

# A Data-driven Study of RR Lyrae Near-IR Light Curves: Principal Component Analysis, Robust Fits, and Metallicity Estimates

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# Abstract

RR Lyrae variables are widely used tracers of Galactic halo structure and kinematics, but they can also serve to constrain the distribution of the old stellar population in the Galactic bulge. With the aim of improving their near-infrared photometric characterization, we investigate their near-infrared light curves, as well as the empirical relationships between their light curve and metallicities using machine learning methods. We introduce a new, robust method for the estimation of the light-curve shapes, hence the average magnitudes of RR Lyrae variables in the  $K_{\rm S}$  band, by utilizing the first few principal components (PCs) as basis vectors, obtained from the PC analysis of a training set of light curves. Furthermore, we use the amplitudes of these PCs to predict the light-curve shape of each star in the *J*-band, allowing us to precisely determine their average magnitudes (hence colors), even in cases where only one *J* measurement is available. Finally, we demonstrate that the  $K_{\rm S}$ -band light-curve parameters of RR Lyrae variables, together with the period, allow the estimation of the metallicity of individual stars with an accuracy of ~0.2–0.25 dex, providing valuable chemical information about old stellar populations bearing RR Lyrae variables. The methods presented here can be straightforwardly adopted for other classes of variable stars, bands, or for the estimation of other physical quantities.

*Key words:* methods: data analysis – methods: observational – methods: statistical – stars: variables: RR Lyrae – techniques: photometric

# 1. Introduction

RR Lyrae variables are tracers of the old stellar populations of galaxies, conveniently identifiable due to their characteristic light curves and well-defined mean brightnesses. The nearinfrared (near-IR) light curves of RR Lyrae stars are diminished in amplitude and contain fewer features than at optical wavelengths, making them challenging to discover. This is especially true of first-overtone (RRc) and double-mode (RRd) subtypes. Therefore, in this study, we are only investigating the near-IR properties of fundamental-mode RR Lyrae variables (RRab subtype, from here on RRLs). Nevertheless, the lower absorption in the infrared bands, as well as the slightly lowered effect of metallicity on the absolute magnitudes, makes near-IR observations a tempting option for the determination of the properties of old stellar populations. The reduced dependence on metallicity in the near-IR is, however, still a subject of some debate. In particular, the reduction is not seen in the calculations by Marconi et al. (2015, see their Table 6), whereas it is clearly present in the relations provided by Catelan et al. (2004). Note that in the latter study, evolutionary effects were fully taken into account, using synthetic horizontal branch modeling; indeed, it is shown that one of the main advantages of the near-IR regime is the reduced dependence on evolutionary effects, which become increasingly more important as one moves toward bluer bandpasses.

Generally, to derive distances to these variables, their mean apparent magnitudes in at least two bands have to be determined, allowing the estimation of the line-of-sight extinction toward each star. For classical radial pulsators, such

as Cepheids and RRLs, it is customary to fit the light curves with a truncated Fourier series (see, e.g., Simon & Lee 1981), and to use the intercept of this fit as a measure of the mean apparent magnitude. The lower amplitudes and relatively high scatter of near-IR photometry present a challenge for this technique; in time-series with a limited amount of measurements, there are not enough data to accurately determine the coefficients of the high-order truncated Fourier series needed to describe sharp features, such as the region of the minimum light of the light curves. Alternatively, light-curve templates, such as those of Jones et al. (1996), could be used as model representations of the time-series. However, this approach has drawbacks: the Jones et al. (1996) templates only cover the K band, and the J-band light curves of fundamental-mode RR Lyrae are markedly different; with only four RRL templates, not all possible light-curve shapes are represented.

Although the effect of metallicity on the absolute magnitudes is lessened in the near-IR compared with the optical (e.g., Bono et al. 2003; Catelan et al. 2004), knowledge of individual RRL metallicities can still provide valuable insight into the star formation histories of the oldest populations of the Milky Way in parts not accessible by optical spectroscopy and/or photometry. Relationships between the light-curve shape, period, and the iron abundance [Fe/H], such as the widely used relation of Jurcsik & Kovács (1996), provide a convenient estimate in the optical regime. Despite their obvious usefulness, no such relation has been established so far in the near-IR bands.

The main motivation for the current study is the VISTA Variables in the Vía Láctea (VVV) ESO Public Survey (Minniti et al. 2010), an extensive  $K_{\rm S}$ -band variability survey, also including *YZJH* imaging, conducted with the VIRCAM

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near-IR camera at the VISTA 4.1 m telescope at Paranal Observatory, Chile. As it surveys the most crowded regions of the Milky Way, the Galactic bulge and southern disk, VVV has the capability of uncovering previously uncharted parts of our Galaxy, utilizing different stellar tracers, such as Cepheids and RRLs. For example, RRLs identified by the Optical Gravitational Lensing Experiment variability survey (OGLE; Soszyński et al. 2011) have been combined with VVV photometry to determine the distance of the Galactic bulge and to constrain the spatial distribution of its old component (Dékány et al. 2013). Due to the limiting factor of the diminished amplitudes, searches for new RRLs in the VVV fields have been generally limited to RRab variables. Notwithstanding, directed searches have already led to the discovery of new RRLs (Gran et al. 2015, 2016; Minniti et al. 2017a) in the VVV. These have also led to the discovery of a new Galactic globular cluster by the spatial clustering of RRLs in the VVV disk fields (Minniti et al. 2017b). The need for the reproducible, automatic classification of RRLs has led to the development of a machine-learned classifier for finding RRab variables in the  $K_{\rm S}$ -band (Elorrieta et al. 2016), which enables the discovery of thousands of new RRLs among the hundreds of millions of stellar sources of the VVV survey.

In this article, we introduce several new methods, motivated by the desire of making maximum use of the RRLs in the VVV survey, for the analysis of variable stars in general, and for the near-IR light curves of RRLs in particular. Utilizing a highquality RRL sample (Section 2.1), we apply principal component analysis (PCA) to the  $K_{\rm S}$ -band light curves with the aim of decreasing the parameters required to accurately describe the various light-curve shapes of RRab variables (Section 2). We demonstrate how the J-band light-curve shape can be approximated, utilizing the K<sub>s</sub>-band principal component amplitudes, allowing the determination of the J-band average magnitudes, even from a single observation (Section 2.4). Utilizing the first few principal components as basis vectors, we describe a robust nonlinear fitting technique (Section 3). Finally, we demonstrate, on a selected sample of OGLE-IV RRLs, that the light-curve shapes of RRLs in the  $K_{\rm S}$ band, together with the pulsation period, can be used to estimate their metallicities, similar to their optical light curves (Section 4). The methods developed here are utilized in an accompanying paper (Dékány et al. 2018) for the characterization of the RRL population of the VVV disk fields.

#### 2. Model Representation of Near-IR RRab Light Curves

We have elected to utilize PCA to give a compact representation of the near-IR RRL light curves. PCA is a widely used dimensionality reduction procedure to transform an original set of variables by an orthogonal transformation into a new set of linearly uncorrelated variables called principal components (PCs). Generally, the first few PCs contain most of the variation in the original data. The procedure was first described more than a century ago by Pearson (1901), and then rediscovered and named by Hotelling (1933). PCA has two main uses for data analysis: (1) reducing the number of dimensions of a data set by keeping only the most significant PCs and (2) identifying hidden trends in the data. As astronomical data sets are inherently multidimensional (e.g., images, spectra, individual element abundances of stars, etc.), PCA has been adopted for both purposes by the astronomical community. A non-exhaustive list of PCA applications

includes spectral classification of galaxies (Galaz & de Lapparent 1998), stars (Singh et al. 1998), and quasars (Yip et al. 2004); modeling systematics in light curves (Jordán et al. 2013); removal of galaxies from images with the aim of finding gravitionally lensed background galaxies (Paraficz et al. 2016); analysis of galaxy velocity curves (Kalinova et al. 2017); as well as looking for correlations between diffuse interstellar band features (Ensor et al. 2017).

In the context of variable stars, the most important applications of PCA for the study of Cepheids and RR Lyrae stars were done by Kanbur et al. (2002), Kanbur & Mariani (2004), Tanvir et al. (2005), and Bhardwaj et al. (2017), while Deb & Singh (2009) analyzed a range of variable star classes with the aim of evaluating PCs as a metric for light curves in large databases.

## 2.1. The Light-curve Training Set

We have collected high-quality  $K_{\rm S}$  and J-band photometry available from the literature for 101 RRLs. Table 1 summarizes the training set. The available data can be categorized into three main types: near-IR photometry taken with the aim of Baade-Wesselink analysis of RRLs (Barnes et al. 1992; Cacciari et al. 1992; Fernley et al. 1989, 1990; Jones et al. 1987a, 1987b, 1988, 1992; Liu & Janes 1989; Skillen et al. 1993); serendipitous observations of RRLs in the 2MASS (Szabó et al. 2014) and WFCAM (Ferreira Lopes et al. 2015) calibration fields; and the extensive J and  $K_{\rm S}$  variability study of  $\omega$  Centauri by Navarrete et al. (2015). There are other sources of near-IR timeseries photometry available in the literature for RRLs (see, e.g., Angeloni et al. 2014), but we decided to utilize only photometry where the phase coverage and quality of the observations are adequate for the accurate determination of the near-IR light-curve shapes, at least in the  $K_{\rm S}$  band.

Examining Table 1 reveals that the variables in the field of  $\omega$  Centauri (Navarrete et al. 2015) dominate the sample. Although this could introduce a heavy bias toward a particular metallicity,  $\omega$  Cen contains multiple stellar populations (e.g., Valcarce & Catelan 2011; Gratton et al. 2012, and references therein), with two contributing to the RRL sample (Sollima et al. 2006, and references therein). These two populations have [Fe/H]  $\sim -1.2$  and  $\sim -1.7$ . The variables utilized from the photometry of Ferreira Lopes et al. (2015) are all members of the globular cluster M3, with a metallicity of  $\sim -1.5$  (Carretta et al. 2009). Furthermore, the field RRLs have a wider metallicity distribution, allowing us to cover the physical parameter space of RRLs.

As for the period distribution, the sample covers most of the possible period range of RRLs, from 0.39 to 1.01 days. While very metal-rich RRLs can have periods as short as about 0.35 days, their optical light-curve shapes are not drastically different from those with periods around 0.4 days. Therefore, we can surmise that the present data set can be considered representative of RRLs and their light-curve shapes.

#### 2.2. Data Preparation

To apply PCA to our data outlined in Section 2.1, we have to describe their phase-folded light-curve shape. As mentioned in Section 1, it can be hard to accurately represent the near-IR RRL light curves as a Fourier series. Furthermore, our light curves have vastly different numbers of data points: NSV 660 has almost 3000, while for the RRL found in  $\omega$  Centauri,

 Table 1

 Collection of RR Lyrae near-IR Photometric Observations

						Coll	ection of RI	R Lyrae near-I	R Photo	ometric Observa	ations						
II	) <sup>a</sup>	Period <sup>b</sup>	J <sup>c</sup>	Reference	[Fe/H] <sup>d</sup>	I	D <sup>a</sup>	Period <sup>b</sup>	J <sup>c</sup>	Reference	[Fe/H] <sup>d</sup>	Ι	D <sup>a</sup>	Period <sup>b</sup>	J <sup>c</sup>	Reference	[Fe/H] <sup>d</sup>
AV	Peg	0.390375	+	10	+0.08	$\omega$ Cen	V055	0.581724	+	11		$\omega$ Cen	V041	0.662942	+	11	
V445	Oph	0.397020	+	4	+0.01	$\omega$ Cen	V181	0.588370	+	11		$\omega$ Cen	V013	0.669039	+	11	
W	Crt	0.412012	+	12	-0.45	$\omega$ Cen	V025	0.588500		11		$\omega$ Cen	V114	0.675307	+	11	-1.61
AR	Per	0.425549	+	10	-0.14	$\omega$ Cen	V045	0.589116	+	11		$\omega$ Cen	V149	0.682728	+	11	
SW	And	0.442262	+	1, 9, 10	-0.06	$\omega$ Cen	V125	0.592888	+	11	-1.81	$\omega$ Cen	V046	0.686971	+	11	
RR	Leo	0.452390	+	10	-1.30	$\omega$ Cen	V108	0.594458		11	-1.63	$\omega$ Cen	V088	0.690211	+	11	
Cen	V112	0.474359	+	11		RV	Phe	0.596400	+	2		$\omega$ Cen	V102	0.691396	+	11	-1.65
BB	Pup	0.480532	+	12	-0.35	TT	Lyn	0.597436	+	10	-1.50	$\omega$ Cen	V097	0.691898	+	11	-1.74
Cen	V130	0.493250		11		$\omega$ Cen	V033	0.602324	+	11	-1.58	$\omega$ Cen	V007	0.713026	+	11	
VVSC	054 <sup>e</sup>	0.501267	+	5	-1.50	$\omega$ Cen	V090	0.603404	+	11	-1.78	VY	Ser	0.714094	+	4, 8	
Cen	V074	0.503209	+	11		$\omega$ Cen	V049	0.604650	+	11		$\omega$ Cen	V116	0.720074		11	-1.11
Cen	NV457	0.508619	+	11		UU	Cet	0.606074	+	2		$\omega$ Cen	V034	0.733967	+	11	
Cen	V023	0.510870	+	11	-1.35	$\omega$ Cen	V079	0.608276	+	11		$\omega$ Cen	V172	0.738049		11	
Cen	V107	0.514102	+	11		$\omega$ Cen	V118	0.611618	+	11	-2.04	$\omega$ Cen	V085	0.742758	+	11	
VVSC	055 <sup>e</sup>	0.514810	+	5	-1.50	$\omega$ Cen	V020	0.615559	+	11	-1.52	$\omega$ Cen	V109	0.744098	+	11	-1.70
Cen	V005	0.515274		11	-1.24	$\omega$ Cen	V027	0.615680	+	11	-1.16	$\omega$ Cen	V111	0.762905	+	11	-1.79
VVSC	047 <sup>e</sup>	0.519611	+	5	-1.50	$\omega$ Cen	V062	0.619770	+	11		$\omega$ Cen	V099	0.766181		11	-1.91
Cen	V008	0.521329	+	11		$\omega$ Cen	NV458	0.620326	+	11		$\omega$ Cen	V054	0.772915	+	11	-1.80
VVSC	046 <sup>e</sup>	0.529820	+	5	-1.50	$\omega$ Cen	V032	0.620347		11		$\omega$ Cen	V038	0.779061	+	11	-1.64
Cen	V120	0.548537	+	11	-1.15	$\omega$ Cen	V018	0.621689	+	11		$\omega$ Cen	V026	0.784720	+	11	-1.81
VVSC	050 <sup>e</sup>	0.551535	+	5	-1.50	$\omega$ Cen	V096	0.624527	+	11		$\omega$ Cen	V057	0.794402	+	11	
Cen	V100	0.552745	+	11		SS	Leo	0.626344	+	4	-1.56	$\omega$ Cen	V015	0.810642	+	11	-1.68
R	Cet	0.553038	+	10	-1.29	$\omega$ Cen	V004	0.627320	+	11		$\omega$ Cen	V268	0.812922	+	11	-1.76
U	UMa	0.557648	+	1	-1.15	$\omega$ Cen	V115	0.630469	+	11	-1.64	$\omega$ Cen	V063	0.825943	+	11	
Cen	V067	0.564446		11	-1.19	$\omega$ Cen	V146	0.633092		11		$\omega$ Cen	V128	0.834988	+	11	
Cen	V044	0.567545	+	11	-1.29	$\omega$ Cen	V040	0.634072	+	11	-1.62	$\omega$ Cen	V144	0.835320	+	11	
Cen	V056	0.568023		11		$\omega$ Cen	V122	0.634929	+	11	-1.79	$\omega$ Cen	V003	0.841258	+	11	
W	Dra	0.569670	+	7	-0.95	NSV	660	0.636985	+	13	-1.31	$\omega$ Cen	V411	0.844880	+	11	
Cen	V106	0.569903		11	-1.90	W	Tuc	0.642230	+	2	-1.37	$\omega$ Cen	V104	0.866308	+	11	
V	Oct	0.571163	+	12	-1.08	$\omega$ Cen	V086	0.647844	+	11	-1.99	$\omega$ Cen	V091	0.895225	+	11	-1.81
Cen	V113	0.573375	+	11		X	Ari	0.651180	+	3,6	-2.10	$\omega$ Cen	V150	0.899367	+	11	
Cen	V051	0.574152	+	11	-1.84	$\omega$ Cen	V134	0.652903	+	11		$\omega$ Cen	NV455	0.932517	+	11	
WΥ	Ant	0.574341	+	12	-1.39	$\omega$ Cen	V069	0.653195		11		$\omega$ Cen	V263	1.012158	+	11	-1.73
U Cen	V073	0.575215	+	11		$\omega$ Cen	V132	0.655656		11							

# Notes.

<sup>a</sup> Unique identifier of the variable.

<sup>b</sup> Period in days.

<sup>c</sup> Flag whether the *J*-band data are present and utilized in Section 2.4.

<sup>d</sup> Sources of the iron abundances: ω Cen stars: Sollima et al. (2006); M3 stars: Carretta et al. (2009); NSV 660: Szabó et al. (2014); all other stars: Jurcsik & Kovács (1996).

<sup>e</sup> The variables WVSC 054, 055, 047, 046, and 050 are V83, V116, V108, V55, and V40, respectively, from the globular cluster Messier 3 (Clement et al. 2001; Ferreira Lopes et al. 2015).

References. (Photometric system): 1—Barnes et al. (1992) (CIT), 2—Cacciari et al. (1992) (ESO), 3—Fernley et al. (1989) (AAO), 4—Fernley et al. (1990) (SAAO), 5—Ferreira Lopes et al. (2015) (WFCAM), 6—Jones et al. (1987a) (CIT), 7—Jones et al. (1987b) (CIT), 8—Jones et al. (1988) (CIT), 9—Jones et al. (1992) (CIT), 10—Liu & Janes (1989) (CIT), 11—Navarrete et al. (2015) (VISTA), 12—Skillen et al. (1993) (SAAO), 13—Szabó et al. (2014) (2MASS). For the definitions of the photometric systems, we refer to González-Fernández et al. (2018, VISTA), Hodgkin et al. (2009, WFCAM), Carpenter (2001), and references therein (all other systems).

Navarrete et al. (2015) only acquired a total of 100 and 42 epochs in the  $K_S$  and J bands, respectively.

As an alternative, folded light curves can be described as the linear sum of a series of different basis functions of choice, such as Gaussians, which can be aligned to the phased light-curve points with ordinary least squares (OLS) regression. We have chosen to adopt the Gaussian sum fitting example of Figure 8.4 of Ivezić et al. (2014), to the  $K_S$  and J-band RRL light curves. As RRL light curves are strictly periodic, we have decided to replace the Gaussians with their periodic analog, the circular normal (also called Von Mises) distributions of the form

$$f(x) = \frac{e^{\kappa \cos(x-\mu)}}{2\pi I_0(\kappa)},\tag{1}$$

where  $\mu$  is the measure of location (analogous to the mean of the Gaussian distribution),  $\kappa$  is the measure of concentration (where  $1/\kappa$  is analogous to the variance,  $\sigma^2$  of a Gaussian), and  $I_0(\kappa)$  is the modified Bessel function of order 0. To model the light-curve shapes, we define the sum of 100 of these basis functions, distributed evenly between phases 0 and 1:

$$LC = m + \sum_{i=0}^{99} A_i \frac{e^{\kappa \cos[2\pi (x - \frac{i}{100})]}}{2\pi I_0(\kappa)},$$
(2)

where  $A_i$  are the individual amplitudes of the circular normal distributions, and *m* is the intercept of the fit. Although we could use OLS to find the amplitudes of Equation (2), most of the light curves have less than 100 points available, leading to an underdetermined problem. In such cases, regularization can be introduced to penalize the magnitude of independent parameters (in our case, the amplitudes  $A_i$ ), by modifying the loss function. We do so by utilizing the least absolute shrinkage and selection operator (LASSO or L1 regularization; Tibshirani 1996), which adds the sum of absolute values of the regression coefficients multiplied by a regularization parameter  $\alpha$  to the loss function. We note that by utilizing LASSO, generally most coefficients end up being zero.<sup>6</sup>

We have utilized the linear regression routines of scikit-learn (Pedregosa et al. 2011) to fit the folded  $K_S$  and J light curves of each RRL with Equation (2) utilizing LASSO regularization. Our fit has two hyperparameters:  $\kappa$  and  $\alpha$ . We have utilized both cross validation (leave-one-out for stars with few light-curve points, N-fold otherwise) and manual inspection of the resulting light-curve fits to determine the optimal values of these parameters. High values for the concentration parameter  $\kappa$ grants our model the ability to fit sharp features, such as the rising branches of certain variables, but can cause an overfit in phase ranges with few points. Conversely, at low values of  $\kappa$  the model cannot fit sharp features. We have found that a numerical value of  $\kappa = 6$  is optimal for our model for both the  $K_{\rm S}$  and J-band light curves of RRLs. As for the regularization parameter  $\alpha$ , our cross validation resulted in different optimal values for different stars, typically in the range between  $10^{-4}$  and  $10^{-5}$ . As visual inspection did not reveal significant differences when

changing the regularization parameter between these two values, we have adopted an intermediate value of  $10^{-4.5}$  for all of our stars.

Figure 1 illustrates the quality of the light curves and their fits. During our fitting process, some outlying points have been removed manually and the periods of some variables have been revised, when tension existed between different photometric sources. Table 1 contains the revised periods for all variables. Figure 2 compares our fit with Fourier series of different orders. As can be seen, our method provides a better representation for variables with light-curve gaps.

The utilized light curves had been obtained in a variety of photometric systems, as detailed in Table 1. Therefore, the resulting light-curve fits were all transformed to the photometric system of VISTA. For variables not in the VISTA, WFCAM, or 2MASS systems, first they were transformed to the system of 2MASS utilizing the updated transformation formulae<sup>7</sup> of Carpenter (2001). Then, the 2MASS and WFCAM photometry was transformed to the VISTA system with the help of the CASUVERS 1.4 transformations given by the Cambridge Astronomical Survey Unit (CASU).<sup>8</sup>

# 2.3. Application of PCA on the K<sub>s</sub>-band Light Curves

To apply PCA, we sample our light-curve fits on a grid of 100 phase points, evenly distributed phases from 0.0 to 0.99. In the analysis of pulsating variables, it is customary to set the light-curve maxima at phase 0.0; however, inspecting the  $K_{\rm S}$ -band light curve examples of Figure 1 reveals that the maxima of RRLs in the  $K_S$  band are ill defined: the timing of the maxima of the light curve depends heavily on the strength of the bump on the rising branch. Therefore, we have chosen to align our light curves by the much sharper feature of the lightcurve minima (similar to the case of eclipsing binaries). Furthermore, in PCA, the sample values (the magnitudes in our case) are usually normalized to a mean of 0 and scatter of 1 along each dimension (phase). As our goal is to describe the light-curve shapes of RRLs in the  $K_{\rm S}$  band as a linear combination of PCs, we have chosen to normalize each light curve independently to have a mean of 0 and scatter of 1. These aligned, normalized input light curves can be seen in Figure 3.

We carry out PCA by utilizing singular value decomposition, adopted from the PCA module of scikit-learn (Pedregosa et al. 2011). Figure 4 shows the first six PCs, according to our decomposition. As we have chosen not to normalize in each phase point, the first PC contains the average light-curve shape of the normalized light curves. Including further components to describe the light curves modifies this average shape, and this can be easily understood in the context of the individual light curves, for PCs of low order: the second component can make the bump at the end of the rising branch (around phase 0.15) more or less pronounced, while the third component is important to reproduce the double-peaked light curves displayed by some of the variables.

<sup>&</sup>lt;sup>6</sup> Another popular choice for these kinds of problems is the Tikhonov regularization, also known as L2 regularization or Ridge regression, where the sum of the squares of the fit coefficients times  $\alpha$  is added to the loss function. In contrast to LASSO, Ridge regression does not result in sparse solutions (i.e., most parameters do not end up being zero), making the interpretation of the results of the fit harder to interpret.

<sup>&</sup>lt;sup>7</sup> http://www.astro.caltech.edu/~jmc/2mass/v3/transformations/

<sup>&</sup>lt;sup>8</sup> http://casu.ast.cam.ac.uk/surveys-projects/vista/technical/photometric-

properties. The transformations between the VISTA and WFCAM systems were carried out using the relations updated on 2014 July 30. Very recently, updated transformations were provided by González-Fernández et al. (2018). We have checked that the changes in the resulting magnitudes are less than 0.005 mag. In addition, because this only affects the five stars from Ferreira Lopes et al. (2015), none of the results in this paper are significantly affected by the choice of transformation equations between these two systems.

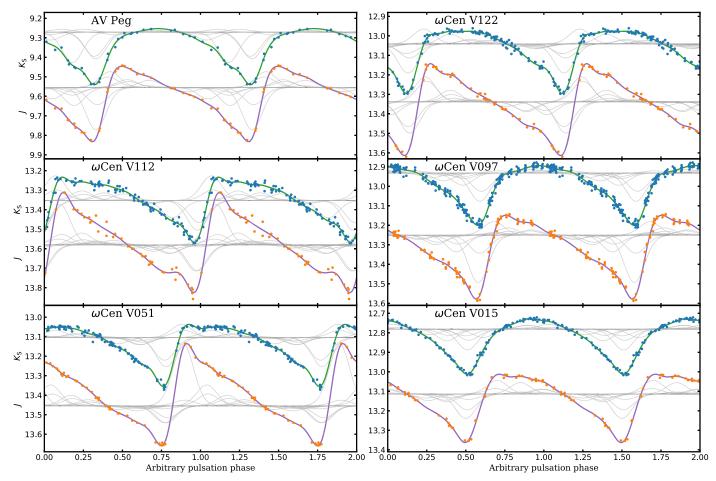


Figure 1. Typical folded  $K_S$  (top, blue points) and J-band (bottom, orange points) light curves of our RRL sample. Each light curve is modeled with a sum of periodic Gaussian (von Mises) basis functions, following Equation (2). As we utilize LASSO regularization, most periodic Gaussians have zero amplitudes. The individual periodic Gaussians with non-zero amplitudes are illustrated by the faint gray lines for each band and variable, while the green and purple curves illustrate the light curve fit (sum of the individual periodic Gaussians) in the  $K_S$  and J-band, respectively.

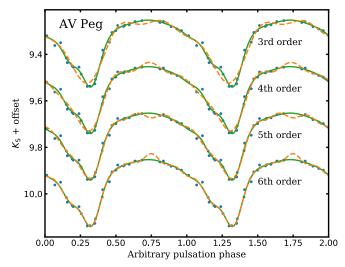
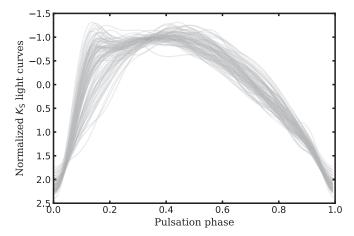


Figure 2. Comparison between our fit with the circular normal distributions (Equation (2), continuous lines), and Fourier series of different orders (dashed lines).

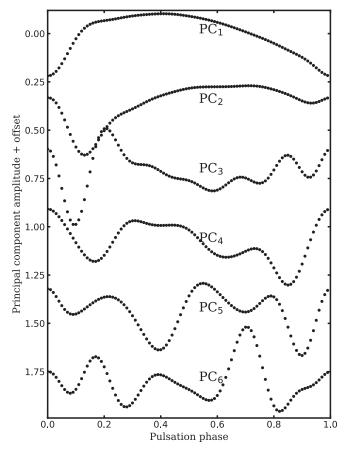
The power of the PCA lies in the fact that generally the first few PCs, together with their amplitudes, are sufficient to describe the original input data, and the rest of the PCs can be discarded. The number of significant PCs can be decided by examining the fraction of the variance explained by the components. As is obvious, the first PC dominates the explained variance, because we normalized each light curve individually instead of normalizing the magnitudes along each dimension (phase). If we look at the variance explained by the rest of the PCs, they explain 3.68, 1.39, 0.96, 0.72, etc. percent of the total variance, or 88.8, 3.4, 2.3, 1.7, etc. percent of the residual variance, when the variance explained by the first PC is subtracted from the total variance. On the basis of these values, we deem the first four PCs are sufficient for describing the  $K_{\rm S}$ -band light-curve shapes of RRab stars.

We applied the PCA method on the normalized light curves of our sample of variables to emphasize the light-curve shape differences among RRLs in the  $K_S$  band, instead of the different pulsation amplitudes of each variable. Consequently, the individual amplitudes (also called eigenvalues) of the first four PCs,  $u_{1j}$ ,  $u_{2j}$ ,  $u_{3j}$  and  $u_{4j}$ , where j = 1..101 is the index of the variables, carry no amplitude information on the original light curves. By multiplying these amplitudes with the normalization constants used to normalize each of the light curves, or by utilizing OLS to directly align the PCs to the light curves, we can determine the PC amplitudes  $U_{1j}$ ,  $U_{2j}$ ,  $U_{3j}$  and  $U_{4j}$  for each star in our sample. The sum of the PCs multiplied with these amplitudes recover the original light-curve shapes (and amplitudes) with high accuracy.

The distribution of the amplitudes  $U_{ij}$  is illustrated in Figure 5. The shapes of the light curves of RRLs, represented

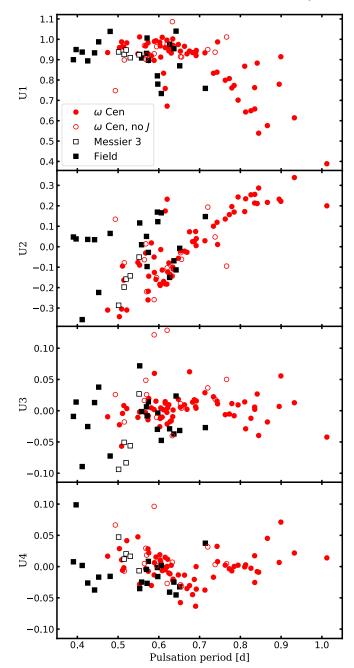


**Figure 3.** Normalized, folded, and minima-aligned  $K_S$ -band light curves of our RRL sample.



**Figure 4.** From top to bottom: PCs 1 through 6, according to our decomposition of the  $K_{\rm S}$ -band RRL light curves.

by the amplitudes of these first few PCs, potentially contain information on the physical properties of the variables themselves, when the pulsation periods of the stars are taken into account. The first amplitude,  $U_1$  (from now on, we omit the *j* index for simplicity), can be viewed as an analog to the total amplitude in amplitude-period (also called Bailey) diagrams. As described before, additional PCs modify the light-curve shape. Keeping this in mind, it is immediately obvious that, at a given period, the amplitudes  $U_1$  and  $U_2$ separate two sequences, which we can associate with the Oosterhoff type I and II groups of variables (Oo;

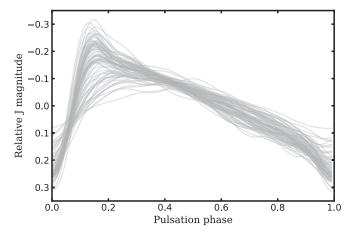


**Figure 5.** Amplitudes  $U_i$  of the PCs for the training sample. Circles and squares denote RRLs from the  $\omega$  Centauri photometry of Navarrete et al. (2015) and other sources, respectively. Empty squares denote RRLs in Messier 3 from the photometry of Ferreira Lopes et al. (2015), while empty circles denote variables not considered for the approximation of the *J*-band light-curve shapes in Section 2.4.

Oosterhoff 1939; see Catelan & Smith 2015, for a recent review and references). As these groups at least partly correlate with metallicity, we are going to explore the possibility of estimating the metallicity of individual RRLs when their  $K_{\rm S}$ -band light curves are described by the PC amplitudes  $U_i$ , in Section 4.

# 2.4. Approximation of the J-band Light-curve Shape

The light-curve shapes and amplitudes of pulsating variables vary among bandpasses, due to the difference in the relative contribution of the change of radius and photospheric



**Figure 6.** Phase- and average-aligned *J*-band light curves of the training sample utilized for the approximation of the *J*-band light-curve shapes. Note that these light curves are aligned in phase by the  $K_S$ -band light-curve minima. Only data for stars with high-quality *J*-band light curves, marked with a "+" sign in column 3 of Table 1, are shown.

temperature during the pulsation cycle to the emitted flux at different wavelengths (Catelan & Smith 2015, and references therein). Comparing the  $K_S$  and J-band light curves in Figure 1 clearly demonstrates this.

During the course of the VVV survey, each field has been observed only a few times in the *J*-band. To determine accurate mean *J* magnitudes for the RRLs, it is necessary to describe the light-curve shapes of the stars in the *J*-band, i.e., the deviation of the *J*-band magnitude from its average in each phase of the pulsation cycle. Usually, the difference between the light-curve shapes is ignored, and the  $K_{\rm S}$ -band light-curve shape is used to estimate the average *J*-band magnitude. However, this method, depending on the light-curve phases of the observations, introduces additional scatter in the derived magnitudes.

As the PC amplitudes  $U_i$  provide a concise description of the  $K_{\rm S}$ -band light-curve shapes, we can evaluate whether the *J*-band light-curve shapes can be approximated with their use. Besides these amplitudes, we consider the period as a possible additional parameter, to assess its effect on variables that otherwise possess the same light-curve shapes in the  $K_{\rm S}$  band.

The  $\omega$  Centauri photometry of Navarrete et al. (2015) contains only 42 epochs in the *J*-band; therefore, depending on their periods, some variables have gaps in their *J*-band light curves at critical phases. A few other stars possess photometric anomalies in their light curves, preventing us from obtaining a good light-curve fit for them in this band. We decided to omit these stars, marked in Table 1, and continue with the analysis of the remaining 87 RRLs. These *J* light curves are first sampled in the same 100 phase points, as has been done for the  $K_{\rm S}$ -band light curves in Section 2.3, where phase 0.0 is the phase of the  $K_{\rm S}$ -band lightcurve minimum for each of the variables.<sup>9</sup> These curves are aligned to have a mean of 0, but in contrast to the PCA analysis of the  $K_{\rm S}$ -band light curves, they are not normalized.

The training set of *J*-band light curves is shown in Figure 6. We are searching for the ideal set of variables to describe these light-curve shapes, e.g., the deviations from the mean in each phase.

We can consider this as a separate problem at each light-curve phase for which we are looking for a linear model description, as a combination of yet to be determined parameters as the basis. The connection between the *J* magnitudes in different light-curve phases is that these linear models should depend on the same parameters. Our candidate parameters are the period *P*, the four PC amplitudes of the  $K_{\rm S}$ -band light curves  $U_1$ ,  $U_2$ ,  $U_3$ , and  $U_4$ , as well as their polynomial combinations up to the third order  $(P^2, PU_1, PU_2, PU_3, PU_4, U_1^2, \text{ etc.})$ .

We performed an exhaustive search for the best parameter combination, where we considered various permutations containing up to six candidate parameters. As our sample size is fairly small, we have chosen not to utilize a hold-out set, nor N-Fold Cross Validation for the validation of our results, as the exact (random) choice of the hold-out set or the N-Folds could lead to heavily biased results (for example, if all the short period variables are selected to be in one fold). Therefore, for each candidate parameter set, we have opted to perform Leave-One-Out Cross Validation (LOOCV) the following way:

- 1. we evaluate 87 separate sub-cases, where we omit the light curve of each of the 87 variables;
- 2. in each case, we optimize a linear solution using OLS for the *J*-band magnitudes, using the candidate parameter set as basis;
- 3. with these solutions, we approximate the *J*-band light curve of the omitted variable, and calculate the residuals; and
- 4. for each candidate parameters set, we assess the total root mean square error (RMSE) as a performance metric.

We have found that out of all the possible parameter combinations, by only considering the PC amplitudes of the  $K_S$ -band light curves  $U_1$ ,  $U_2$ ,  $U_3$ , and  $U_4$ , we reach a total RMSE of our linear prediction of just 0.012 mag. Although this value can still be lowered by including combinations with the period, such as  $PU_1$ , we noticed that this mainly decreases the errors for variables coming from  $\omega$  Centauri, where multiple stars have very similar periods and  $K_S$  and J light-curve shapes, while the errors for other RRLs, especially the short period ones, increase drastically. Figure 7 illustrates the residuals both as a function of the pulsation phase (top panel) and for individual stars from the LOOCV (bottom panel).

Based on our analysis, we conclude that the *J*-band light-curve shapes of RRLs can be estimated using their  $K_{\rm S}$ -band PCA amplitudes with high precision. Figure 8 illustrates this process by comparing the  $K_{\rm S}$ -band PCs derived in Section 2.3 (left-hand panel) with the coefficients found by OLS for the prediction of the *J*-band light-curve shape (right-hand panel), when all of the 87 RRLs with good *J* light curves are considered.

#### 3. Robust Fitting of RRL K<sub>s</sub>-band Light Curves

Our goal is to provide an accurate and convenient method for determining the light-curve shapes, periods and average magnitudes of RRLs in the VVV survey. As we have seen in Section 2, the PC amplitudes  $U_i$  provide a compact description of the  $K_S$ -band light-curve shape; furthermore, the *J*-band light-curve shapes can also be approximated using the same  $U_i$  coefficients as obtained from the  $K_S$ -band data.

The PCs themselves have been utilized before by Tanvir et al. (2005) in the case of Cepheids, to produce a series of realistic template light curves in the optical, which could be used to determine specific parameters, such as the periods or

<sup>&</sup>lt;sup>9</sup> We could reach an incorrect solution if the photometry presented in Table 1 would have  $K_s$  and J-band photometries combined from different epochs, as an incorrect period or (slow) period change could affect the phases. However, as all stars have simultaneous photometry in these two bands, this is not a problem in our case.

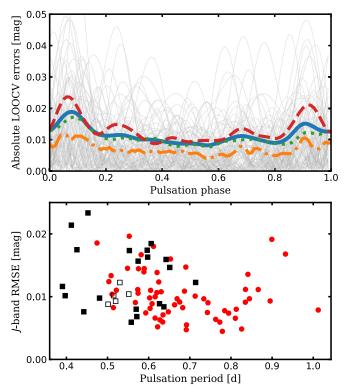


Figure 7. Top: absolute J-band residuals for our approximation of the J-band light-curve shape using the PC amplitudes, as a function of the pulsation phase. Thin gray lines show the individual absolute residuals, according to our LOOCV. The continuous, dashed, and dotted lines mark the average errors for the complete sample, the stars of  $\omega$  Centauri, and the sample omitting the  $\omega$  Centauri stars, respectively. The dotted–dashed line shows the median error for each phase bin. Bottom: J-band RMSEs from our LOOCV. The symbols are the same as in Figure 5.

the magnitude at maximum light, from relatively scarce photometry. In contrast to the study of Tanvir et al. (2005), we implement a method to fit the PCs directly to the VVV  $K_{\rm S}$ -band light curves. In Section 2.3, we determined the PCs on a grid of light-curve phases. However, to fit a target light curve, we have to discern their PC amplitudes  $U_i$ , using the PCs as basis functions. As PCs are not continuous functions, we transform them to a Fourier representation of the form

$$\mathcal{F}_{\text{PC},i}(x) = \sum_{k=1}^{12} [A_{ik} \sin(2\pi kx) + B_{ik} \sin(2\pi kx)], \qquad (3)$$

where x is the light-curve phase, using OLS fitting of the individual PCs. The continuous lines on the left panel of Figure 8 demonstrate the transformed PCs. These new representations can be used as basis functions to describe the light-curve shapes of RRLs as

$$\mathrm{LC}(t) = m_0 + \sum_{i=1}^{4} \left[ U_i \times \mathcal{F}_{\mathrm{PC},i} \left( \frac{t - E_0}{P} \right) \right],\tag{4}$$

where t is the Julian Date, and the free parameters are the PC amplitudes  $U_i$ , the magnitude average  $m_0$ , the period P and the zero epoch  $E_0$ .

The light-curve representation proposed by Equation (4) is not a linear function; therefore, we have to use nonlinear regression. We developed a method to adjust these seven free parameters, and the model presented in Equation (4), to the light curves of RRLs in the VVV survey. Unfortunately, in the more crowded fields and near bright stars, VVV photometry tends to contain a fraction of outlying points. Traditionally, iterative threshold rejection is utilized to omit outlying measurements in massive time-series analysis. However, in cases of unfortunate data distribution, and due to the fact that in the first iteration erroneous observations can have a large effect on the fit, this can result in the flagging of good data as outliers. To avoid this problem, we replace the normally used squared error loss with the Huber loss function (Huber 1964) of the form

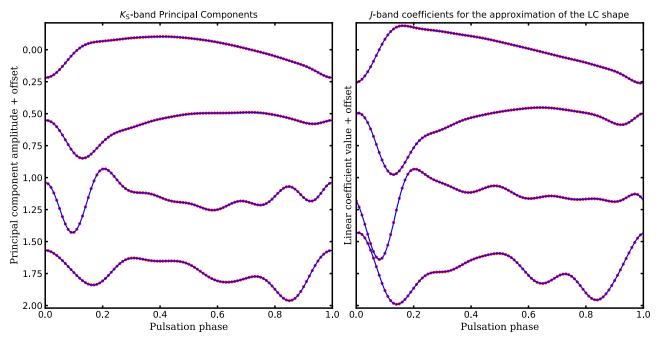
$$L_{\delta}(y, f(x)) = \begin{cases} \frac{1}{2}(y - f(x))^2 & \text{for } |y - f(x) < \delta| \\ \delta |y - f(x)| - \frac{1}{2}\delta^2 & \text{otherwise.} \end{cases}$$
(5)

This loss function behaves identically to the squared error loss for data points with residuals smaller than  $\delta$ . For residuals bigger than  $\delta$ , the loss grows linearly with increasing residuals. Therefore, outlying points weigh less than they would, if squared error loss was utilized, a convenient feature in the case of outliers.

The distribution of PC amplitudes of our PCA sample (Figure 5) gives us a priori information on the possible shapes of RRL light curves, if we accept the sample analyzed in Section 2 as representative of RRLs. We utilize this information as follows. We require the first PC amplitude to be in the range  $U_1 \in [U_{1,\max} + \delta_{U_1}, U_{1,\min} - \delta_{U_1}]$ , where  $U_{1,\max}$ and  $U_{1,\min}$  are the largest and smallest  $U_1$  amplitudes of the PCA training sample (see Section 2.3), respectively, and  $\delta_{U_1} = (U_{1,\text{max}} - U_{1,\text{min}})/10$ . In other words, we limit the  $U_1$ amplitudes to not be larger or smaller than the largest or smallest  $U_1$  amplitude in the PCA sample,  $\pm 10\%$  of the total range of  $U_1$  amplitudes given by the PCA. As for the other amplitudes, we require their relative values,  $U_2/U_1$ ,  $U_3/U_1$  and  $U_4/U_1$ , to adhere to analogous requirements on their ranges. These requirements ensure that the fit light curves have shapes that are similar to those in the PCA sample.

The following steps have been implemented to ensure that the light-curve shapes and  $JHK_S$  average magnitudes are derived with the highest possible accuracy, based on the VVV data:

- 1. all  $K_{\rm S}$ -band light-curve points brighter than 9 mag, fainter than 20 mag, as well as points farther than  $\pm 0.5$  mag from the median are discarded;
- 2. the remaining light-curve points are fit with the light-curve model of Equation (4), while utilizing the Huber loss function in the form of Equation (5), with  $\delta = 0.05$  mag, and the constraints on the ranges of  $U_1$  and  $U_{2,3,4}/U_1$ , as described above;
- 3. we omit points with a residual larger than  $3.5\sigma$ , where  $\sigma$  is determined from the absolute median deviation of the fit from step 2;
- 4. the fit is repeated on the remaining data in the same way as described in step 2;
- 5. as the light-curve shapes of RRLs are very similar in the H and  $K_{\rm S}$  bands (e.g., they are the same with a scatter of  $\sim 0.02$  mag, depending on the pulsation phase; see Figures 1 and 2 of Barnes et al. 1992), we use the  $K_{\rm S}$  light-curve shapes to determine the average magnitudes by fitting them directly to the *H*-band measurements; and
- 6. the *J*-band light-curve shapes are predicted using the  $U_i$  magnitudes, as described in Section 2.4, and these are fit to the *J*-band measurements to determine the mean *J* magnitude.



**Figure 8.** Left: the first four principal components (red circles) and their continuous Fourier representations (blue lines), utilizing Equation (3). Right: the coefficients for the approximation of the *J*-band light-curve shape in each phase (red points) and their Fourier representations (blue lines). Note that the coefficients utilized to reconstruct the *J*-band light curves describe similar light-curve features as the  $K_s$ -band PCs, albeit with larger amplitudes, especially around the rising branch. This behavior is in line with the fact that RRLs have larger amplitudes at shorter wavelengths.

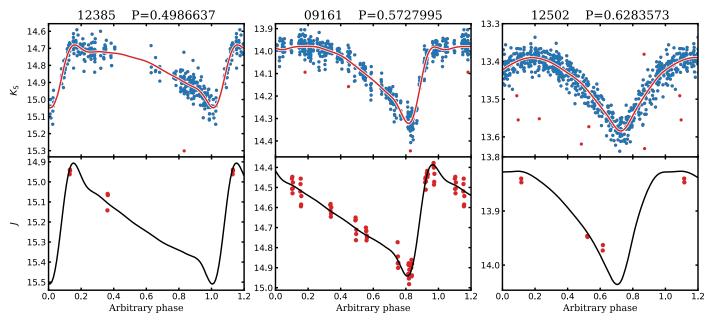


Figure 9. Examples of the near-IR light-curve fitting methods implemented in Section 3. OGLE IDs and periods are shown on the top of the panels. Top:  $K_S$ -band light curves and their fits using the first four PCs derived in Section 2. Bottom: *J*-band light-curve points and the approximations of the *J*-band light-curve shapes from the PC amplitudes, as described in Section 2.4. Variable no. 12385 (left panel) illustrates the case of a single large gap in the folded light curve, a challenge for traditional Fourier fitting. No. 09161 has an unusually large number of *J*-band light-curve points, allowing us to demonstrate the quality of the predicted *J* light curve. No. 12502 has a much more symmetric light-curve shape, but our method is flexible enough to find a good fit for it.

The principal outputs of this procedure are the  $K_{\rm S}$ -band lightcurve parameters, as well as the robust magnitude estimates of the target RRLs in the *JHK*<sub>S</sub> bands. Figure 9 illustrates the VVV  $K_{\rm S}$ -band fits (top), as well as the predicted *J*-band light curves of three RRLs from the OGLE bulge RRL sample of Soszyński et al. (2014). The implementation of all these steps in a single, convenient routine is available on GitHub.<sup>10</sup>

# 4. Metallicity Estimation from K<sub>S</sub>-band Photometry

The idea of estimating the metallicity of RRLs from their Fourier light-curve parameters originates from Kovács & Zsoldos (1995). Their original method was substantially improved by Jurcsik & Kovács (1996), who found a simple linear relationship between the iron abundances of RRLs, their periods and the epoch-independent Fourier phase differences  $\phi_{31} (=\phi_3 - 3\phi_1$ , where  $\phi_3$  and  $\phi_1$  are the Fourier phases of the

<sup>&</sup>lt;sup>10</sup> https://github.com/gerhajdu/pyfiner

third and first harmonics of the light curve, respectively; Simon & Lee 1981) of their V-band light curves. Their formula was calibrated by spectroscopic measurements and corresponds to a metallicity scale established by high-dispersion spectroscopy (Jurcsik 1995). Following the example set by Jurcsik & Kovács (1996), Smolec (2005) developed similar formulae for the Cousins I-band, notable as the band in which most observations of the OGLE surveys are being carried out (Udalski et al. 2015). Smolec (2005) gave two alternative formulae for the derivation of the iron abundance, a 2- and a 3-term formula, and he argued that the latter, which also includes the amplitude of the second Fourier harmonic, provided better results. We emphasize that both the V- and I-band formulae have a residual scatter of only  $\sim 0.14$  dex, which is similar to the accuracy of spectrophotometric methods, such as the  $\Delta S$  method (Preston 1959), and are on the same metallicity scale defined by Jurcsik (1995).

A similar calibration between [Fe/H] and the near-IR lightcurve parameters has long been lacking, despite that with the advent of large time-domain near-IR photometric surveys, such as the VVV and the VISTA Survey of the Magellanic System (Cioni et al. 2011), among others, such an empirical calibration is of key importance, as a large fraction of the newly discovered distant RRL stars along the Galactic plane are beyond the faint magnitude limit of optical surveys. In the following, we are going to detail our calibration of such a relationship, utilizing the overlapping RRLs found by the OGLE project (Soszyński et al. 2014) and observed by the VVV survey.

#### 4.1. VVV Photometry of Bulge RR Lyrae Variables

We employed our light-curve fitting method described in Section 3 to determine the  $K_{\rm S}$ -band light-curve parameters of RRab variables listed in the OGLE-IV bulge RR Lyrae catalog (Soszyński et al. 2014). Toward this end, we made use of the data processed by the VISTA Data Flow System (VDFS; Emerson et al. 2004), provided by the CASU. The aperture photometry of detector frame stacks (pawprints) was extracted using the Starlink Tables Infrastructure Library (STIL; Taylor 2006) at the coordinates of the OGLE RRLs. This process has resulted in near-IR light curves for 24217 RRLs.

The VDFS provides photometry in circular apertures of different radii, but due to the source crowding in the Galactic bulge fields, the adopted radii have to be chosen carefully. Generally, smaller apertures provide better photometry for dimmer objects, but in relatively uncrowded cases, this does not always hold true. Furthermore, variables found in overlapping regions of tiles and pawprints can have different offsets between them, but generally the best aperture also minimizes this offset. Therefore, for each star, the optimal aperture has to be chosen on a case-by-case basis.

We utilized our fitting procedure described in Section 3 on the five smallest apertures provided by the VDFS for each OGLE RRL. It has been found that the final value of the Huber cost function divided by the number of data points after the  $3.5\sigma$  clip is a good indicator of the quality of the photometry obtained with a given aperture with respect to the others. Hence, in this section we are going to utilize the PC amplitudes determined for each star using the aperture for which this value is the smallest.

# 4.2. Unbiased Photometric Metallicities from the RR Lyrae I-band Light Curves

Due to the large number of RRLs with both OGLE-IV and VVV photometry, we decided to search for correlations between the abundances determined on the former, and the light-curve shapes of the latter. As the main bulge population of RRLs have spectroscopic iron abundances of  $[Fe/H] \sim -1$  (Walker & Terndrup 1991), this component is expected to be dominant in the calculated photometric metallicities as well.

The OGLE-IV (or OGLE-III, if the former was not available) *I*-band light curves of RRLs in the common OGLE/VVV sample were fit with a sixth-order Fourier series to calculate the Fourier parameters necessary for the metal abundance formulae of Smolec (2005).

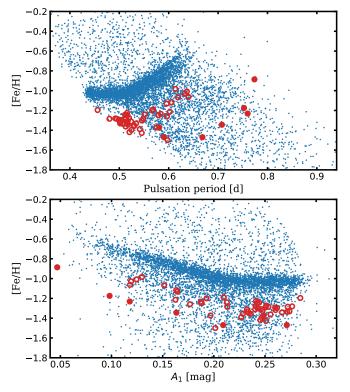
As our goal is to provide a method for estimating metallicity based on the near-IR light-curve parameters, we eliminate all variables where either estimate is uncertain.

By far, the most common reason for the elimination of variables was the presence of the Blazhko effect, which leads to distorted light-curve shapes in RRLs, which are never equivalent to those of non-modulated RRLs (Jurcsik et al. 2002). Prudil & Skarka (2017) studied the Blazhko effect on a subsample of bulge RRLs in the OGLE sample, and found an incidence rate of 40%. We have inspected each I-band light curve and its fit visually to reveal the presence of the Blazhko effect. In dubious cases, we have inspected the Discrete Fourier Transform of the residual light curve for signs of the characteristic side peaks of modulation in the Fourier domain. We opted to eliminate all variables from our selection where inspection of the light curve gave hints of the Blazhko effect, which accounted for approximately the same fraction of stars as found by Prudil & Skarka (2017). Furthermore, this inspection revealed some dubious RRL light curves, as well as many light curves where the Fourier parameters cannot be determined with high accuracy (faint variables, too few lightcurve points, gaps in the folded light curve, etc.), which were also removed in order to preserve the purity of the sample.

Besides requiring good-quality light curves in the *I*-band, the  $K_{\rm S}$ -band data were also subjected to the following quality cuts: we discarded all variables where either the first PC amplitude,  $U_1$ , or the PC amplitude ratios  $U_2/U_1$ ,  $U_3/U_1$ , or  $U_4/U_1$  exceeded the limit described in Section 3, as these indicate problematic photometry in the VVV data, such as severe blending.

We utilized Equations (2) and (3) of Smolec (2005) to calculate the metal abundances of the 6215 remaining variables. The top panel of Figure 10 reveals a striking systematic on top of the overall linear trend in the photometric metallicity calculated with Equation (2) of Smolec (2005) as a function of the pulsation period. In the main locus of stars, denoting the Oosterhoff I population of the bulge RRLs, longer-period RRLs seem to have systematically higher metallicities than their shorter-period counterparts. Moreover, a similar trend can be seen on the bottom panel: the lower-amplitude stars have systematically higher metallicities.

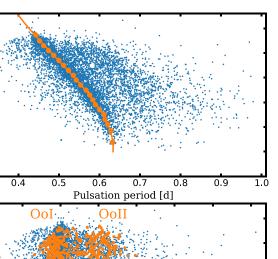
In globular clusters, for both Oosterhoff groups of fundamental-mode variables, the amplitude decreases with increasing periods. It is hard to conjure up a scenario where systematically higher-metallicity RRLs would only populate the lower-amplitude, longer-period part of the main locus on this diagram, while the lower-metallicity variables occupy the higher-amplitude, shorter-period part. Therefore, we conclude



**Figure 10.** Photometric iron abundances calculated from the *I*-band light-curve parameters with Equation (2) of Smolec (2005) as a function of the pulsation period (top) and as a function of the amplitude of the first harmonic of the Fourier series fit (bottom), both times indicated by small blue points. Oo I and II variables from M3 are overplotted as large empty and filled circles, respectively. Both the bulge field and the M3 RRLs display a systematic bias toward higher abundances as the pulsation period increases and the pulsation amplitude decreases. Furthermore, at a given period or amplitude, the metallicities derived for the Oo I and II stars in M3 (calculated from the photometry of Jurcsik et al. 2017) at similar amplitudes are systematically offset.

that the metal abundance formula described by Equation (2) of Smolec (2005) suffers from a systematic bias as a function of amplitude/period. This finding is confirmed by a demonstrably monometallic data set: on both panels of Figure 10, empty and filled circles illustrate the *I*-band photometric metallicities of the Oosterhoff I and II variables in the globular cluster M3, which is well known to be monometallic (Kraft et al. 1992; Cohen & Meléndez 2005), highlighting the systematic offset between these two groups of objects as well.

The second formula given by Smolec (2005) in the form of Equation (3) gives much more consistent abundance estimates as a function of the amplitude. However, to compare the behavior of the calculated abundances of the two Oosterhoff groups of variables, we have to separate them. This has been done using the period-amplitude diagram, utilizing a cut that is dependent on the calculated iron abundance, as illustrated by Figure 11. Comparing the top and middle panels of Figure 12 indicates that there is still a slight trend for the Oosterhoff I variables as a function of the amplitude, as well as an offset with respect to the Oo II stars, which have a calculated average [Fe/H] of -1.05 dex across the whole amplitude range. We correct for this offset using the ridge of Oo I RRLs, derived from a KDE of amplitude bins of the estimated abundances of Oo I stars. The resulting rectified Oo I metallicity estimates are consistent with the Oo II variables, as well as across the whole range of RRL amplitudes (bottom panel of Figure 12).



All stars

KDE

03

0.2

-1.2 < [Fe/H] < -1.1

04

05

0.30

0.25

0.20

0.15

0.10

0.05

0.00

0.30

0.25

0.20

0.15

0.10

0.05

0.00

0.2

-0.1

0.0

A1 [mag]

A1 [mag]

Rectified pulsation period Figure 11. Iron abundance-dependent separation of Oosterhoff classes. Top: the distribution of the amplitudes of the first harmonic of the I-band Fourier fits of bulge RRLs as a function of the pulsation periods. The main ridge of Oo I stars is localized using the kernel density estimate (KDE) maxima in different amplitude bins (circles). These are fit with a third-order polynomial, illustrated by the continuous line. Bottom: same as above, but as a function of the rectified pulsation periods, calculated as the difference between the real pulsation period and the position of the Oo I ridge of the top panel, as marked by the continuous line. The RRLs in the abundance bin -1.2 < [Fe/H] < -1.1 are marked, and the continuous line at the bottom shows their KDE distribution. The local minima in the middle of the KDEs of different abundance bins change, reflecting the metallicity dependence of the Oosterhoff phenomenon. We classify RRLs as Oo I or II based on their position in this diagram, using a criterion that is a linear function of their calculated iron abundance. The decision criteria for [Fe/H] = 0.0 and -2.0 are marked by the vertical black lines.

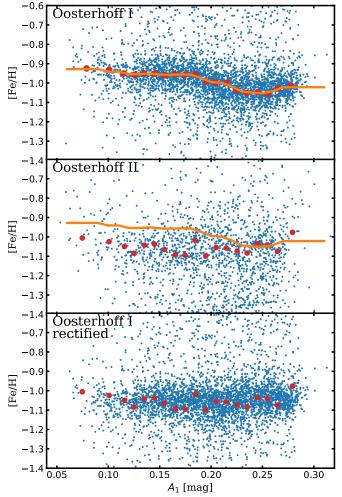
01

We do note that not only the iron abundance formulae of Smolec (2005) suffer from obvious biases when applied on a population of monometallic RRLs, nor is the problem limited to the *I*-band. As a stopgap measure (until new, carefully calibrated photometric abundance formulae become available) we make available a routine on GitHub<sup>11</sup> that implements the procedure outlined here to rectify the [Fe/H] values of Oo I RRLs calculated from Equation (3) of Smolec (2005), given the period, as well as the *I*-band Fourier parameters  $A_1$ ,  $A_2$  and  $\phi_{31}$ .

# 4.3. Iron Abundance Estimation from the K<sub>S</sub>-band Light-curve Shapes

To assess whether it is possible to determine the iron abundance of RRLs from the  $K_S$  band light-curve shapes, we made use of the  $K_S$ -band light-curve parameters of OGLE RRLs in the VVV fields determined in Section 4.1, combined with the photometric metallicities determined in Section 4.2. As the RRL sample on which Smolec (2005) calibrated the relation used to determine the initial [Fe/H] of the RRLs in

<sup>&</sup>lt;sup>11</sup> https://github.com/gerhajdu/pyrime



**Figure 13.** Dependence of the rectified photometric (*I*-band) iron abundances on the near-IR light-curve parameters. Top: comparison of the KDE of the period distribution of the bulge (blue contiguous line) and the PCA training sample (Table 1). Although the period distributions of the two samples are different, the example light-curve fits of Figure 9 reveal that the method described in Section 3 is well suited for fitting the bulge RRL light curves. Middle: the iron abundances show a significant overlap when only the period and the first PC amplitude are considered. Bottom: at a given period, and  $U_1$ amplitude stars with different [Fe/H] possess different  $U_2$  amplitudes, illustrating the iron abundance dependence of the near-IR light curve shape parameters.

As additional features, we also consider Fourier parameters

for this regression problem. During the light-curve fitting, the

individual PCs are represented as Fourier series (Equation (3)),

and the total light curve is given as the sum of these series

multiplied by the PC amplitudes (Equation (4)). Therefore, traditional Fourier parameters, such as the Fourier amplitudes and epoch-independent phase differences (Simon & Lee 1981),

can be calculated straightforwardly. Of these, we have decided

to use the amplitude of the Fourier harmonics  $A_1$ ,  $A_2$ , and  $A_3$ , as

well as the epoch-independent phase differences  $\phi_{21}$  and  $\phi_{31}$ .

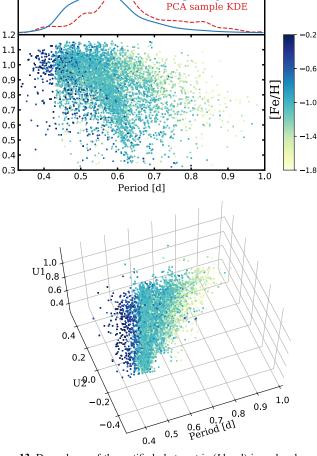
Together with the pulsation period, the total number of

Figure 12. Photometric iron abundances (small blue points) calculated from the I-band light-curve parameters with Equation (3) of Smolec (2005). Top: the photometric metallicity estimates of Oosterhoff I variables. As can be seen, these still display a dependency on the amplitude of the first Fourier harmonic, albeit to a much smaller degree than found using Equation (2) of Smolec (2005) (Figure 10). Red filled circles denote the average metallicities, as calculated from the maximum value of a KDE of data points in different amplitude bins. K-nearest regression (with K = 2, marked as a continuous orange line) was utilized to generalize this effect, and this function was used to rectify the individual abundance estimates. Middle: photometric metallicity estimates of Oosterhoff II variables. The KDE estimates (red filled circles) of the amplitude bins reveal no significant dependency on the estimated metallicities. The contiguous orange line from the top panel is drawn for comparison purposes. Bottom: the rectified metallicity estimates of the Oosterhoff I RRLs. The KDE estimates from the middle panel (filled red circles) are repeated for comparison purposes.

question only extends down to [Fe/H] = -1.7, using even the rectified abundance estimates below this threshold is uncertain. Nevertheless, because the typical abundance of Oo II globular clusters bearing RRLs is about  $[Fe/H] \sim -2.2$  dex, we are only discarding the RRLs below this limit. This final cut results in a final sample of 6193 RRLs.

Figure 13 shows the optical photometric metallicity as a function of the period, and the first two PC amplitudes. Some general trends are apparent; for example, as expected from the form of Equation (3) of Smolec (2005), longer-period RRLs have higher iron abundances, but even at the same period and  $U_1$ , stars of different  $U_2$  have systematically different iron abundances.





U

 Table 2

 Parameters of Our Hyperparameter Grid Search

Hyperparameter	Candidate Parameter Value
Hidden layers <sup>a</sup>	(20), (50), (100) (20, 20), (20, 50), (20, 100) (50, 20), (50, 50), (50, 100) (100, 20), (100, 50), (100, 100)
α <sup>b</sup>	$\begin{array}{c} 10^{1.0}, \ 10^{0.5}, \ 10^{0.0}, \ 10^{-0.5}, \ 10^{-1.0} \\ 10^{-1.5}, \ 10^{-2.0}, \ 10^{-2.5}, \ 10^{-3.0} \\ 10^{-3.5}, \ 10^{-4.0}, \ 10^{-4.5}, \ 10^{-5.0} \end{array}$
Activation function <sup>c</sup>	logistic, tanh, relu
Solver <sup>c</sup>	lbfgs, sgd, adam
Independent variables	$\begin{array}{c} P, \ U_{14} \\ P, \ U_{13}, \ \phi_{31} \\ P, \ U_{13}, \ \phi_{21}, \ \phi_{31} \\ P, \ A_{13}, \ \phi_{21}, \ \phi_{31} \\ P, \ U_{14}, \ A_{13}, \ \phi_{21}, \ \phi_{31} \end{array}$

#### Notes.

<sup>a</sup> Number of neurons of hidden layers of the neural network.

<sup>b</sup> Regularization parameter.

<sup>c</sup> Activation function and solver of optimization; see the scikit-learn documentation for description at http://scikit-learn.org/stable/modules/generated/sklearn.neural\_network.MLPRegressor.html.

with metallicities around [Fe/H] = -1 dex biases any kind of regression without resampling or weighting the input data.

The best results were achieved using the MLPRegressor routine, implementing a multilayer perceptron regressor, a type of neural network. Artificial neural networks are inspired by and modeled after biological neural networks, and have been used in many branches of science as well as in commercial applications for their ability to learn highly nonlinear relations inherent in many types of data (see Haykin 2011 for a general description of neural networks).

Neural networks have many different parameters that have to be decided (hyperparameters) before the training of the network. As these can vastly influence their performance, we decided to train the MLPRegressor with a grid of different hyperparameter values, and use cross validation to determine the best combination of hyperparameters. The grid of considered hyperparameter values is documented in Table 2. This grid only contains possible neural network architectures up to two hidden layers of neurons; although adding more hidden layers increases the flexibility of the neural network, doing so did not significantly improve the performance in our tests. Besides the parameters of the network, we considered different combinations of the 10 independent variables as an additional hyperparameter.

Due to the heavily biased nature of the bulge photometric metallicity distribution (most stars have abundances  $[Fe/H] \sim -1$ ), we implemented a special sampling method. We grouped observations in ten 0.2 dex wide bins (Table 3). Because there are only a few variables with [Fe/H] above 0 and below -2 dex, these were merged with neighboring bins. In each bin, 10 stars were selected randomly to be part of the validation sample. Of the remaining stars, 80 variables were selected randomly with replacement to be part of the training sample, resulting in a training set of 800 data points. As there are less than 90 stars in the most metal-poor and metal-rich bins, some variables at the extremes of the [Fe/H] distributions are

 Table 3

 Number of RR Lyrae Variables in Different Metallicity Bins

[Fe/H] Bin	Number of RRLs
[-0.2, +0.2]	47
[-0.4, -0.2]	102
[-0.6, -0.4]	225
[-0.8, -0.6]	295
[-1.0, -0.8]	487
[-1.2, -1.0]	3888
[-1.4, -1.2]	628
[-1.6, -1.4]	375
[-1.8, -1.6]	84
[-2.2, -1.8]	62

selected multiple times due to the selection with replacement. The variables that were not selected for training were added to the cross-validation sample. In case of an uneven distribution of stars between the two extremes in a metallicity bin, the bin was split into sub-bins to guarantee a more even sampling as a function of metallicity.

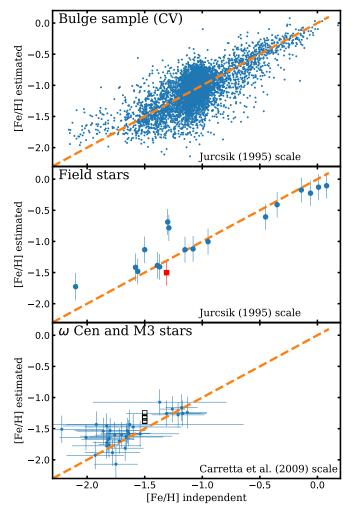
This separation into training and cross-validation samples guarantees that the training set is large enough for the training of neural networks; the training set is balanced on the total range of possible abundances; by repeating this separation, variables in bins with few stars get selected for the crossvalidation sample at least a few times.

This separation was repeated 40 times, resulting in 40 training and cross-validation samples. The neural networks were trained on all 40 training samples with all combinations of the possible hyperparameters detailed in Table 2. In all cases, the predictions of the trained neural networks were calculated for the corresponding cross-validation samples. Finally, for each star and for each hyperparameter combination, the predictions were averaged over the 40 repeats. Every star appeared at least six times in the cross-validation samples.

Due to the imbalanced data set, neural networks with the highest  $R^2$  score (also called the coefficient of determination) calculated on the complete sample of averaged predictions is heavily biased toward solutions predicting [Fe/H]  $\sim -1$ , irrespective of the values of the independent parameters. Therefore, we have instead selected the best hyperparameter combination by using the sum of the  $R^2$  scores calculated separately for each of the abundance bins for the averaged cross-validation predictions. The cross-validation result with the highest score is illustrated by the top panel of Figure 14, with features P,  $A_1$ ,  $A_2$ ,  $A_3$ ,  $\phi_{21}$ , and  $\phi_{31}$ ; hyperparameters  $\alpha = 10^0$ ; two hidden layers with 100 and 20 neurons; activation function relu; and solver lbfgs.

After determining the optimal hyperparameters, 100 new training samples were drawn by randomly selecting 80 stars with replacement from each metallicity bin, but without withholding any stars for cross validation. The MLPRegressor was trained on each of these samples, and the predicted metallicity is the average of these 100 trained regressors. The code estimating the iron abundance using these neural networks from the  $K_{\rm S}$ -band light-curve parameters is available on GitHub.<sup>12</sup>

<sup>&</sup>lt;sup>12</sup> https://github.com/gerhajdu/pymerlin



**Figure 14.** Top: cross-validated (CV) average  $K_S$ -band abundance estimate for 6193 RRLs in the OGLE/VVV sample against the rectified *I*-band photometric abundance estimates (Section 4.2). The dashed orange curve represents equality between the compared abundance estimates. Middle: comparison of estimated photometric abundances of field RRLs (circles) with their spectro-photometric ( $\Delta S$ -based) abundances from Table 1 of Jurcsik & Kovács (1996) and the star NSV 660 (red square), the field variable with a spectroscopic SEGUE metallicity estimate (Lee et al. 2011; Szabó et al. 2014). Individual errors are estimated as the quadratic sum of the scatter of the 100 abundance estimates (Section 4.3) and a fixed uncertainty term of 0.2 dex. Bottom: comparison of the estimated photometric abundances of stars in the clusters  $\omega$  Cen (blue dots) and M3 (black open squares) with the individual spectroscopic measurements (Sollima et al. 2006) and the overall cluster iron abundance (Carretta et al. 2009), respectively.

## 4.4. Validation of the Metallicity Estimates

To provide an unbiased evaluation of the performance of a regression model, it is important to test the performance on a test data set not used during the fit for either training or validation. In this case, the model was used to predict the iron abundance from the shape parameters determined directly from the PCA analysis for the variables in Table 1. Then, these predictions can be compared to independent determinations of their abundances from the literature.

The cross-validation results in the top panel of Figure 14 indicate that the method described in Section 4.3 gives reasonable photometric abundance estimates with a scatter of  $\sim 0.22$  dex in the -1.7 < [Fe/H] < 0.0 abundance range. The estimated iron abundance is biased for stars below -1.7 toward higher values, probably caused by the combination of

overestimation of the *I*-band photometric abundances, as well as contamination of the parameter space of low-abundance stars with those of higher estimated abundances, due to the higher relative photometric errors in low-amplitude, longperiod stars (see Figure 13).

The middle panel of Figure 14 compares the photometric abundances of 17 stars from our Table 1 to the abundances listed in Table 1 of Jurcsik & Kovács (1996). Additionally, the photometric abundance of NSV 660 is compared with the [Fe/H] = -1.31 value determined from the SDSS spectra of the SEGUE Stellar Parameter Pipeline (Lee et al. 2011; Szabó et al. 2014). The two values for this sample of stars generally agree. We note, however, that there is a hint that the abundances of high-metallicity RRLs are slightly underestimated. Furthermore, the [Fe/H] of the most metal-poor star is overestimated, in concordance with the cross-validation results. The abundances of two stars, RR Cet and RR Leo, are overestimated by 0.5 and 0.6 dex, respectively. As there is a single source of photometry for both stars, we believe that the original  $K_{\rm S}$ -band photometry of both stars is affected by systematic trends on their rising branches, causing the estimated light-curve shapes to be deformed, resulting in the overestimation of their abundances.

The bottom panel of Figure 14 compares the  $K_{\rm S}$ -band photometric abundances of the RRLs in  $\omega$  Cen with their spectroscopic measurements by Sollima et al. (2006). The photometric abundances for five variables from the globular cluster M3 are also compared to the average cluster abundance, [Fe/H] = -1.5 dex, given by Carretta et al. (2009). We do note that so far all abundances were on the metallicity scale defined by Jurcsik (1995). To compare the  $K_{\rm S}$ -band photometric abundance estimates of the  $\omega$  Cen and M3 variables in a consistent way to their spectroscopic determinations, we convert them to the metallicity scale established by Carretta et al. (2009). The conversion between the two scales can be determined by comparing the common clusters with measured abundances in Table 1 of Jurcsik (1995, J95) and Table A1 of Carretta et al. (2009, C09) (neglecting the cluster NGC 5927, where the difference between the two sources is almost 0.8 dex):

$$[Fe/H]_{C09} = 1.044 [Fe/H]_{J95} - 0.037,$$
 (6)

with a residual scatter of 0.1 dex. Furthermore, the iron abundances of Sollima et al. (2006) are also converted to the Carretta et al. (2009) scale by shifting them by -0.02 dex, following Braga et al. (2016).

The photometric abundances of  $\omega$  Cen stars reproduce the bimodal distribution of iron abundances (Sollima et al. 2006). When taking into account the uncertainties of individual measurements made by Sollima et al. (2006), there is no offset for the metal-rich group at -1.2 dex. For the metal-poor group of variables, there is a hint that our method systematically overestimates abundances by about 0.1 dex, similarly to the results on the top and middle panels of the Figure 14. As for the M3 RRLs, their average  $K_{\rm S}$ -band photometric abundance estimate is [Fe/H] = -1.33 dex, only 0.17 dex higher than the cluster metallicity given by Carretta et al. (2009).

During the calibration of the abundance prediction, we discarded Blazhko variables, as their modulated light curves lead to systematic biases in their calculated photometric abundances (see, e.g., Figure 4 of Jurcsik & Kovács 1996). Nevertheless, a significant fraction of the total population of

RRLs suffers from this light-curve modulation. The nature of the Blazhko effect in the near-IR has not been established until very recently, due to the lack of well-populated light curves covering sufficiently long time spans. In Jurcsik et al. (2018), we show that the modulation is indeed present in the  $K_{\rm S}$ -band light curves, but with diminished amplitudes. We have no abundance estimates for the 22 Blazhko RRLs studied in Jurcsik et al. (2018), but we surmise that the majority of them must come from the bulge population of RRLs. We calculated the abundances based on their individual  $K_{\rm S}$ -band mean light curves, for which we get a mean of [Fe/H] = -1.11 dex with a standard deviation of 0.28 dex. These values are in agreement with those of the original bulge sample, where the mean and standard deviation are [Fe/H] = -1.06 dex and 0.26 dex, respectively. Therefore, we surmise that the iron abundances of RRLs should not show systematic biases when estimated with the parameters of the complete, long-term near-IR light curves, even when the Blazhko effect is present.

In summary, the  $K_{\rm S}$ -band light-curve parameters allow the estimation of the iron abundance of RRLs to an accuracy of 0.20–0.25 dex, with slight hints of systematic trends (at the <0.05 dex level) in the range of  $-1.7 < [{\rm Fe}/{\rm H}] < 0$  (underestimation for high abundances, overestimation for low ones), but abundances lower than this range are systematically overestimated. To establish better relations, one will need high-quality  $K_{\rm S}$ -band observations of a large sample of RRLs in globular clusters covering a wide metallicity range, and/or  $K_{\rm S}$ -band observations of the hundreds of bright RRLs in the Solar neighborhood, and/or spectroscopic iron abundances for the OGLE field RRLs, with emphasis on the high and low-abundance extremes.

# 5. Summary

In this study, we have analyzed the near-IR light-curve properties of RRab variables. The principal component analysis of the  $K_{\rm S}$ -band light curves of a sample of 101 RRLs revealed that their varied shapes can be described as a linear combination of a low number of PCs. It has to be stressed that compared with most previous studies involving PCA of light curves, we decided to phase our variables by the light-curve minimum, as the near-IR maxima of most RRLs is nearly flat and hard to localize. We advocate the exploration of alternatives to phasing the light curves of pulsating variables by their maxima, even for studies in the optical. Possibilities include, among others, phasing by the light-curve minima, as was done here; by the middle of the rising or descending branches; or by the phase of the first harmonic of a Fourier series.

The comparison of the  $K_{\rm S}$ -band light curve parameters and the *J*-band light curves of 87 variables led to the conclusion that the former can be used to reliably predict the light-curve shape in the *J* band. This finding is of great interest for timeseries surveys where most data points are taken in a single filter (as is the case of the VVV survey), as the accurate approximation of the light-curve shape in a different band can greatly reduce the number of observations needed to determine accurate mean magnitudes in complementary bands, hence truly representative colors. This can be especially useful for the estimation of the amount of foreground reddening toward individual stars.

We developed a method to fit the RRL light-curve shapes in the  $K_S$  band with the PCs as basis vectors, while also taking into account some of the peculiarities of the VVV data. The routine takes advantage of the robustness of the Huber loss function to decrease the effect of outliers in VVV light curves. The *J*-band light-curve shapes are also approximated from the PC amplitudes, resulting in accurate mean magnitudes in the  $JHK_S$  bands.

The possibility of deriving metallicities from the  $K_{\rm S}$ -band light curve parameters, analogous to what was done in the optical (Jurcsik & Kovács 1996), was explored with the help of RRLs in common between the VVV and OGLE surveys. The PC amplitudes of these stars were determined using our fitting method on the VVV light curves. However, the photometric metallicities of these variables, as calculated using Equation (2) of Smolec (2005) on the I-band light-curve parameters, revealed a systematic trend with the period/amplitude of the main locus of variables, representing the RRL population of the Galactic bulge. The same trend is also seen in other populations, as demonstrated by Figure 10 for the RRLs of the monometallic globular cluster Messier 3. This problem with photometric metallicities is especially worrying in view of their prevalent use when tracing old populations with RRLs. Therefore, we emphasize the necessity of calibration and validation of such relations on stars spanning the whole range of possible amplitudes, periods, lightcurve shapes and iron abundances of RRLs.

The relation presented by Smolec (2005) in the form of Equation (3) has a smaller, but still detectable trend as a function of the amplitude, which we have corrected for. These rectified photometric iron abundance estimates were used to look for relations between [Fe/H] and the  $K_{\rm S}$ -band light-curve shapes of the bulge RRL sample. As the apparent relations are nonlinear, and the data set is unbalanced, we have decided to use the aggregate of many separate neural networks trained on balanced sub-sets of the total data set. The resulting accuracy for the determination of individual iron abundances based on the  $K_{\rm S}$ -band light curves is about 0.20–0.25 dex.

The methods developed in this work are utilized in a companion paper (Dékány et al. 2018) to characterize the RRL population of the disk area of the VVV survey, allowing the estimation of the metallicity distribution function of RRL stars in that area. Some of these approaches could be adopted relatively easily for the exploration of other sources of timeseries photometry, for example where the small number of data points or the disparate amount of multiband observations do not lend themselves well to traditional (i.e., Fourier based, in the case of pulsating stars) methods. Naturally, in order to utilize such *prior* knowledge, a high-quality (as well as preferably expansive) training sample is needed to characterize the population of variable stars in question.

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Facility: ESO:VISTA.

Software: scikit-learn (Pedregosa et al. 2011), scipy (Jones et al. 2011).

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