$	$41.67_{-0.14}^{+0.20}$	$-1.61^{+0.23}_{-0.25}$	$38.42^{+0.20}_{-0.13}$	0.3/1	43.5
Norman et al. (2004), hi-z	$-4.38^{+0.47}_{-0.94}$	$42.05^{+0.65}_{-0.26}$	$-2.03^{+0.28}_{-0.12}$	$38.27^{+0.41}_{-0.14}$	19.8/2	99.5
Early-type galaxies, hi-z	$-4.31^{+0.53}_{-1.06}$	$41.77_{-0.32}^{+0.74}$	$-1.84^{+0.41a}_{-0.22}$	$37.77^{+0.42}_{-0.20}$	4.2/1	92.2
Late-type galaxies, hi-z	$-3.41^{+0.29}_{-0.37}$	$41.14_{-0.17}^{+0.22}$	$-1.55^{+0.50}_{-0.38}$ a	$37.83^{+0.40}_{-0.20}$	0.9	N/A

Note.—Best-fitting parameters from fitting a Schechter function to the XLFs.

^a Parameter is moderately constrained by prior. Luminosities are in ergs s⁻¹ in the 0.5–2.0 keV bandpass; ρ is in ergs s⁻¹ Mpc⁻³ in the 0.5–2.0 keV bandpass; ρ_{χ^2} gives the χ^2 probability at which the model fit can be rejected (note that χ^2 is computed excluding upper limits).

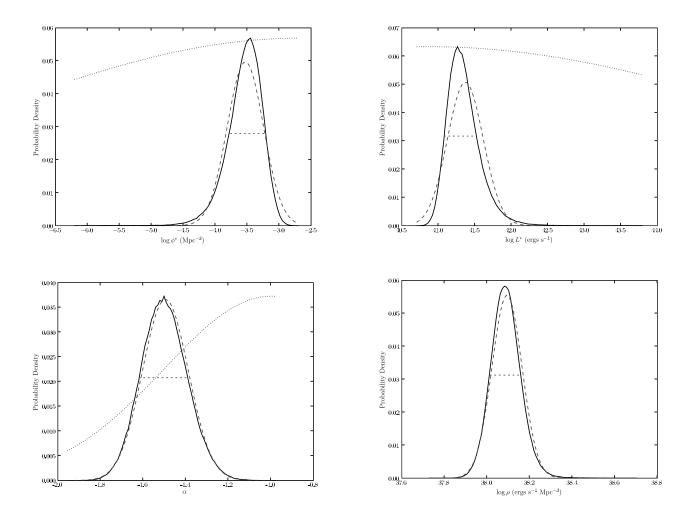


Fig. 22.—Marginalized posterior probabilities for the fit parameters $\log \phi^*$, $\log L^*$, and α for the Schechter function fit to the z < 0.5 early-type galaxy XLF. Also shown is the probability distribution for the luminosity density ρ derived from the fit parameters. Lines are as shown in Fig. 19. [See the electronic edition of the Journal for a color version of this figure.]

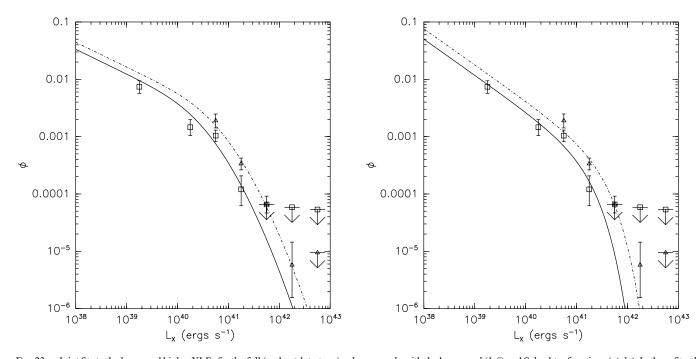


Fig. 23.—Joint fits to the low-z and high-z XLFs for the full (early-+ late-type) galaxy sample with the lognormal (left) and Schechter functions (right). In these fits, the function parameters are tied between the two XLFs and the offset in $\log L^*$ is introduced as an additional fit parameter (i.e., pure luminosity evolution is assumed). The low-z XLF points are marked with squares, and the high-z XLF points are marked with triangles. The dashed lines and dot-dashed lines show the fit to the low-z and high-z XLFs. [See the electronic edition of the Journal for a color version of this figure.]

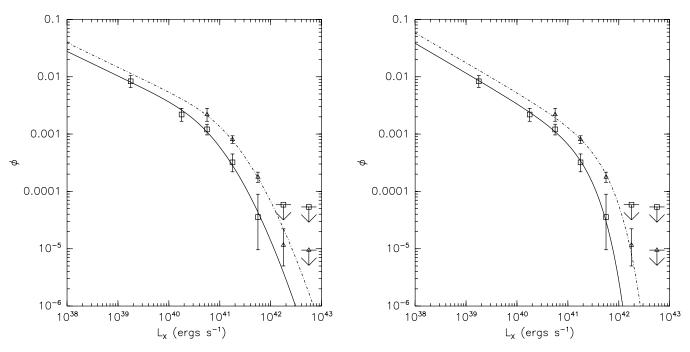


Fig. 24.—As in Fig. 23, but for the optimistic galaxy sample. [See the electronic edition of the Journal for a color version of this figure.]

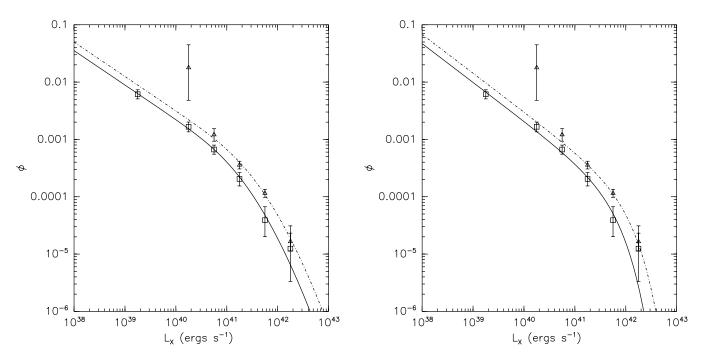


Fig. 25.—As in Fig. 23, but for the N04 sample. [See the electronic edition of the Journal for a color version of this figure.]

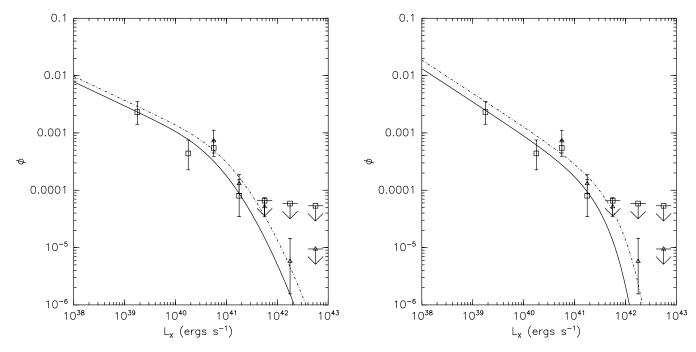


Fig. 26.—As in Fig. 23, but for the early-type galaxy sample. [See the electronic edition of the Journal for a color version of this figure.]

In all cases, the posterior probability for $\Delta \log L^*$ was nearly Gaussian, and we show the cases of the lognormal fits to the early-type and late-type galaxy XLFs in Figure 28.

4. DISCUSSION

We used the GOODS survey to derive X-ray luminosity functions for sources segregated by optical/NIR spectral type. We split the sources into low (z < 0.5) and high (0.5 < z < 1.2) redshift samples in order to investigate evolution. We also explored an "optimistic" sample (when our galaxy selection criterion is relaxed) and the N04 XLFs for comparison. We implemented MCMC techniques for linear, lognormal, and Schechter function fits to the binned XLFs. This has given us a reliable statistical

assessment of the XLFs and any evolution. While in general either a lognormal or Schechter function could fit a given XLF well, this is due in part to the relatively sparse sampling in the XLFs presented here. Better data would be required to constrain the shapes of the XLFs. A consequence of this is that the faint-end slopes of the XLFs are somewhat uncertain, with the lognormal and Schechter function fits giving divergent predictions for the numbers of galaxies expected in deeper exposures.

4.1. Comparison with Local XLFs for Different Spectral Types

We showed that the difference in the early and late-type galaxy XLFs was only significant at the \sim 90% and \sim 97% level for the $z\sim0.25$ and $z\sim0.75$ XLFs. This suggests that there is a difference

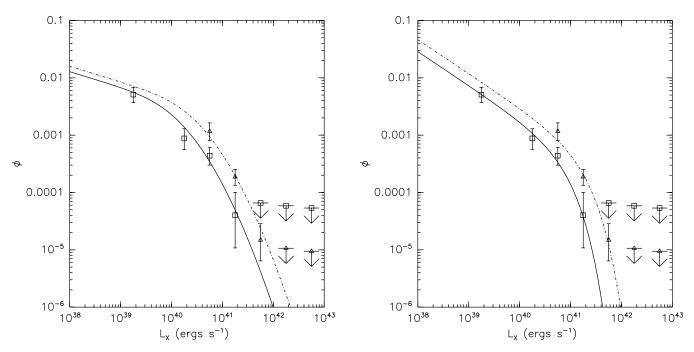


Fig. 27.—As in Fig. 23, but for the late-type galaxy sample. [See the electronic edition of the Journal for a color version of this figure.]

TABLE 6
MCMC JOINT LOGNORMAL FITS

Sample	$\log \phi^*$	$\log L^*$	α	σ	$\Delta \log L^*$	$\Delta \log \rho$	χ^2 /dof	<i>p</i> _{χ²} (%)
All galaxies	$\begin{array}{l} -2.23^{+0.16}_{-0.29} \\ -2.47^{+0.29}_{-0.13} \\ -2.96^{+0.29}_{-0.24} \\ -2.98^{+0.31}_{-0.28} \\ -2.28^{+0.18}_{-0.29} \end{array}$	$\begin{array}{c} 39.74^{+0.48}_{-0.42} \\ 40.14^{+0.30}_{-0.42} \\ 40.52^{+0.32}_{-0.35} \\ 40.08^{+0.53}_{-0.52} \\ 39.43^{+0.50}_{-0.47} \end{array}$	$\begin{array}{c} 1.43^{+0.16b}_{-0.30} \\ 1.43^{+0.13b}_{-0.19} \\ 1.60^{+0.10}_{-0.10} \\ 1.42^{+0.21b}_{-0.27} \\ 1.27^{+0.34b}_{-0.25} \end{array}$	$\begin{array}{c} 0.72^{+0.075}_{-0.07} \\ 0.69^{+0.08a}_{-0.06} \\ 0.74^{+0.08a}_{-0.07} \\ 0.73^{+0.08a}_{-0.07} \\ 0.73^{+0.07a}_{-0.07} \\ 0.69^{+0.08a}_{-0.06} \end{array}$	$\begin{array}{c} 0.28^{+0.09}_{-0.09} \\ 0.35^{+0.08}_{-0.08} \\ 0.25^{+0.08}_{-0.06} \\ 0.23^{+0.15}_{-0.16} \\ 0.34^{+0.12}_{-0.11} \end{array}$	$\begin{array}{c} 0.29_{-0.11}^{+0.09} \\ 0.35_{-0.08}^{+0.08} \\ 0.27_{-0.09}^{+0.07} \\ 0.23_{-0.18}^{+0.15} \\ 0.33_{-0.11}^{+0.14} \end{array}$	8.2/3 2.6/4 2.4/6 5.0/3 3.4/2	95.7 38.0 12.2 82.7 81.4

Note.—Best-fitting parameters from fitting a lognormal function jointly to the low- and high-z XLFs, allowing only $\log L^*$ to vary (i.e., assuming PLE).

between the spectral type-selected XLFs at each redshift, but clearly better data will be required to strengthen this result. Earlyand late-type XLFs were also derived in Georgantopoulos et al. (2005, hereafter G05) for a sample of galaxies at z < 0.2. In Figure 29, we plot our low-z early- and late-type normal/starburst sample XLFs, with the G05 XLFs also shown. There is good overall agreement between the two sets of XLFs, with the exception of the G05 early-type point at $L_{\rm X} \sim 10^{39}~{\rm ergs~s^{-1}}$ being marginally higher than our corresponding point. The mean redshift of our low-z galaxy XLFs is \sim 0.3 for both the early-type and late-type samples, tentatively implying that there has been little or no evolution between $z \sim 0.1$ and ~ 0.3 . However, $\sim 40\%$ of the G05 galaxy sample was comprised of CDF sources, most of which are probably also in our normal/starburst galaxy sample. Having overlapping sources between our GOODS and the G05 samples would obviously dilute any difference between the XLFs.

4.2. Evolution

In all of the samples discussed here, the low-z and high-z XLFs differ at a confidence of >99%, showing that there is statistically significant evolution with redshift. When lognormal and Schechter functions are fit to the individual XLFs, not surprisingly, the resulting ϕ^* and L^* values differ for a given sample; however, the implied luminosity densities ρ do not in general differ significantly. In other words, our computed values of ρ are insensitive to whether the lognormal or Schechter functions are used, giving additional confidence in the resulting values. We also fit the lowand high-redshift XLFs simultaneously with the lognormal and Schechter functions, with the main function parameters (i.e., $\log \phi^*$, $\log L^*$, and the slopes) tied between the two XLFs, but assuming pure luminosity evolution by fitting also for $\Delta \log L^*$. The results for $\Delta \log L^*$ also do not depend strongly on whether the lognormal or Schechter functions are used. The corresponding $\Delta \log \rho$ values were nearly identical to the $\Delta \log L^*$ values, which

might be expected, since $\log \rho$ depends linearly on $\log L^*$ when all other parameters are fixed (and the differences observed between $\Delta \log L^*$ and $\Delta \log \rho$ show the impact of jointly varying $\log \phi^*$ and the slopes).

The early- and late-type galaxy $\Delta \log L^*$ values were 0.23 and 0.34 (from the lognormal fits), showing stronger evolution for late-type galaxies. This corresponds to early- and late-type galaxies being factors of ~ 1.7 and ~ 2.2 brighter, respectively, between $z \sim 0.25$ and $z \sim 0.75$. It is not clear whether the evolution observed in early-type galaxies is due to the passive evolution of the low-mass X-ray binary population (Ghosh & White 2001; Ptak et al. 2001) or enhanced star formation, as expected in the case of late-type galaxies. However, we note that our galaxy type selection is based on galaxy SED type rather than morphological type, and these galaxies therefore have red colors. This suggests that they are not actively star forming. The lack of any redshift dependence in the F_X /NIR flux ratio (see Appendix C) may be constraining to LMXRB evolution models, since LMXRB do not "turn on" instantaneously, as discussed in Ghosh & White (2001); however, the scatter here is large.

Pure luminosity evolution is often expressed as $L^*(z)=(1+z)^pL^*(z=0)$. With that parameterization and $\Delta \log L^*$ being measured between two redshifts z_1 and z_2 , $p=(\Delta \log L^*)/[\log (1+z_2)-\log (1+z_1)]$ or $p\sim 6.8\Delta \log L^*$ for $z_1\sim 0.25$ and $z_2\sim 0.75$. The $\Delta \log L^*$ values then correspond to $p_{\rm early}=1.57^{+1.09}_{-1.03}$ and $p_{\rm late}=2.33^{+0.75}_{-0.82}$. The late-type galaxy evolution is consistent with the FIR evolution of $p\sim 3$. Also note that Georgakakis et al. (2007) similarly found $p\sim 2.4$ for star-forming galaxies from the GOODS-N, using methods somewhat independent of those discussed here (although in both studies, $\log L_X$, X-ray hardness, and X-ray/optical flux ratio were among the selection criteria). The full sample ΔL^* was 0.29, not surprisingly intermediate to the early-type and late-type galaxy XLF values. The optimistic sample resulted in a $\Delta \log L^*$ value of 0.35, basically the same as the late-type

TABLE 7
MCMC JOINT SCHECHTER FITS

Sample	$\log \phi^*$	$\log L^*$	α	$\Delta \log L^*$	$\Delta \log \rho$	χ^2/dof	<i>p</i> _{χ²} (%)
All galaxies	$\begin{array}{l} -3.70^{+0.19}_{-0.22} \\ -3.50^{+0.12}_{-0.17} \\ -4.23^{+0.17}_{-0.15} \\ -4.26^{+0.30}_{-0.31} \\ -3.62^{+0.28}_{-0.24} \end{array}$	$41.24_{-0.14}^{+0.19} \\ 41.32_{-0.10}^{+0.14} \\ 41.32_{-0.10}^{+0.16} \\ 41.75_{-0.11}^{+0.16} \\ 41.47_{-0.23}^{+0.34} \\ 40.91_{-0.17}^{+0.20}$	$\begin{array}{l} -1.63^{+0.10}_{-0.09} \\ -1.52^{+0.07}_{-0.09} \\ -1.68^{+0.06}_{-0.06} \\ -1.58^{+0.14}_{-0.13} \\ -1.59^{+0.16}_{-0.13} \end{array}$	$\begin{array}{c} 0.29^{+0.11}_{-0.09} \\ 0.35^{+0.07}_{-0.09} \\ 0.24^{+0.07}_{-0.08} \\ 0.26^{+0.17}_{-0.18} \\ 0.35^{+0.11}_{-0.12} \end{array}$	$\begin{array}{c} 0.30^{+0.10}_{-0.11} \\ 0.34^{+0.08}_{-0.09} \\ 0.23^{+0.08}_{-0.08} \\ 0.24^{+0.18}_{-0.18} \\ 0.35^{+0.11}_{-0.14} \end{array}$	6.3/4 1.0/5 4.1/7 5.0/4 1.6/3	82.0 4.1 23.4 70.8 35.0

Notes.—Best-fitting parameters from fitting a Schechter function jointly to the low- and high-z XLFs, allowing only $\log L^*$ to vary (i.e., assuming PLE). Luminosities are in ergs s⁻¹ in the 0.5–2.0 keV bandpass; ρ is in ergs s⁻¹ Mpc⁻³ in the 0.5–2.0 keV bandpass; p_{χ^2} gives the χ^2 probability at which the model fit can be rejected (note that χ^2 is computed excluding upper limits).

^a Parameter is tightly constrained by prior.

^b Parameter is moderately constrained by prior. Luminosities are in ergs s⁻¹ in the 0.5–2.0 keV bandpass; ρ is in ergs s⁻¹ Mpc⁻³ in the 0.5–2.0 keV bandpass; p_{χ^2} gives the χ^2 probability at which the model fit can be rejected (note that χ^2 is computed excluding upper limits).

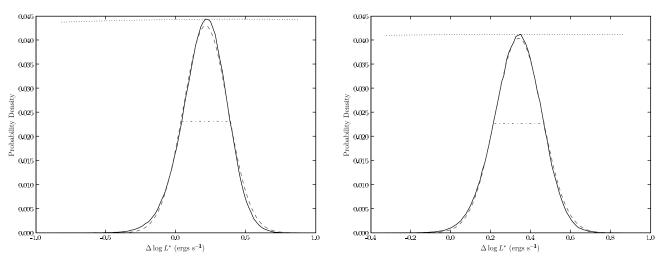


Fig. 28.—Posterior probability distributions for $\Delta \log L^*$ for the early-type (*left*) and late-type (*right*) galaxy samples based on lognormal fits (very similar results were obtained from the Schechter function fits). The solid lines show the posterior probability, the dashed lines shows a Gaussian distribution with the same mean and standard deviation as the posterior, the dotted lines show the prior, and the dot-dashed line shows the 68% confidence interval. [*See the electronic edition of the Journal for a color version of this figure*.]

galaxy value, although this may be somewhat coincidental, since the optimistic sample is most likely also introducing low-luminosity AGNs. However, any AGN activity is probably not dominating the near-IR—optical and the X-ray bandpass, since otherwise the X-ray/optical ratios and/or the X-ray hardnesses would have resulted in an AGN classification (see Appendix A). Nevertheless, the results of our analysis of the optimistic sample probably represent a reasonable limit to the maximum amount of evolution expected for the soft X-ray emission from normal/starburst galaxies between $z \sim 0.25$ and 0.75. We also note that the luminosity densities inferred for the full sample XLFs are very similar to the luminosity densities inferred for the N04 XLFs, while the optimistic sample XLFs result in luminosity densities \sim 0.2 dex higher. Better data (either from deeper *Chandra* exposures or future X-ray missions) would of course result in smaller errors on the X-ray properties of the

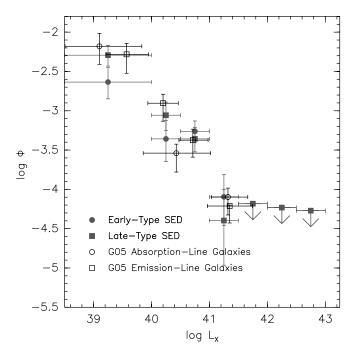


Fig. 29.—Low-z spectral type and galaxy selected XLFs from this work shown along with the XLFs from G05. [See the electronic edition of the Journal for a color version of this figure.]

sources, which would in turn improve the classification probabilities, as well as increasing the number of sources populating the low-luminosity end of the XLFs.

5. SUMMARY

We have computed XLFs for normal/starburst galaxies in GOODS for sources with X-ray counterparts, and fit the XLFs with linear models using "traditional" techniques, as well as Markovchain Monte Carlo techniques. From the photometric redshift fitting procedure, we classified 40 galaxies as early-type and 46 galaxies as late-type based on their SEDs. The early-type galaxy XLFs tend to be slightly flatter than those for late-type galaxies, although from the MCMC analysis, the significance of this result is only at the $1-2 \sigma$ level. The early and late-type galaxy sample XLFs at $z \sim 0.25$ are consistent with the low-redshift early- and latetype XLFs of Georgantopoulos et al. (2005). We used the MCMC approach to also fit the XLFs with lognormal and Schechter functions. The XLFs discussed here all show significant evolution between $z \sim 0.25$ and 0.75. We jointly fit the low- and high-redshift XLFs assuming pure luminosity evolution, allowing only L^* to vary between the XLFs, which resulted in an evolution of $(1+z)^{1.6}$ and $(1+z)^{2.3}$ for the early-type and late-type galaxy samples, respectively. The late-type galaxy evolution derived here is consistent with the star-forming galaxy X-ray evolution given in Georgakakis et al. (2007). Including sources with ambiguous classification results in an "optimistic" galaxy sample with a total galaxy X-ray evolution of $(1+z)^{2.4}$, essentially the same value as derived for late-type galaxies. The optimistic sample XLF evolution suggests that the maximum amount of evolution in the X-ray emission of normal/starburst galaxies at these redshifts is $(1+z)^{2.4\pm0.5}$

The Bayesian fitting approach here could be expanded to include additional uncertainties that might impact this analysis, such as the redshift errors (see, e.g., Dahlen et al. 2005), the uncertainties in the completeness correction, and the uncertainties the X-ray and optical fluxes of the sources. A larger impact on our results would likely result from including radio and FIR data from the GOODS fields, which we will explore in future work. This will improve both the SED fitting (and hence the galaxy type determination) and give an independent star formation rate estimation (see also Georgakakis et al. 2007). It may also be possible to simultaneously fit for the multivariate luminosity functions and

the individual galaxy types, activity types, and redshifts (with any spectroscopic redshifts used as tight priors), at least in an iterative fashion (i.e., where the current luminosity function estimates guide the galaxy type and photometric redshift probabilities). Finally, advanced Bayesian model selection techniques, as discussed in Trotta (2007 and references therein), can be applied here to guide the parameters of future observations by predicting the

ability of future data to prefer a given model and arrive at a given set of constraints.

We thank the anonymous referee for useful comments that improved this paper. A. P. acknowledges the support of NASA grant NNG04GE13G.

APPENDIX A

GOODS X-RAY SAMPLE PROPERTIES AND GALAXY SELECTION

Here we discuss our methodology for classifying the sources, and we also discuss the statistical properties of the sample. Note that the relevant point of our classification is not whether the entire SED is dominated by star formation or an AGN, but which of these is dominating the soft X-ray band.

A1. BAYESIAN SELECTION

N04 selected normal/starburst galaxies from the full CDF-N and CDF-S samples using a Bayesian classification procedure, where priors were constructed from a set of galaxies with well-determined optical types, normal/starburst galaxy, type 1 AGN, and type 2 AGN (hereafter galaxy, AGN1 and AGN2). The product of the prior distributions for a class and the likelihood for the observed parameters for a given source gave the probability that the source was drawn from that class.

Here we follow the same procedure with several improvements. First, sources with small differences between the probabilities in each class should have been considered uncertain, since these conditions occur when the separation of the sources' parameter values and the parent distribution means are small relative to the parameter errors. Here we use the Bayesian "odds ratio," or the ratio of posterior probabilities for the classes being compared. In Bayesian model testing, a model is considered to be favored only when the odds ratio exceeds at least 3 (while odds ratios greater than 10 are preferred). Second, the parameter likelihoods were handled somewhat simplistically, with Gaussian errors assumed on each parameter and a constant error in $\log L_X$ and $\log L_X - \log L_{\rm opt}$ of 0.25. The Gaussian assumption is not correct for hardness ratio errors and $\log L_X$ errors when the number of counts detected for the source is small. However, it turns out that the Gaussian approximation to hardness ratio errors is *conservative* (Park et al. 2006). Here we use the larger of the asymmetric errors on count rate given in Alexander et al. (2003). We assume that $\Delta F_X/F_X = \Delta C/C$, where ΔF_X is the error on X-ray flux and ΔC is the error on X-ray count rate C, given in Alexander et al. We then take the error on $\log L_X$ to be given by $(\Delta L_X)/[L_X \log (10)]$, where $\Delta L_X = L_X(\Delta F_X/F_X)$. This is valid only when $\Delta F_X/F_X \ll 1$; however, this is the case for the majority of the X-ray sources. Finally, N04 did not account for k-corrections in the optical data when computing X-ray/optical flux ratios (k-corrections to the X-ray data were not necessary, since an energy index of 1.0 was assumed for every source). As discussed below, k-corrections are now included in the computation of the priors and in the galaxy classification.

A2. PRIORS

Szokoly et al. (2004) reported classifications based on the optical spectra alone, in the classes "ABS" (no or only absorption lines are present in the spectrum), "LEX" (a low-ionization emission-line spectrum), "HEX" (a high-ionization emission-line spectrum),

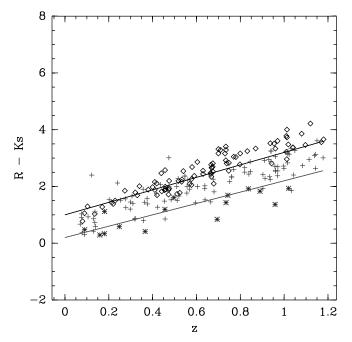


Fig. 30.— $R - K_s$ color plotted as a function of redshift for the early-type, late-type, and irregular/starburst galaxies in the full sample. Also plotted are the lines used to separate these spectral types in the computation of priors in the Szokoly sample. [See the electronic edition of the Journal for a color version of this figure.]

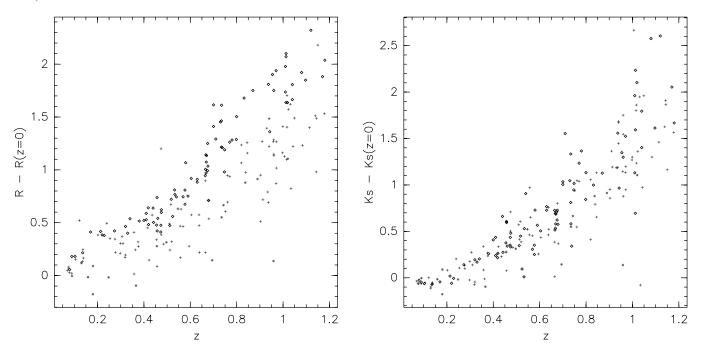


Fig. 31.—K-corrections plotted as a function of redshift for R-band (left) and K_s-band (right) magnitudes. The quantity plotted is the difference between the observed and k-corrected magnitude. Early-type galaxies are plotted with diamonds, late-type galaxies are plotted with plus symbols, and irregular/starbursts are plotted with asterisks. [See the electronic edition of the Journal for a color version of this figure.]

and "BLAGN" (broad emission lines are found). All of the BLAGN sources should correspond to AGN1 sources by definition; however, a broad-line AGN may be present in the other classes where there was not sufficient signal to detect a broad-line component. The HEX and LEX classes should be dominated by type 2 AGN and star-forming galaxies, respectively. However, the LEX classification includes some AGNs where low signal-to-noise or dilution of the nuclear spectrum due to aperture effects (Moran et al. 2002) has precluded the detection of high-excitation emission lines. For several LEX sources, the statistics are sufficient for an AGN component to be identified from line-ratio diagnostics. We therefore derive priors using only ABS sources as galaxies, HEX sources and LEX sources with AGN line ratios as AGN2, and BLAGN sources as AGN1. We required that the X-ray sources have a corresponding entry in the Alexander et al. catalog, since that catalog is used for the X-ray properties of the GOODS sample.

We initially only selected sources with z <= 1.2, but this resulted in only eight AGN2 sources. Since AGN2s are known to have flat X-ray spectra, and hence a minimal k-correction, we relaxed the redshift constraint to z <= 2 for that class. The final tally was then 11 AGN2 sources, 11 AGN1 sources, and 15 (normal/starburst) galaxy sources. The mean offsets between the R- and K-band magnitudes reported in Szokoly et al., and the R and K_s bands used in the GOODS survey were computed in order to adopt the priors based on the Szokoly et al. source for use with GOODS data. This was done regardless of spectral type, and we found offsets of 0.22 mag in R and 1.9 mag in K, with standard deviations of 0.16 and 0.21 mag.

As shown in Figure 3, k-corrections to the X-ray/R- and X-ray/K-band flux ratios can be significant. Since only K- and R-band magnitudes are listed in Szokoly et al., we cannot apply the same interpolation procedure for k-corrections as was done with GOODS data. However, we can use the R-K color to estimate the spectral type of the source and apply a mean k-correction based on that type. We show in Figure 30 the $R-K_s$ colors for the full sample, with early-type, late-type, and irregular/starburst galaxies marked separately. While there is some overlap, we manually selected linear functions to delineate the spectral type as a function of redshift, as shown in the figure. The early-/late-type galaxy separation is given by $R-K_s=1.0+2.2z$, and the late-type/starburst galaxy separation is given by $R-K_s=0.2+2.0z$.

TABLE 8
BAYESIAN PRIOR PARAMETERS

Class	Parameter	Mean	σ
Normal/Starburst Galaxies	HR	-0.19	0.46
	$\log L_{\mathrm{X}}$	40.6	0.7
	$\log F_{\rm X}/F_{\rm opt}$	-3.2	0.7
	$\log F_{\rm X}/F_{\rm K}$	-3.4	0.7
AGN2	HR	0.16	0.37
	$\log L_{ m X}$	41.1	1.1
	$\log F_{\rm X}/F_{\rm opt}$	-2.2	1.0
	$\log F_{\rm X}/F_{K}$	-2.7	0.7
AGN1	HR	-0.51	0.05
	$\log L_{ m X}$	42.9	0.4
	$\log F_{\rm X}/F_{\rm opt}$	-1.2	0.4
	$\log F_{\mathrm{X}}/F_{K}$	-1.4	0.5

TABLE 9 CDF-N X-Ray Sample

XID	Z	Flag	$F_{\rm X}$	Class	log(odds1)	log(odds2)	$\log(L_{\rm X})$	$\log (F_{\rm X}/F_{\it R})$	$\log (F_{\rm X}/F_{\rm K})$	Type	OBXF ID	Z _{G07}
8	1.01	4	0.27	AGN2	38.53	0.90	42.16	-1.97	-2.99	Early		
5	0.64	1	0.12	G 1			41.30	-2.78	-3.02	Late		0.008
5 7	0.13	4	0.20	Gal	14.16	0.51	39.94	-2.72 -3.21	-2.30 -3.33	Late	E (0.28)	0.08
	0.38 0.42	2 4	0.15 0.05	Gal	12.78	1.36	40.87 40.54	-3.21 -2.42	-3.33 -2.52	Late Early	E (0.38)	0.55
	0.42	4	0.03	Gal	18.12	1.21	39.58	-2.42 -4.56	-2.32 -4.44	Early	A (0.09)	
	0.64	1	0.14	Gal	8.68	0.98	41.39	-2.87	-3.30	Late	11 (0.05)	0.64
	0.94	1	0.11				41.68	-2.36	-2.97	Late		
	0.75	1	0.18	AGN2	1.39	1.14	41.66	-1.72	-2.24	Early		0.75
	0.38	4	0.10	Gal	10.92	0.84	40.70	-3.12	-3.28	Early		
	0.68	1	0.14				41.44	-2.58	-2.87	Early		
	0.14	1	0.12	Gal	20.57	1.31	39.76	-3.39	-3.20	Early	A (0.14)	
	1.14	1	0.16	AGN2	13.13	0.84	42.06	-2.09	-2.88	Late		
	0.28	1	5.08	AGN2	1.80	2.12	42.08	-1.70	-1.75	Late	. (0.45)	
1	0.45	1	0.08	Gal	12.60	1.19	40.81	-3.08	-3.30	Early	A (0.45)	0.45 0.81
3	0.97	2	0.28	AGN2	1.51	0.66	42.13	-2.05	-2.62	Late		0.81
5 0	0.33 1.01	4 4	0.03 0.34	Gal AGN2	14.66 0.89	0.96 1.05	40.03 42.26	-3.31 -1.98	-3.41 -2.54	Early Early		
1	0.52	1	0.04	Gal	13.33	1.05	40.64	-3.17	-2.54 -3.57	Late	E (0.52)	0.52
3	0.32	2	1.59	AGN1	4.98	1.13	42.74	-3.17 -1.29	-3.37 -1.90	Late	E (0.32)	0.32
4	0.53	1	0.04	Gal	17.54	0.65	40.63	-3.22	-3.53	Early		0.04
5	0.68	1	4.67	AGN1	5.48	1.43	42.98	-1.21	-1.38	Late		0.68
9	0.47	1	0.08	Gal	12.31	1.14	40.83	-3.11	-3.35	Early	E (0.47)	0.47
0	0.69	1	0.09	Gal	7.17	0.51	41.27	-2.78	-2.75	IR	(,	
6	0.77	4	0.03	Gal	10.68	0.93	40.90	-3.01	-3.60	Late		0.84
2	0.64	1	0.07	Gal	11.36	0.67	41.06	-3.01	-3.26	Late	E (0.65)	0.64
6	0.47	1	0.05	Gal	13.55	1.16	40.65	-3.20	-3.37	Early	A (0.47)	0.47
8	0.48	1	0.07	Gal	10.81	0.80	40.80	-2.92	-3.05	Late	E (0.48)	0.48
9	0.93	4	0.76	AGN1	2.87	0.75	42.52	-1.60	-2.20	Late		1.01
2	0.75	1	0.29	AGN2	7.05	0.97	41.87	-2.02	-2.60	Early		
)	0.63	4	0.13	AGN2	23.99	1.38	41.35	-1.85	-2.26	Early		
3	1.01	2	0.17	AGN2	20.44	0.49	41.96	-2.16	-3.33	Early		
0	0.82	4	0.12				41.57	-2.73	-2.91	Late		
6	0.46	1	0.04	0.1	12.04		40.50	-2.78	-2.63	Late	1 (0.04)	0.46
9	0.31	4	0.10	Gal	13.84	1.10	40.48	-3.08	-3.09	Late	A (0.84)	
0	0.63	4	0.20	AGN2	11.44	0.77	41.53	-2.13	-2.62	Early		1.01
7 0	1.02 0.46	1	0.25 0.23	AGN2 Gal	2.73 9.37	0.66 1.05	42.14 41.25	-2.07 -2.97	-2.99 -3.35	Late		1.01 0.46
7	0.40	1 4	0.23	Gal	14.53	0.49	41.23	-2.80	-3.68	Early Late		0.40
8	1.15	5	0.04	AGN2	4.05	0.49	41.71	-2.80 -1.93	-3.08 -2.47	Late		
9	0.41	1	0.17	AGINZ	•••		41.01	-2.57	-2.77 -2.71	Early		
4	0.56	1	1.62	AGN1	6.01	2.94	42.30	-1.19	-1.41	Late		0.56
7	0.08	1	0.05	Gal	28.79	0.90	38.90	-3.81	-3.43	Late	E (0.08)	0.08
0	0.97	1	0.04				41.31	-2.72	-3.18	Late	_ (****)	
3	1.14	2	0.03	Gal	10.14	0.98	41.27	-3.02	-3.96	Late		
9	0.51	2	0.23				41.36	13	-2.47	Late		0.51
0	0.70	4	0.09	Gal	8.43	0.87	41.30	-2.71	-3.30	Early		
1	0.76	4	0.06	Gal	8.76	0.51	41.23	-2.82	-3.10	Late		0.85
2	0.94	1	0.13				41.77	-2.28	-2.76	Early		
4	1.04	4	0.03	Gal	8.63	0.75	41.27	-2.93	-3.63	Early		
5	1.01	2	0.03				41.17	-2.22	-3.63	Late		
7	0.54	4	0.04	Gal	24.31	0.79	40.63	-3.38	-3.97	Early		
8	0.09	1	0.10	Gal	21.17	0.65	39.31	-3.05	-2.65	Late	E (0.09)	
9	0.70	4	0.03	Gal	10.65	0.85	40.76	-3.02	-3.64	Early		0.00
2	0.77	4	0.59	AGN2	3.17	1.29	42.21	-1.88	-2.60	Early	F (0.56)	0.86
7	0.52	4	0.04	Gal	13.17	1.03	40.57	-3.14	-3.34	Late	E (0.56)	0.56
)	1.01	1	0.08	Gal	9.67	1.11	41.62	-2.99 2.80	-3.80	Early		0.45
1 1	0.45 0.97	1 1	0.09 0.05	Gal Gal	10.56 9.90	0.84 0.91	40.85 41.36	-2.89 -3.06	-2.97 -3.67	Late Late		0.43
1 5	0.32	1	0.03	Gal	15.27	0.91	39.90	-3.00 -3.32	-3.07 -3.09	Late	E (0.32)	0.32
9	0.32	2	0.02	Gai		0.61	41.30	-3.32 -2.53	-3.09 -2.71	Early	L (0.34)	0.32
1	0.47	1	0.24	Gal	22.92	1.31	39.59	-2.59 -3.59	-3.24	Late	E (0.14)	0.14
6	0.60	1	0.09	Jui			41.10	-2.65	-2.72	Late	2 (0.17)	0.17
7	0.09	1	0.04	Gal	30.11	1.08	38.92	-4.09	-3.93	Early	A (0.09)	
8	0.75	1	0.04	Gal	9.97	0.79	40.96	-2.83	-3.34	Late	(0.07)	0.75
0	0.47	1	0.06	Gal	10.82	0.86	40.71	-2.69	-3.13	Late		0.47
2	0.87	4	0.67	AGN1	4.47	1.93	42.39	-1.31	-1.91	Early		0.81
4	0.32	1	0.05				40.25	-2.54	-2.29	Late		0.32
5	0.41	1	0.09	Gal	12.74	1.30	40.74	-3.03	-3.35	Early	C (0.41)	0.41
								- · · · -				

XID	z	Flag	$F_{\rm X}$	Class	log(odds1)	log(odds2)	$\log{(L_{\rm X})}$	$\log \left(F_{\rm X}/F_{\it R} \right)$	$\log (F_X/F_K)$	Type	OBXF ID	$z_{\rm G07}$
269	0.36	2	0.05				40.36	-2.56	-2.42	Late		0.36 ^b
274	0.32	1	0.24	Gal	10.66	1.13	40.91	-3.01	-3.13	Early	A (0.32)	
278	1.02	2	0.03		• • • •		41.24	-2.49	-3.16	Late		
279	0.89	2	0.03		• • •	• • •	41.07	-2.46	-2.92	IR		
280	0.96	4	0.04	Cal	18.94	1.10	41.25	-2.23	-2.90	Early	E (0.20)	0.08 ^{bc}
282 286	0.20 0.95	2 1	0.06 0.41	Gal AGN2	0.71	1.10 1.16	39.87 42.28	-3.42 -1.92	-3.42 -2.61	Late	E (0.20)	0.08 ^b
288	0.93	4	0.41	Gal	8.51	0.88	42.28	-1.92 -2.61	-2.61 -3.54	Early Early		0.93 0.79 ^b
291	0.71	1	0.11	Gal	14.27	1.20	40.60	-3.35	-3.62	Early		0.75
292	0.50	2	0.04	Gai			40.34	-2.27	-2.46	IR		
294	0.47	2	0.31	AGN2	1.29	1.50	41.42	-1.50	-1.93	Late		0.47^{b}
295	0.85	1	0.06				41.31	-2.36	-2.68	Late		
296	0.66	1	0.07	Gal	11.67	0.55	41.12	-2.89	-3.30	Early		
300	0.14	2	0.07	Gal	25.94	1.48	39.54	-3.91	-3.75	Late	E (0.14)	0.14^{b}
304	0.68	1	3.22	AGN1	5.87	2.60	42.81	-1.30	-1.75	Late	. /	
305	0.30	1	0.05	Gal	17.69	1.30	40.15	-3.43	-3.47	Late	E (0.30)	0.30^{b}
309	1.14	2	0.68	AGN1	4.25	1.77	42.70	-1.37	-2.28	Late		1.14 ^b
310	0.76	1	0.03	Gal	11.44	1.02	40.91	-3.05	-3.64	Late		0.76^{b}
311	0.91	1	0.05	Gal	7.15	0.49	41.33	-2.56	-3.18	Late		0.91 ^b
313	0.80	1	0.07	Gal	9.70	0.88	41.30	-3.08	-3.71	Early		
320	0.14	1	0.05	Gal	23.11	1.04	39.40	-3.53	-3.25	Late	E (0.14)	0.14 ^b
323	0.51	1	6.48	AGN2	11.69	3.65	42.82	-1.37	-1.55	Late		
326	0.36	2	0.05		• • •	• • •	40.29	-3.00	-3.24	Late		
332	0.56	1	0.04	Gal	16.42	0.71	40.70	-3.13	-3.54	Late		h
333	0.38	1	0.86	AGN2	0.95	0.97	41.62	-1.96	-2.01	Late		0.38^{b}
337	0.90	1	0.03	Gal	12.22	0.90	41.04	-3.31	-3.79	Late	F (0.05)	0 2 5 h
339	0.25	1	0.05	Gal	22.38	1.14	39.95	-3.82	-3.84	Late	E (0.25)	0.25 ^b
344	0.09	4	15.70	AGN1	4.26	0.99	41.50	-1.06	-0.78	IR		1 00h
346	1.02	2	0.02	Gal	7.37	0.69	40.98	-2.80	-3.76	Late		1.02 ^b 1.25 ^{bc}
349	1.12	4 1	0.53	AGN1	3.71	1.23	42.56	-1.14 -2.64	-2.45	Early		0.94 ^b
351 352	0.94 0.91	4	0.09 0.03	Gal Gal	7.85 17.62	0.50 0.69	41.59 41.13	-2.64 -3.12	-3.53 -3.84	Late Late		0.94
353	0.42	1	0.03	Gal	11.41	1.05	40.80	-3.12 -2.97	-3.06	Late	C (0.42)	0.42 ^b
354	0.57	2	0.10	Gai		1.03	41.05	-2.57 -2.53	-2.70	Early	C (0.42)	0.72
367	0.78	4	0.17	AGN2	2.27	1.19	41.68	-1.65	-2.41	Early		
373	0.48	1	0.15	AGN2	22.40	0.52	41.11	-2.35	-2.68	Early		
378	1.08	2	0.04	110112			41.40	-2.35	-2.98	Late		1.08 ^b
383	0.17	4	0.05	Gal	25.18	1.47	39.63	-4.05	-3.88	Early	A (0.11)	0.11 ^b
384	1.02	2	0.09				41.70	-2.30	-3.39	Early	, ,	
387	0.97	4	0.08				41.59	-2.20	-2.99	Early		1.01 ^{bc}
389	0.56	1	0.05				40.79	-2.52	-3.15	Late		
392	0.40	4	0.05	Gal	14.47	1.13	40.48	-3.24	-3.50	Early	E (0.41)	0.41^{b}
401	0.94	1	0.07	Gal	7.52	0.63	41.52	-2.63	-3.44	Early		
404	0.11	1	0.07	Gal	25.62	1.11	39.27	-3.76	-3.56	Late	A (0.11)	
405	0.94	4	0.05				41.38	-2.80	-3.56	Late		
414	0.80	1	0.09				41.45	-2.46	-2.91	Early		
418	0.28	1	0.09	Gal	17.04	1.26	40.34	-3.50	-3.47	Late		0.28^{b}
426	0.16	2	0.08	AGN2	12.68	0.75	39.75	-2.17	-1.80	IR		
428	0.30	2	0.10	Gal	14.84	1.05	40.47	-3.23	-3.13	Late	E (0.30)	
433	1.04	4	0.08	Gal	7.72	0.69	41.66	-2.70	-3.61	Early		1.02 ^b
437	0.82	4	0.85	AGN1	2.18	0.54	42.44	-1.80	-2.37	Late		0.84 ^b
453	0.73	4	0.38	AGN2	2.17	0.59	41.96	-2.13	-2.74	Early		0.84 ^b
454	0.46	4	0.39				41.48	-2.26	-2.45	Early		0.46^{b}
458	0.07	1	0.31				39.57	-3.03	-2.78	Late	(0.07)	
462	0.51	1	0.03	a .			40.44	-3.21	-3.36	Early		
466	0.44	1	0.09	Gal	13.55	1.16	40.81	-3.15	-3.65	Early		
471	1.17	2	0.10	Gal	7.89	0.64	41.90	-2.71	-3.71	Early		
473	0.31	1	24.90	AGN1	7.71	3.56	42.88	-1.12	-1.25	Late		
477	0.44	1	0.10	A CINTO	4.40	2.94	40.86	-2.67	-2.82	Early		
478	0.08	4	2.27	AGN2	4.40	2.84	40.55	-1.03	-0.82	Early		
480	0.46	1	0.20				41.19	-2.28	-2.77	Early		

Notes.—XID is the ID number from Alexander et al. (2003). z gives the adopted redshift. Flag describes the source and quality of the redshift: 1 for high-quality spectroscopic redshift, 2 for unknown-quality spectroscopic redshift, 3 for low-quality spectroscopic redshift, 4 for high-quality photometric redshift, and 5 for low-quality photometric redshift. F_X gives the 0.5–2.0 keV X-ray flux from Alexander et al. in units of 10^{-15} ergs cm⁻² s⁻¹. Class gives the results of the Bayesian classification, if any. Log(odds1) and log(odds2) give the logarithms of the Bayesian odds ratios for the alternate models (AGN1 and AGN2 in the case of galaxies, galaxies and AGN2 in the case of AGN1, and AGN1 and galaxies in the case of AGN2). L_X is the 0.5–2.0 keV luminosity of the source in erg s⁻¹. Here $\log(F_X/F_R)$ and $\log(F_X/F_K)$ give the X-ray/R-band and X-ray/K-band flux ratios. Type gives the SED type computed in the photometric analysis. OBXF ID gives the spectral type and redshift of the sources found in Hornschemeier et al. (2003). Here z_{G07} gives the redshift from Georgakakis et al. (2007), if present.

^a Source is in the G07 infrared-faint sample.

b Source is in the G07 infrared-bright sample.

^c The G07 redshift is a photometric redshift.

TABLE 10 CDF-S X-ray Sample

XID	Z	Flag	$F_{\rm X}$	Class	log(odds1)	log(odds2)	$\log(L_{\rm X})$	$\log (F_{\rm X}/F_R)$	$\log(F_X/F_K)$	Туре
29	0.57	1	0.24	Gal	7.12	0.81	41.51	-3.25		Early
44	0.57	1	0.60				41.91	-2.96		Early
53	0.67	1	0.14			•••	41.43	-2.57	-2.95	Early
60	0.54	1	0.23	G 1			41.44	-2.53		Late
73	0.42	1	0.10	Gal	13.41	1.22	40.78	-3.62	-3.81	Early
75	1.00	1	0.08	Gal	6.74	0.48	41.61	-2.81	-3.45	Late
80	0.58 0.68	1 1	0.23 0.06	AGN2 Gal	6.17 9.35	0.77 0.76	41.49 41.11	-2.21 -3.00	-2.38 -3.39	Early
83 84	1.03	1	0.00	Gai			41.11	-3.00 -2.30	-3.39 -3.28	Early Late
88	0.60	1	2.48	AGN2	21.15	2.29	42.58	-1.87	-2.12	Late
94	0.12	1	0.73	AGN2	30.05	2.26	40.45	-1.43	-1.69	Late
103	0.68	1	1.76	AGN1	3.64	1.47	42.54	-1.68	-2.06	Early
106	0.67	1	0.27	AGN2	1.67	0.50	41.72	-1.98	-2.45	Early
113	0.52	1	0.09				40.97	-2.45	-2.49	Early
115	0.34	1	0.06	Gal	18.63	1.06	40.33	-3.66	-3.84	Early
117	0.57	1	4.50	AGN1	7.26	3.69	42.77	-1.19	-1.46	Late
118	1.10	1	0.55	AGN2	38.53	1.68	42.56	-1.80	-2.60	Early
121	0.73	1	0.28	Gal	6.78	0.80	41.84	-2.97	-3.38	Early
122	0.18	1	0.36	AGN2	5.20	2.97	40.52	-0.91	-0.98	IR
124	0.96	1	0.08	ACNO	 5 0 4	2.05	41.56	-2.51	-3.06	Late
126	0.59 0.23	4 1	0.33 0.22	AGN2 Gal	5.84 14.37	2.85 1.15	41.67 40.54	-1.17 -3.23	-1.55 -3.19	Early
129 134	0.23	1	0.22	Gai			41.00	-3.23 -2.64	-3.19 -3.13	Early Early
146	0.73	1	0.07				41.33	-2.44	-3.13 -3.26	Early
149	0.13	1	0.23				40.02	-2.73	-2.50	Late
152	0.25	1	0.06	Gal	15.29	0.90	40.06	-3.36	-3.08	IR
155	0.53	1	1.85	AGN1	3.26	1.37	42.32	-1.82	-1.88	Early
158	0.36	1	0.09	Gal	11.29	0.73	40.61	-2.94	-2.92	Late
159	1.05	1	0.91	AGN2	4.15	3.43	42.73	-1.00	-2.01	Late
161	0.73	1	0.31				41.87	-3.09		Late
162	1.04	1	0.08	AGN2	8.43	0.78	41.65	-2.17	-2.54	Late
167	0.58	1	0.09	Gal	11.41	1.00	41.06	-3.28	-3.51	Early
169	1.02	4	0.05				41.43	-3.65		Late
171	1.03	1	0.18	AGN2	0.64	1.56	42.00	-1.83	-1.72	Late
176	0.73	1	0.61	AGN2	6.11	1.71	42.17	-1.79	-2.31	Early
179	0.66 0.74	1 1	0.49 0.25	AGN2	30.04	1.77	41.97 41.79	-1.87 -2.32	-2.13 -2.61	Late
181 182	0.74	1	0.23	Gal	10.83	1.08	40.95	-2.32 -3.13	-2.01 -3.06	Early Early
184	0.62	1	1.32	Gai			42.33	-2.86	-3.00	Early
185	0.18	4	0.06	AGN2	11.38	0.60	39.74	-2.33	-1.98	IR
189	0.08	1	0.81	Gal	21.22	2.08	40.06	-3.60	-3.42	Late
190	1.02	1	0.06				41.48	-2.09	-2.70	Late
192	0.08	1	0.57	Gal	21.21	1.53	39.91	-3.56	-3.43	Late
193	0.96	1	1.81	AGN1	6.23	3.07	42.93	-1.27	-1.96	Late
196	0.67	1	0.07				41.13	-2.74	-2.75	Late
203	1.18	4	0.04	Gal	6.08	0.72	41.54	-3.29	-4.12	Early
207	0.10	1	0.18	Gal	27.50	1.25	39.69	-4.09	-4.02	Early
210	0.83	1	0.07			• • •	41.38	-2.18	-2.71	Early
212	0.67	1	0.17	A CN1	 5.25	1.00	41.53	-2.52	-2.80	Early
214 224	0.84 0.55	1 1	1.51 0.12	AGN1 Gal	5.35 9.46	1.98 0.88	42.71 41.14	-1.22 -2.90	-1.63 -3.31	Late
225	0.55	1	0.12	Gal	11.56	1.05	40.83	-2.90 -3.21	-3.62	Late Late
226	1.00	1	0.12	Gai	11.30		41.79	-2.19	-3.02 -2.71	Late
227	0.67	1	4.01	AGN1	7.49	2.95	42.89	-0.90	-1.19	Early
229	0.67	1	1.86	AGN1	3.40	1.91	42.55	-1.59		Late
236	0.46	1	0.08	Gal	9.70	0.59	40.78	-2.84	-2.91	IR
238	0.24	1	0.06	Gal	17.46	1.25	40.05	-3.42	-3.66	Late
244	0.58	1	0.12				41.21	-2.54	-2.99	Early
247	0.62	2	0.54	AGN2	22.40	0.76	41.95	0.59		Early
248	0.67	4	0.03	Gal	10.45	1.02	40.83	-3.35	-3.74	Early
256	0.74	1	3.66	AGN1	9.56	4.09	42.96	-0.55	-0.88	IR
260	0.27	1	0.51	Gal	7.85	0.95	41.08	-2.68	-2.82	Early
262	0.42	1	0.13	Gal	10.01	0.71	40.89	-2.91	-3.05	Late
263	1.01	1	0.56	AGN2	12.02	2.67	42.48	-1.53	-1.85	Early
265	0.46	1	0.14	Gal	9.69	0.88	41.04	-3.48	• • •	Late

TABLE 10—Continued

XID	Z	Flag	F_{X}	Class	log(odds1)	log(odds2)	$log(L_X)$	$\log (F_{\rm X}/F_{\it R})$	$\log\left(F_{\rm X}/F_K\right)$	Type
266	0.44	1	0.09	Gal	10.56	0.81	40.80	-2.93	-3.13	Late
267	0.10	1	0.11				39.45	-2.96		IR
269	0.66	1	0.22				41.63	-3.82		Early
271	1.18	1	0.12				41.96	-2.32	-3.04	Late
273	1.11	1	0.09				41.79	-2.30	-3.02	Late
276	0.67	1	0.15				41.47	-2.99	-3.43	Early
277	1.15	4	0.78	AGN1	2.04	1.16	42.76	-0.31		Late
292	0.37	1	0.22				41.01	-2.70	-2.42	IR
293	1.14	1	0.18	Gal	8.46	0.49	42.13	-4.43	-2.47	Early
303	0.25	1	2.70				41.71	-1.95		Early

Note.—Columns are as in Table 9.

Having established a crude spectral type color selection, we proceeded by plotting the R and K_s k-corrections as a function of redshift for each spectral type, shown in Figure 31. For each spectral type, we determined the k-correction/redshift correlation. Again, while the dispersion in k-correction is large, this resulted in at least an approximate k-correction that was applied to the sources when computing the priors.

We list in Table 8 the mean and standard deviation for the parameters $L_{\rm X}$, HR (X-ray hardness), X-ray/R-band flux ratio, and X-ray/ $K_{\rm s}$ -band flux ratio computed using the prior samples. The priors were then modeled as Gaussians with these values for the Gaussian mean and standard deviation, as discussed in N04. The full sample is listed in Tables 9 and 10, where we list the final classification along with the Bayesian odds ratios for the models. For example, if the preferred class for a source is "galaxy," then the odds ratios are given for galaxy versus AGN1 and galaxy versus AGN2. For comparison, 30/43 of the OBXF sources from Hornschemeier et al. (2003) are in our GOODS-N sample, and all but one were classified as (normal/starburst) galaxies. The remaining source (XID = 458) had "galaxy" as the highest probability, but only at a factor of 2.1 higher than the AGN2 probability, and hence is only in the "optimistic" sample. A spectral type was given for 28 of the 30 OBXF sources with matches, which is also listed in Table 9. Overall there was good agreement between our SED types and the spectral types, with 7/9 of the absorption-line galaxies having an early-type SED classification and 15/17 emission-line galaxies having a late-type SED classification (the remaining 2 sources were "composite" galaxies with both absorption and emission lines). Similarly, we also list the redshift found in Georgakakis et al. (2007), if present, for the GOODS-N sources, where it can be seen there is good overall agreement between the redshifts.

APPENDIX B

STATISTICAL PROPERTIES OF THE SAMPLES

In Figure 32, we show histograms of the photon indices, binned separately by spectral type. The photon index distributions are similar, peaking at $\Gamma \sim 1.2$ –1.4. However, note that this is due in part to the adoption of $\Gamma = 1.4$ in Alexander et al. 2003 for sources with low signal-to-noise. The number of irregular/starburst sources is very low and is likely to be similar to the rest of the sample. We list in Table 11 the number of sources in each sample (i.e., divided by field, host galaxy type, and bandpass) along with the mean, standard deviation, and minimum and maximum (k-corrected) magnitude. We also give the similar statistics for $\log F_X$ and $\log L_X$.

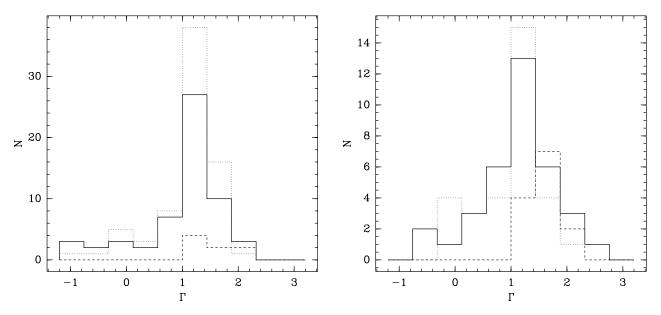


Fig. 32.—Distribution of photon indices for the CDF-N (*left*) and CDF-S (*right*) sources, binned by galaxy spectral type (solid line for early-type, dotted line for late-type, dashed line for starburst/irregular). [See the electronic edition of the Journal for a color version of this figure.]

TABLE 11
X-RAY AND OPTICAL STATISTICAL PROPERTIES

Field	Band	SED Type	N	Mean	σ	Minimum	Maximum
CDFN	В	Early	57	22.26	1.07	19.13	24.27
CDFN	R	Early	57	21.29	1.16	18.03	23.45
CDFN	K_s	Early	57	18.90	0.94	17.09	21.32
CDFN	J	Early	57	20.32	1.02	17.45	22.32
CDFN	$\log F_{\mathrm{X}}$	Early	57	-16.16	0.48	-17.38	-14.64
CDFN	$\log L_{ m X}$	Early	57	40.96	0.70	38.91	42.42
CDFN	B	Late	73	21.94	1.16	19.64	24.66
CDFN	R	Late	73	21.20	1.15	18.59	23.81
CDFN	K_s	Late	73	19.18	0.85	17.62	21.69
CDFN	J	Late	73	20.41	0.95	18.09	22.53
CDFN	$\log F_{\mathrm{X}}$	Late	73	-16.03	0.79	-17.39	-13.64
CDFN	$\log L_{ m X}$	Late	73	41.01	1.03	38.89	43.33
CDFN	B	IR	8	22.11	1.48	19.91	24.54
CDFN	R	IR	8	21.73	1.38	19.84	23.88
CDFN	K_s	IR	8	20.20	1.55	18.27	22.47
CDFN	J	IR	8	21.18	1.39	19.69	23.13
CDFN	$\log F_{\mathrm{X}}$	IR	8	-15.27	1.27	-16.74	-13.81
CDFN	$\log L_{ m X}$	IR	8	41.73	1.52	39.72	43.54
CDFS	B	Early	36	21.62	1.48	18.11	25.95
CDFS	R	Early	36	20.54	2.08	16.25	27.61
CDFS	K_s	Early	29	18.73	1.22	16.07	21.59
CDFS	J	Early	29	19.86	1.38	16.27	22.80
CDFS	$\log F_{\mathrm{X}}$	Early	36	-15.80	0.56	-16.88	-14.51
CDFS	$\log L_{ m X}$	Early	36	41.32	0.65	39.65	42.77
CDFS	B	Late	32	22.03	1.68	17.34	25.95
CDFS	R	Late	32	21.08	1.70	16.71	24.97
CDFS	K_s	Late	26	19.08	1.27	15.94	21.82
CDFS	J	Late	26	20.25	1.50	16.03	22.39
CDFS	$\log F_{\mathrm{X}}$	Late	32	-15.83	0.63	-16.74	-14.39
CDFS	$\log L_{ m X}$	Late	32	41.31	0.83	39.91	42.90
CDFS	B	IR	13	21.45	1.55	19.24	24.65
CDFS	R	IR	13	21.00	1.69	18.00	24.30
CDFS	K_s	IR	7	20.76	1.33	19.58	22.89
CDFS	J	IR	7	21.45	1.30	19.98	23.64
CDFS	$\log F_{\mathrm{X}}$	IR	13	-14.90	1.13	-16.27	-13.36
CDFS	$\log L_{ m X}$	IR	13	41.98	1.78	39.44	44.03

 $\label{eq:table 12} \textbf{X-Ray and Optical Statistical Properties, Galaxies Only}$

Field	Band	SED Type	N	Mean	σ	Minimum	Maximum
CDFN	В	Early	27	21.54	0.91	19.13	22.91
CDFN	R	Early	27	20.50	1.00	18.03	21.88
CDFN	$K_{\scriptscriptstyle S}$	Early	27	18.31	0.61	17.09	19.56
CDFN	J	Early	27	19.60	0.81	17.45	20.75
CDFN	$\log F_{\rm X}$	Early	27	-16.32	0.31	-16.93	-15.60
CDFN	$\log L_{ m X}$	Early	27	40.63	0.66	38.91	41.47
CDFN	B	Late	36	21.45	1.01	19.64	23.65
CDFN	R	Late	36	20.73	1.04	18.59	22.87
CDFN	$K_{\scriptscriptstyle S}$	Late	36	18.95	0.66	17.78	20.59
CDFN	J	Late	36	20.02	0.80	18.09	21.45
CDFN	$\log F_{\rm X}$	Late	36	-16.39	0.32	-17.12	-15.72
CDFN	$\log L_{ m X}$	Late	36	40.41	0.64	38.89	41.41
CDFN	$\log F_{\mathrm{X}}$	IR	1				
CDFN	$\log L_{ m X}$	IR	1				
CDFS	B	Early	13	20.47	1.01	18.11	21.52
CDFS	R	Early	13	19.15	1.38	16.25	20.96
CDFS	K_s	Early	11	17.63	0.76	16.07	18.74
CDFS	J	Early	11	18.58	1.07	16.27	20.09
CDFS	$\log F_{\mathrm{X}}$	Early	13	-15.97	0.46	-16.59	-15.18
CDFS	$\log L_{ m X}$	Early	13	40.92	0.62	39.65	41.85
CDFS	B	Late	10	20.52	1.61	17.34	21.83
CDFS	R	Late	10	19.70	1.58	16.71	21.23
CDFS	$K_{\scriptscriptstyle S}$	Late	9	18.08	1.21	15.94	19.54
CDFS	J	Late	9	18.92	1.56	16.03	20.11
CDFS	$\log F_{\rm X}$	Late	10	-15.94	0.43	-16.36	-15.07
CDFS	$\log L_{ m X}$	Late	10	40.62	0.49	39.91	41.43
CDFS	B	IR	2	21.03	0.79	20.47	21.59
CDFS	R	IR	2	20.63	0.74	20.11	21.15
CDFS	K_s	IR	2	19.66	0.11	19.58	19.74
CDFS	J	IR	2	20.32	0.48	19.98	20.66
CDFS	$\log F_{\mathrm{X}}$	IR	2	-16.24	0.05	-16.27	-16.20
CDFS	$\log L_{ m X}$	IR	2	40.35	0.48	40.01	40.69

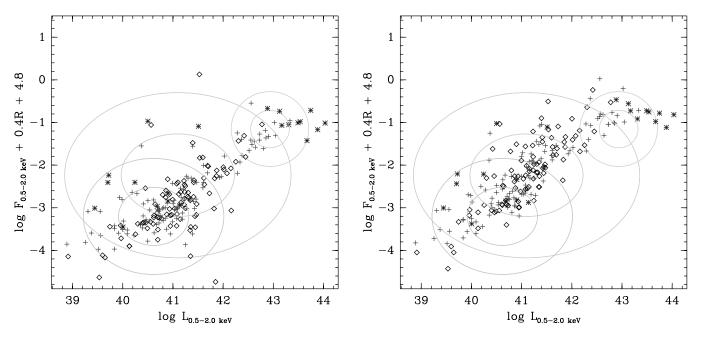


Fig. 33.—X-ray/R-band flux ratio plotted as a function of X-ray luminosity, including (left) and not including (right) k-corrections. Early-type, late-type, and irregular/ starburst SED sources are marked with diamonds, plus symbols, and asterisks, respectively. The ellipses show the 1 σ and 2 σ intervals computed for the prior probability distributions for galaxies, AGN1, and AGN2 for $\log L_X$ and $\log F_X/F_R$. Note that these intervals correspond to standard deviations computed separately for these parameters (i.e., the widths of the ellipses give the 68% and 95% probability intervals for $\log L_X$ for each type), which assumes no correlation between L_X and F_X/F_R . [See the electronic edition of the Journal for a color version of this figure.]

Note that K_s magnitudes were not always available. An immediate conclusion is that the AGN contribution is most significant for the irregular/starburst samples, since the mean X-ray luminosities are $\sim 10^{42}$ ergs s⁻¹, the luminosity at which AGN emission starts to dominate the X-ray band. The corresponding values after galaxy selection are listed in Table 12.

APPENDIX C

X-RAY FLUX RATIOS

Since X-ray/R-band and X-ray/ K_s -band flux ratios are used in selection criteria for this paper, we discuss here the potential impact of any luminosity or redshift dependence of these quantities. In Figures 33 and 34, we plot the X-ray/R-band and X-ray/ K_s -band flux ratios as a function of luminosity, before and after applying k-corrections. Figures 35 and 36 show the flux ratios plotted as a function of redshift. In the luminosity versus flux ratio plots, we show the regions resulting from our Bayesian prior analysis. These ellipses

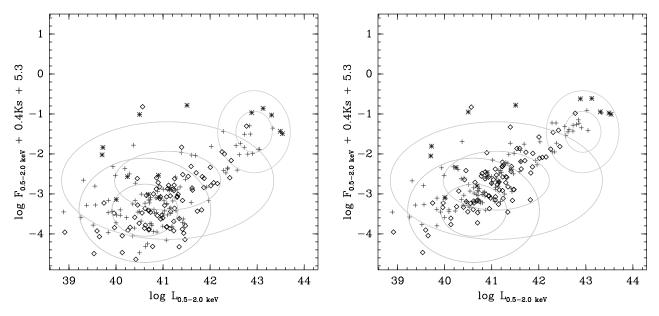


Fig. 34.—X-ray/ K_s -band flux ratio plotted as a function of X-ray luminosity, including (*left*) and not including (*right*) k-corrections. Symbols and ellipses are as in Fig. 33. [See the electronic edition of the Journal for a color version of this figure.]

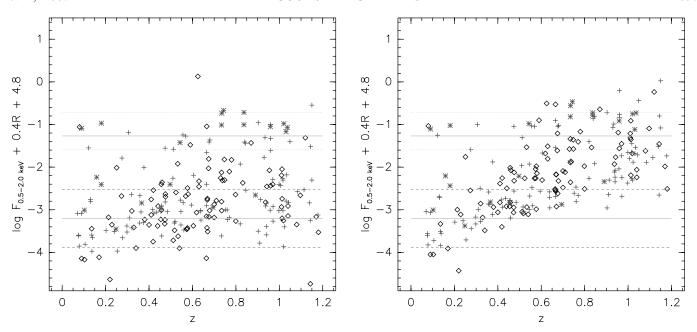


Fig. 35.—X-ray/R-band flux ratio plotted as a function of redshift, including (left) and not including (right) k-corrections. Early-type, late-type, and irregular/starburst SED sources are marked with diamonds, plus symbols, and asterisks, respectively. Horizontal lines are drawn showing the 1 σ probability intervals for the flux ratio for galaxies, AGN1, and AGN2. [See the electronic edition of the Journal for a color version of this figure.]

show the 1 and 2 σ regions where the σ value is based on the standard deviation of the parent populations. In other words, the width and height of the 1 σ regions are the standard deviations of the X-ray luminosity and given flux ratio for the corresponding class (galaxy, AGN1, or AGN2) and correspond to the 68% and 95% probability intervals for that parameter. However, this assumes that the X-ray luminosity and flux ratios are not correlated, and since this is not the case, these regions should *not* be interpreted as joint confidence regions. The intention here is to simply show the regions in the $\log L_{\rm X} - \log F_{\rm X}/F_{\rm R}$ and $\log L_{\rm X} - \log F_{\rm X}/F_{\rm K_s}$ planes from which our normal/starburst samples are being selected.

In the redshift versus flux ratio plots, we show the 1 σ regions for the flux ratios, as in the luminosity/flux ratio plots, except in this case the regions are simply marked with horizontal lines (i.e., since redshift is not a selection criterion). From these plots we conclude that any evolution in the flux ratios with redshift is dominated by the k-corrections and the scatter in the flux ratios. In other words, after k-correction, the flux ratios are consistent with no redshift dependence. This emphasizes the importance of k-correcting the data prior to utilizing flux ratios as a selection criterion. On the other hand, there is clearly a luminosity dependence in the flux ratios. This is due in large part to the increased prevalence of AGN activity at higher X-ray luminosities (e.g., Fiore et al. 2003; Barger et al. 2005).

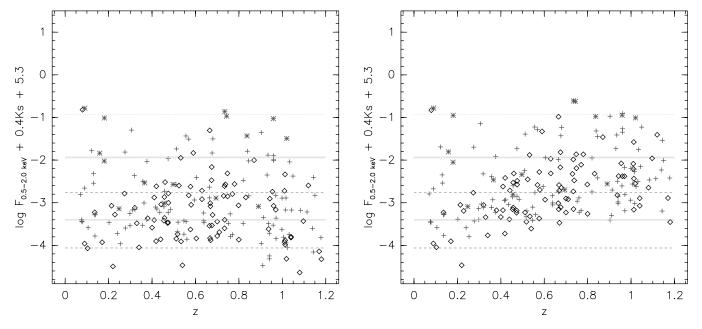


Fig. 36.—X-ray/ K_s -band flux ratio plotted as a function of redshift, including (*left*) and not including (*right*) *k*-corrections. Early-type, late-type, and irregular/starburst SED sources are marked with diamonds, plus symbols, and asterisks, respectively. Horizontal lines are drawn showing the 1 σ probability intervals for the flux ratio for galaxies, AGN1 and AGN2. [See the electronic edition of the Journal for a color version of this figure.]

858 PTAK ET AL.

The correlation between X-ray luminosity and the flux ratios mitigates the benefit of including both luminosity and the ratios as selection criteria. However, in practice there are a significant number of sources in the luminosity range $10^{41}-10^{42}$ ergs s⁻¹, consistent with either AGNs or normal/starburst galaxies, where the X-ray/R-band and/or X-ray/ K_s -band flux ratio is within the 1 σ region for AGN2 galaxies. Thus, including the X-ray/optical and X-ray/ K_s -band flux ratios as a selection criteria should improve the separation, and hence selection, of normal/starburst galaxies from type II AGN.

REFERENCES

Alexander, D. M., et al. 2003, AJ, 126, 539

Barger, A. J., Cowie, L. L., Mushotzky, R. F., Yang, Y., Wang, W., Steffen, A. T., & Capak, P. 2005, AJ, 129, 578

Bauer, F. E., Alexander, D. M., Brandt, W. N., Schneider, D. P., Treister, E., Hornschemeier, A. E., & Garmire, G. P. 2004, AJ, 128, 2048

Benitez, N. 2000, ApJ, 536, 571

Blanton, M. R., et al. 2003, AJ, 125, 2348

Bolzonella, M., Miralles, J.-M., & Pello, R. 2000, A&A, 363, 476

Coleman, G., Wu, C.-C., & Weedman, D. W. 1980, ApJS, 43, 393

Dahlen, T., Mobasher, B., Somerville, R., Moustakas, L., Dickinson, M., Ferguson, H., & Giavalisco, M. 2005, ApJ, 631, 126

Dahlen, T., et al. 2007, ApJ, 654, 172

Fiore, F., et al. 2003, A&A, 409, 79

Ford, E. 2005, AJ, 129, 1706

——. 2006, ApJ, 642, 505

Gehrels, N. 1986, ApJ, 303, 336

Gelman, A., Carlin, J., Stern, H., & Rubin, D. 2004, Bayesian Data Analysis (2nd ed.; New York: Chapman and Hall)

Georgakakis, A., Rowan-Robinson, M., Babbedge, T., & Georgantopoulos, I. 2007, MNRAS, 377, 203

Georgantopoulos, I., Georgakakis, A., & Koulouridis, E. 2005, MNRAS, 360, 782 (G05)

Ghosh, P., & White, N. E. 2001, ApJ, 559, L97

Giavalisco, M., et al. 2003, ApJ, submitted

Ho, L. C., Filippenko, A. V., & Sargent, W. L. W. 2003, ApJ, 583, 159

Hornschemeier, A. E., Brandt, W. N., Alexander, D. M., Bauer, F. E., Garmire, G. P., Schneider, D. P., Bautz, M. W., & Chartas, G. 2002, ApJ, 568, 82 Hornschemeier, A. E., et al. 2003, AJ, 126, 575

——. 2004, ApJ, 600, L147

Isobe, T., Feigelson, E., & Nelson, P. 1986, ApJ, 306, 490

Kinney, A., et al. 1996, ApJ, 467, 38

Kraft, R. P., Burrows, D. N., & Nousek, J. A. 1991, ApJ, 374, 344

Madau, P. 1995, ApJ, 441, 18

Mobasher, B., Rowan-Robinson, M., Georgakakis, A., & Eaton, N. 1996, MNRAS, 282, L7

Mobasher, B., et al. 2004, ApJ, 600, L167

Moran, E. C., Filippenko, A. V., & Chornock, R. 2002, ApJ, 579, L71

Norman, C., et al. 2004, ApJ, 607, 721 (N04)

Page, M. J., & Carrera, F. J. 2000, MNRAS, 311, 433

Park, T., Kashyap, V., Siemiginowska, A., van Dyk, D., Zezas, A., Heinke, C., & Wargelin, B. 2006, ApJ, 652, 610

Persic, M., & Rephaeli, Y. 2003, A&A, 399, 9

Peterson, K., Gallagher, S., Hornschemeier, A., Muno, M., & Bullard, E. 2006, AJ, 131, 133

Ptak, A. 2001, in AIP Conf. Proc. 599, X-ray Astronomy: Stellar Endpoints, AGN, and the Diffuse X-ray Background (New York: AIP), 326

Ptak, A., Griffiths, R., White, N., & Ghosh, P. 2001, ApJ, 559, L91

Ranalli, P., Comastri, A., & Setti, G. 2003, A&A, 399, 39

Russell, P., Ponman, T., & Sanderson, A. 2007, MNRAS, 378, 1217

Saunders, W., et al. 1990, MNRAS, 242, 318

Schechter, P. 1976, ApJ, 203, 297

Spergel, D., et al. 2007, ApJS, 170, 377

Szokoly, G. P., et al. 2004, ApJS, 155, 271

Takeuchi, T., Yoshikawa, K., & Ishii, T. 2003, ApJ, 587, L89

Treister, E., et al. 2004, ApJ, 616, 123

Trotta, R. 2007, MNRAS, 378, 819

van Dyk, D., Connors, A., Kashyap, V., & Siemiginowska, A. 2001, ApJ, 548,

Vanzella, E., et al. 2006, A&A, 454, 423

Wirth, G. D., et al. 2004, AJ, 127, 3121